

Recursive Distributional Equations and Recursive Tree Processes : Lecture - II

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Recursive Distributional Equation (RDE)

Definition 1 *The following fixed-point equation on \mathcal{P} is called a Recursive Distributional Equation (RDE)*

$$X \stackrel{d}{=} g\left(\xi; \left(X_j, 1 \leq j \leq^* N\right)\right) \quad \text{on } S,$$

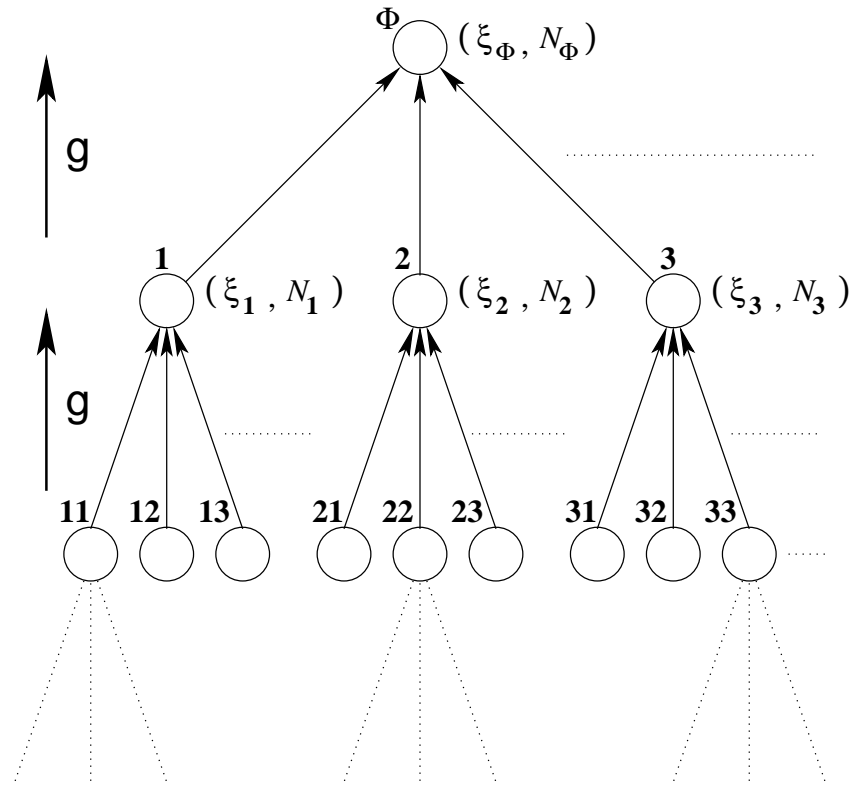
where $(X_j)_{j \geq 1}$ are independent copies of X and are independent of (ξ, N) .

Remark : A more conventional (analysis) way of writing the equation would be

$$\mu = T(\mu)$$

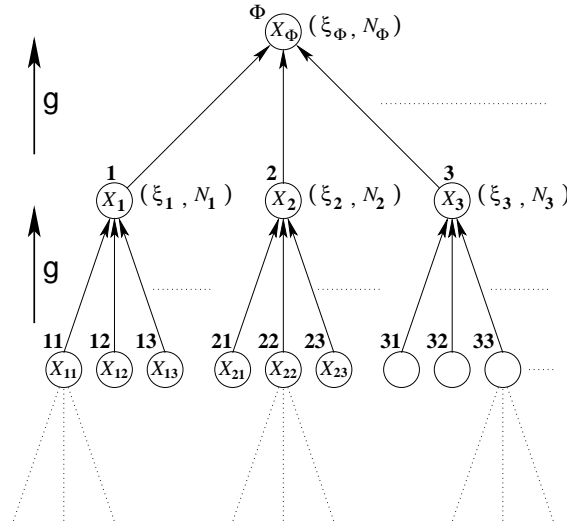
where T is the operator associated with the above equation, which depends on the function g and the joint distribution of the pair (ξ, N) , and μ is the (unknown) law of X .

