

# Statistics - 134 ( Lecture - 2 ), Fall 2002

## Solutions of the Final Exam Problems

1. (a) **TRUE.**

$$\mathbf{P}(A|B) = \frac{\mathbf{P}(A \cap B)}{\mathbf{P}(B)} = \frac{\mathbf{P}(A)}{\mathbf{P}(B)} \geq \mathbf{P}(A).$$

- (b) **FALSE.** Notice that  $\mathbf{E}[X - Y] = 0$  and  $\mathbf{Var}(X - Y) = 2$  so by Chebyshev inequality we get

$$\mathbf{P}(|X - Y| > 2) \leq \frac{1}{2}$$

- (c) **FALSE.** Certainly,  $\mathbf{E}\left[\frac{X}{Y}\right] = \mathbf{E}[X] \mathbf{E}\left[\frac{1}{Y}\right]$ , because  $X$  and  $Y$  are independent. But in general,  $\mathbf{E}\left[\frac{1}{Y}\right] \neq \frac{1}{\mathbf{E}[Y]}$ .
- (d) **TRUE.** Observe that  $\mathbf{Var}(X) = \mathbf{E}[X^2] - (\mathbf{E}[X])^2 \geq 0$ .
- (e) **TRUE.** By convolution formula,  $X - Y$  has a density and hence it is a continuous random variable. So  $\mathbf{P}(X = Y) = \mathbf{P}(X - Y = 0) = 0$ .

2. Apply the change of variable formula to get ( note that the transformation is  $y = e^x$ , which is one-to-one, with differentiable inverse )

$$f_Y(y) = \begin{cases} \frac{1}{\sqrt{2\pi y}} e^{-\frac{1}{2}(\log y)^2} & \text{if } y > 0, \\ 0 & \text{otherwise.} \end{cases}$$

3. (a)

$$\begin{aligned} \mathbf{P}(T_3 > 3) &= \mathbf{P}(N_3 \leq 2) \\ &= e^{-9} \left( 1 + 9 + \frac{9^2}{2!} \right) \\ &\approx 0.0062 \end{aligned}$$

- (b)

$$\begin{aligned} \mathbf{P}(T_6 - T_3 \leq 2) &= \mathbf{P}(W_4 + W_5 + W_6 \leq 2) \\ &= \mathbf{P}(T_3 \leq 2) \\ &= \mathbf{P}(N_2 \geq 3) \\ &= 1 - e^{-6} \left( 1 + 6 + \frac{6^2}{2!} \right) \\ &\approx 0.9380 \end{aligned}$$

4. (a) The joint density of  $(X, Y)$  is given by

$$f(x, y) = \frac{1}{2\pi} e^{-\frac{1}{2}[x^2 + (x-y)^2]} \quad -\infty < x, y < \infty.$$

So to compute the marginal density of  $Y$  we need to integrate  $f(x, y)$  over  $x$ . Now, notice that it is then nothing but a convolution of two  $\text{Normal}(0, 1)$  variables, and hence  $Y \sim \text{Normal}(0, 2)$ .

- (b) Simple calculation will show that given  $[Y = y]$  the random variable  $X \sim \text{Normal}\left(\frac{y}{2}, \frac{1}{2}\right)$ . Thus,  $\mathbf{E}[X|Y = y] = \frac{y}{2}$ .

5. Let  $Z$  be the total profit after 100 draws, then  $Z = X_1 + X_2 + \dots + X_{