Recursive Tree Framework (RTF)



- **Skeleton** : $\mathbb{T}_\infty := (\mathcal{V}, \mathcal{E})$ is the canonical infinite tree with vertex set $\mathcal{V} := \{\mathbf{i} \mid \mathbf{i} \in \mathbb{N}^d, d \geq 1\} \cup \{\emptyset\}$, and edge set $\mathcal{E} := \{e = (\mathbf{i}, \mathbf{ij}) \mid \mathbf{i} \in \mathcal{V}, j \in \mathbb{N}\}$, and root \emptyset .
- **Innovations** : Collection of **i.i.d** pairs $\{(\xi_{\mathbf{i}}, N_{\mathbf{i}}) \mid \mathbf{i} \in \mathcal{V}\}$.
- **Function** : The function $g(\cdot)$.

Recursive Tree Process (RTP)



Consider a **RTF** and let μ be a solution of the associated **RDE**. A collection of S -valued random variables $(X_i)_{i \in \mathcal{V}}$ is called an invariant *Recursive Tree Process (RTP)* with marginal μ if

- $X_i \sim \mu \quad \forall i \in \mathcal{V}$.
- Fix $d \geq 0$ then $(X_i)_{|i|=d}$ are independent.
- $X_i = g(\xi_i; X_{ij}, 1 \leq j \leq N_i)$ a.s. $\forall i \in \mathcal{V}$.
- X_i is independent of $\{(\xi_{i'}, N_{i'}) \mid |i'| < |i|\}$ $\forall i \in \mathcal{V}$.

Remark : Using *Kolmogorov's consistency*, an invariant RTP with marginal μ exists if and only if μ is a solution of the associated RDE.

Questions

Given a RDE

$$X \stackrel{d}{=} g\left(\xi; \left(X_j, 1 \leq j \leq^* N\right)\right) \quad \text{on } S$$

one can ask several questions, such as ...

- (i) Does it have a solution ? (*existence*)
- (ii) If yes, is it unique ? (*uniqueness*)
- (iii) If μ is a solution then for what other measure say ν , do we have $T^n(\nu) \xrightarrow{d} \mu$? (*domain of attraction*)

- (iv) If μ is a solution then is the invariant RTP with marginal μ endogenous ? (*endogeny*)

- (v) If μ is a solution then does the invariant RTP with marginal μ has trivial tail ? (*tail triviality of RTP*)
[will not discuss this issue in this series]

... perhaps many more !

Endogeny

Definition 2 *Let \mathcal{G} be the σ -field generated by the innovation process $\{(\xi_{\mathbf{i}}, N_{\mathbf{i}}) \mid \mathbf{i} \in \mathcal{V}\}$. We will say an invariant RTP is endogenous if X_{\emptyset} is almost surely \mathcal{G} -measurable.*

Endogeny vs Uniqueness

Example (Uniqueness $\not\Rightarrow$ Endogeny) :

$$X \stackrel{d}{=} \text{sign}(X_1) \times \xi \quad \text{on } \mathbb{R},$$

where $\xi \sim \text{Normal}(0, 1)$ and is independent of X_1 .

- Trivially the unique solution is Normal(0, 1).
- Once again invariant RTP can be indexed by non-negative integers. Let $(X_i)_{i \geq 0}$ be invariant RTP with marginal Normal(0, 1), from definition

$$X_i = \text{sign}(X_{i+1}) \times \xi_i \quad \text{a.s.}$$

Let $Y_i = -X_i$, $i \geq 0$ then

$$Y_i = \text{sign}(Y_{i+1}) \times \xi_i \quad \text{a.s.}$$

Thus $(Y_i)_{i \geq 0}$ is also an invariant RTP with marginal Normal(0, 1). So RTP $(X_i)_{i \geq 0}$ can not be endogenous.

Remarks :

- We will see that *uniqueness* of a “modified” RDE will imply endogeny.
- Later we will see that a RDE may have many solutions some endogenous while others are not.

Endogeneity vs The Operator T

RDE - I	RDE- II
$X \stackrel{d}{=} \xi$ on \mathbb{R} , where $\xi \sim \text{Normal}(0, 1)$	$X \stackrel{d}{=} \text{sign}(X_1) \times \xi$ on \mathbb{R} , where $\xi \sim \text{Normal}(0, 1)$

Observations :

- Both RDEs define the same operator T , namely mapping every probability measure to Normal (0, 1).
- Trivially the invariant RTP for **RDE - I** is endogenous, while that for **RDE - II** is not.
- Note T is a “very nice” function ! (It is continuous, monotone, a contraction ...)

Remark : In order to answer the question of endogeneity, we need the whole structure of the RTF. In particular the function g . We will later see [Lecture-III] that if g has some “nice” properties then endogeneity will follow for certain solution(s).

Bivariate Uniqueness

Consider the following **bivariate RDE**,

$$\begin{pmatrix} X \\ Y \end{pmatrix} \stackrel{d}{=} \begin{pmatrix} g(\xi; (X_j, 1 \leq j \leq^* N)) \\ g(\xi; (Y_j, 1 \leq j \leq^* N)) \end{pmatrix}$$

where $(X_j, Y_j)_{j \geq 1}$ are i.i.d and has the same law as of (X, Y) , and are independent of the innovation (ξ, N) .

Note : We will denote the operator associated with this bivariate RDE by $T^{(2)}$.

Definition 3 *An invariant RTP with marginal μ has **bivariate uniqueness** property if the above bivariate RDE has unique solution as $X = Y$ a.s on the space of joint probabilities with both marginals μ .*

Uniqueness vs Bivariate Uniqueness

- It is possible to have an RDE which has unique solution but the solution fail to have bivariate uniqueness.

$$X \stackrel{d}{=} \xi + X_1 \pmod{2} \quad \text{on } \{0, 1\},$$

where $\xi \sim \text{Bernoulli}(p)$ for some $0 < p \leq 1$ and is independent of X_1 .

- ▶ Easy to see that the unique solution is $X \sim \text{Bernoulli}(\frac{1}{2})$.
- ▶ It is also not hard to see that, any joint distribution (X, Y) with marginal $\text{Bernoulli}(\frac{1}{2})$, is a solution of the bivariate RDE.

Remark : Later we will have an example (*Quicksort RDE*) where the RDE has many solutions among which only one has bivariate uniqueness.

Equivalence Theorem

Theorem 1 *Suppose the S is a Polish space. Consider an invariant RTP with marginal distribution μ .*

(a) *If the endogenous property holds then the bivariate uniqueness property holds.*

(b) *Conversely, suppose the bivariate uniqueness property holds. If also $T^{(2)}$ is continuous with respect to the weak convergence on the set of bivariate distributions with marginal μ , then the endogenous property holds.*

(c) *Further, the endogenous property holds if and only if*

$$T^{(2)n}(\mu \otimes \mu) \xrightarrow{d} \mu^{\nearrow},$$

where $\mu \otimes \mu$ is the product measure, and μ^{\nearrow} is the measure concentrated on the diagonal with both marginal μ .

Proof of the Equivalence Theorem

Part (a) (Endogeny \Rightarrow Bivariate Uniqueness) :

$$\begin{pmatrix} X \\ Y \end{pmatrix} \stackrel{d}{=} \begin{pmatrix} g(\xi; (X_j, 1 \leq j \leq^* N)) \\ g(\xi; (Y_j, 1 \leq j \leq^* N)) \end{pmatrix}$$

- ν be a solution of the bivariate RDE with marginal μ .
- Using i.i.d. innovations $(\xi_i, N_i)_{i \in \mathcal{V}}$ construct bivariate RTP $((X_i, Y_i))_{i \in \mathcal{V}}$ such that $(X_i, Y_i) \sim \nu$ for all $i \in \mathcal{V}$.

- Observe for every $n \geq 0$,

$$(X_\emptyset; ((\xi_{\mathbf{i}}, N_{\mathbf{i}}), |\mathbf{i}| \leq n)) \stackrel{d}{=} (Y_\emptyset; ((\xi_{\mathbf{i}}, N_{\mathbf{i}}), |\mathbf{i}| \leq n)).$$

- So if $\Lambda : S \rightarrow \mathbb{R}$ is a bounded measurable function then

$$\mathbf{E} [\Lambda (X_\emptyset) \mid \mathcal{G}_n] = \mathbf{E} [\Lambda (Y_\emptyset) \mid \mathcal{G}_n] \quad \text{a.s.},$$

where \mathcal{G}_n is the σ -field generated by $\{(\xi_{\mathbf{i}}, N_{\mathbf{i}}) \mid |\mathbf{i}| \leq n\}$.

- By martingale convergence theorem and using the endogeneity property we get

$$\Lambda (X_\emptyset) = \Lambda (Y_\emptyset) \quad \text{a.s.}$$

This is true for any bounded measurable function Λ , so

$$X_\emptyset = Y_\emptyset \quad \text{a.s.}$$

Part (b) : Bivariate Uniqueness \Rightarrow Endogeny :

$$\begin{pmatrix} X \\ Y \end{pmatrix} \stackrel{d}{=} \begin{pmatrix} g(\xi; (X_j, 1 \leq j \leq^* N)) \\ g(\xi; (Y_j, 1 \leq j \leq^* N)) \end{pmatrix}$$

- Fix $\Lambda : S \rightarrow \mathbb{R}$, a bounded continuous function.
- Let $(X_i)_{i \in \mathcal{V}}$ be an invariant RTP with marginal μ .
- By martingale convergence theorem

$$\mathbf{E} [\Lambda (X_\emptyset) \mid \mathcal{G}_n] \xrightarrow[\mathcal{L}_2]{\text{a.s.}} \mathbf{E} [\Lambda (X_\emptyset) \mid \mathcal{G}],$$

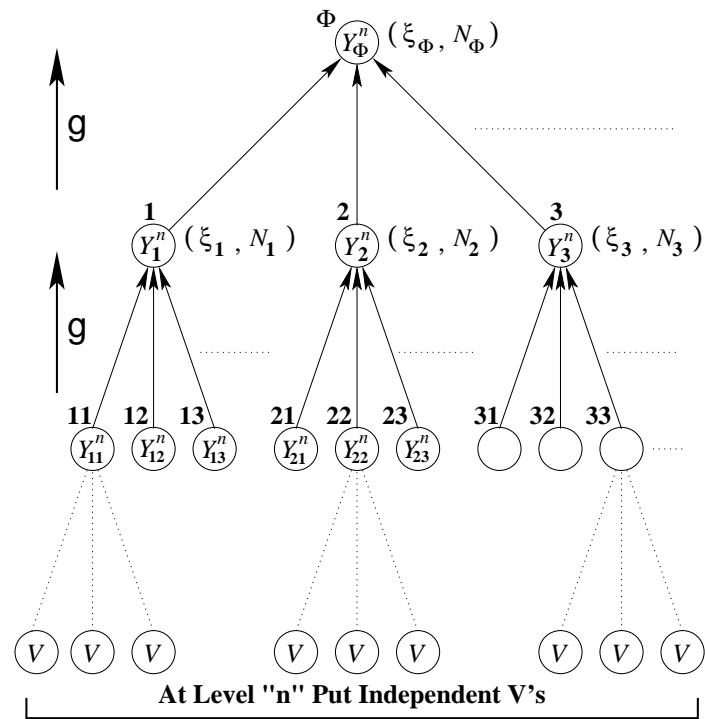
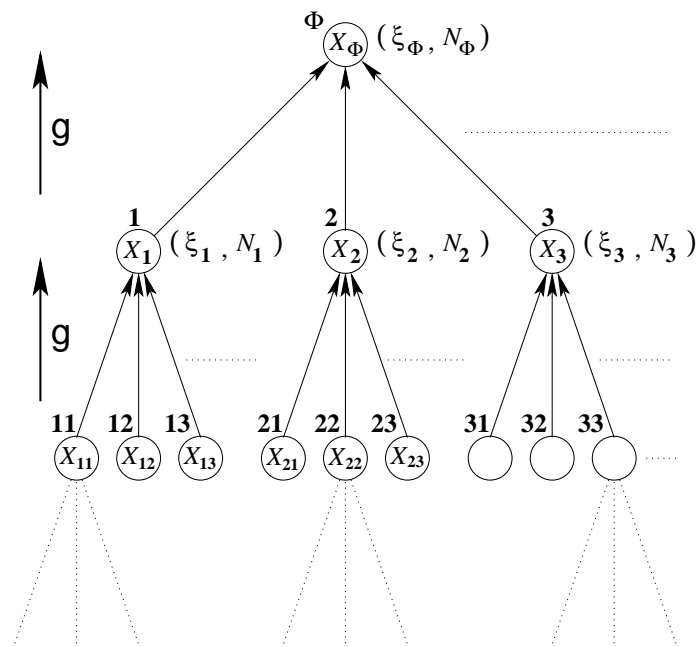
where \mathcal{G}_n is the σ -field generated by $\{(\xi_i, N_i) \mid |\mathbf{i}| \leq n\}$.

- For each $n \geq 0$ we will construct Y_\emptyset^n as follows :
 - ▶ Take $(V_i)_{i \in \mathcal{V}}$ i.i.d. sample from μ which are independent of the innovations $(\xi_i, N_i)_{i \in \mathcal{V}}$ as well as $(X_i)_{i \in \mathcal{V}}$.
 - ▶ Fix $n \geq 0$ and define $Y_i^n = V_i$ for $|\mathbf{i}| = n$.
 - ▶ For $|\mathbf{i}| < n$ define Y_i^n recursively.
- Some immediate consequence of the construction are as follows :
 - ▶ $X_\emptyset \stackrel{d}{=} Y_\emptyset^n \stackrel{d}{=} \mu$ for every $n \geq 0$.
 - ▶ For each $n \geq 0$, the random variables X_\emptyset and Y_\emptyset^n when conditioned on the σ -algebra \mathcal{G}_n , are independent and identically distributed.
 - ▶ Moreover for each $n \geq 0$,

$$\begin{bmatrix} X_\emptyset \\ Y_\emptyset^{n+1} \end{bmatrix} \stackrel{d}{=} T^{(2)} \left(\text{dist} \left(\begin{bmatrix} X_\emptyset \\ Y_\emptyset^n \end{bmatrix} \right) \right).$$

- ▶ Finally we also note that

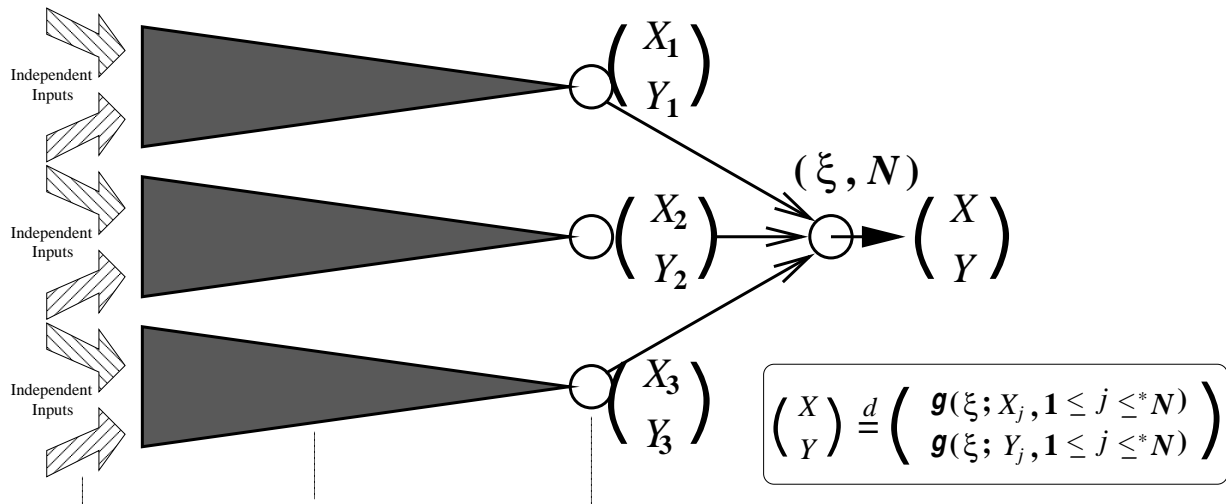
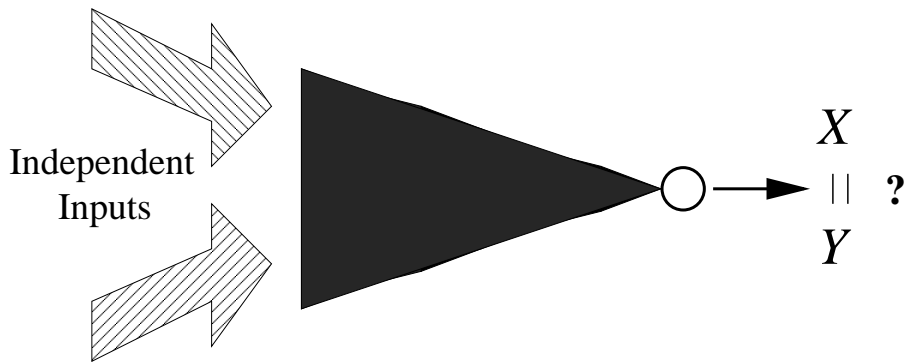
$$\begin{bmatrix} X_\emptyset \\ Y_\emptyset^n \end{bmatrix} \stackrel{d}{=} T^{(2)^n} (\mu \otimes \mu).$$



Input at Infinity

RTF

Output



- Consider the sequence $(X_\emptyset, Y_\emptyset^n)_{n \geq 0}$.
- It is tight because all the marginals are same which is μ .
- If $f, h : S \rightarrow \mathbb{R}$ be two bounded continuous functions then

$$\begin{aligned}
 \mathbf{E} [f (X_\emptyset) h (Y_\emptyset^n)] &= \mathbf{E} [\mathbf{E} [f (X_\emptyset) h (Y_\emptyset^n) \mid \mathcal{G}_n]] \\
 &= \mathbf{E} [\mathbf{E} [f (X_\emptyset) \mid \mathcal{G}_n] \mathbf{E} [h (X_\emptyset) \mid \mathcal{G}_n]] \\
 &\rightarrow \mathbf{E} [\mathbf{E} [f (X_\emptyset) \mid \mathcal{G}] \mathbf{E} [h (X_\emptyset) \mid \mathcal{G}]]
 \end{aligned}$$

- So we conclude that

$$(X_\emptyset, Y_\emptyset^n) \xrightarrow{d} (X^*, Y^*),$$

for some (X^*, Y^*) which has marginal μ .

- From the (technical) continuity assumption of $T^{(2)}$ we get that (X^*, Y^*) is a solution of the bivariate equation with marginal μ , so using bivariate uniqueness we get

$$X^* = Y^* \text{ a.s.}$$

- Let $\sigma_n^2(\Lambda) := \|\mathbf{E}[\Lambda(X_\emptyset) | \mathcal{G}_n] - \Lambda(X_\emptyset)\|_2^2$.
- Easy calculation shows

$$\begin{aligned}\sigma_n^2(\Lambda) &= \mathbf{E}[\text{Var}(\Lambda(X_\emptyset) | \mathcal{G}_n)] \\ &= \frac{1}{2}\mathbf{E}\left[(\Lambda(X_\emptyset) - \Lambda(Y_\emptyset^n))^2\right]\end{aligned}$$

The last inequality follows from the simple fact that for any random variable U with finite second moment,

$$\text{Var}(U) = \frac{1}{2}\mathbf{E}[(U_1 - U_2)^2]$$

where (U_1, U_2) are two independent copies of U .

- Taking $n \rightarrow \infty$ limit we can then conclude that $\sigma_n^2(\Lambda) \rightarrow 0$, because $(X_\emptyset, Y_\emptyset^n) \xrightarrow{d} (X^*, X^*)$.
- So $\Lambda(X_\emptyset) = \mathbf{E}[\Lambda(X_\emptyset) | \mathcal{G}]$ a.s.
- This is true for every bounded continuous function Λ , so we conclude that X_\emptyset is a.s. \mathcal{G} measurable, proving the endogeneity property.

Part (c) [“if”-part] :

- We know that the construction of $(Y_i^n)_{|i| \leq n}$ yields $(X_\emptyset, Y_\emptyset^n)$ has distribution $T^{(2)^n}(\mu \otimes \mu)$.

- So we get

$$(X_\emptyset, Y_\emptyset^n) \xrightarrow{d} \mu^{\nearrow}.$$

- The rest follows from the previous argument.

Part (c) [“only if”-part] :

- Again work with the same construction of $(Y_\emptyset^n)_{n \geq 0}$.
- Let $\Lambda_1, \Lambda_2 : S \rightarrow \mathbb{R}$ be two bounded continuous functions.

$$\begin{aligned} & \mathbf{E} [\Lambda_1 (X_\emptyset) \Lambda_2 (Y_\emptyset^n)] \\ &= \mathbf{E} [\mathbf{E} [\Lambda_1 (X_\emptyset) \Lambda_2 (Y_\emptyset^n) \mid \mathcal{G}_n]] \\ &= \mathbf{E} [\mathbf{E} [\Lambda_1 (X_\emptyset) \mid \mathcal{G}_n] \mathbf{E} [\Lambda_2 (Y_\emptyset^n) \mid \mathcal{G}_n]] \\ &\rightarrow \mathbf{E} [\mathbf{E} [\Lambda_1 (X_\emptyset) \mid \mathcal{G}] \mathbf{E} [\Lambda_2 (Y_\emptyset) \mid \mathcal{G}]] \\ &= \mathbf{E} [\Lambda_1 (X_\emptyset) \Lambda_2 (Y_\emptyset)] \end{aligned}$$

The last equality follows from endogeneity assumption.

- This of course then implies

$$(X_\emptyset, Y_\emptyset^n) \xrightarrow{d} (X_\emptyset, Y_\emptyset),$$

which is same as saying

$$T^{(2)n} (\mu \otimes \mu) \xrightarrow{d} \mu^{\nearrow}.$$

Application to Solutions of the Quicksort RDE

Recall that the *Quicksort RDE* is given by

$$X \stackrel{d}{=} UX_1 + (1 - U)X_2 + C(U) \quad \text{on } \mathbb{R},$$

where (X_1, X_2) are i.i.d. copies of X and are independent of $U \sim \text{Uniform}[0, 1]$, and

$$C(u) := 1 + 2u \log u + 2(1 - u) \log(1 - u).$$

Known :

- If X is a solution then so is $(m + X)$ for any $m \in \mathbb{R}$.
- There is a unique solution with $\mathbf{E}[X] = 0$ and $\mathbf{E}[X^2] < \infty$ [Rösler (1992)].
- Let ν be the solution with mean zero and finite variance then the set of all solutions is given by

$$\{\nu * \text{Cauchy}(m, \sigma^2) \mid m \in \mathbb{R}, \sigma^2 \in \mathbb{R}_+\}$$

[Fill and Janson (2000)]

- Note that the only mean zero solution is ν .

Theorem 2 *A solution of the Quicksort RDE is endogenous if and only if $\sigma^2 = 0$.*

Remark : In other words, the solution ν and its translates are the only endogenous solutions.

Proof of Theorem 2

- We will use the bivariate uniqueness technique.
- Let $\mu = \nu * \text{Cauchy}(m, \sigma^2)$ be a solution of the Quicksort RDE. Consider the bivariate RDE

$$\begin{pmatrix} X \\ Y \end{pmatrix} \stackrel{d}{=} \begin{pmatrix} UX_1 + (1-U)X_2 + C(U) \\ UY_1 + (1-U)Y_2 + C(U) \end{pmatrix},$$

where $(X_j, Y_j)_{j=1,2}$ are i.i.d. copies of (X, Y) and are independent of $U \sim \text{Uniform}[0, 1]$. Further assume $X \stackrel{d}{=} Y \stackrel{d}{=} \mu$.

Proof of the “if”-part

$$\begin{pmatrix} X \\ Y \end{pmatrix} \stackrel{d}{=} \begin{pmatrix} UX_1 + (1-U)X_2 + C(U) \\ UY_1 + (1-U)Y_2 + C(U) \end{pmatrix}$$

- We assume $\sigma^2 = 0$.
- Let $D = X - Y$ and similarly define D_1 and D_2 .
- Then $D \stackrel{d}{=} UD_1 + (1-U)D_2$ on \mathbb{R} .
- Since $\sigma^2 = 0$, so $X \stackrel{d}{=} Y \stackrel{d}{=} \nu * \delta_m$, thus D has finite second moment.
- Simple calculation then shows $\mathbf{E}[D] = 0 = \mathbf{E}[D^2]$.
- Thus $X = Y$ a.s., that is, bivariate uniqueness holds.

Proof of the “only if”-part

$$\begin{pmatrix} X \\ Y \end{pmatrix} \stackrel{d}{=} \begin{pmatrix} UX_1 + (1 - U)X_2 + C(U) \\ UY_1 + (1 - U)Y_2 + C(U) \end{pmatrix}$$

- Suppose $\sigma^2 > 0$.
- We will show that $(Q + Z, Q + W)$ is a solution of the bivariate equation, where Z and W are i.i.d. Cauchy (m, σ^2) and are independent of $Q \sim \nu$.
- Observe that if Z_1 and Z_2 are i.i.d. Cauchy (m, σ^2) and are independent of $U \sim \text{Uniform}[0, 1]$ then

$$Z = UZ_1 + (1 - U)Z_2$$

is also Cauchy (m, σ^2) and it is independent of U (follows by computing the characteristic function).

- Take $(Z_1, Z_2; W_1, W_2)$ i.i.d. Cauchy (m, σ^2) ; (Q_1, Q_2) i.i.d. copies of $Q \sim \nu$; and $U \sim \text{Uniform}[0, 1]$. All are independent.
- Define $X_j := Q_j + Z_j$ and $Y_j := Q_j + W_j$, $j \in \{1, 2\}$.
- Let $Q := UQ_1 + (1 - U)Q_2 + C(U)$ then $Q \sim \nu$.
- If $Z := UZ_1 + (1 - U)Z_2$ and $W := UW_1 + (1 - U)W_2$ then

$$\begin{aligned} Q + Z &= UX_1 + (1 - U)X_2 + C(U) \\ Q + W &= UY_1 + (1 - U)Y_2 + C(U) \end{aligned}$$
- But Z and W are i.i.d. Cauchy (m, σ^2) and are independent of Q .
- Thus $(Q + Z, Q + W)$ is a non-trivial solution of the bivariate RDE and hence bivariate uniqueness fails.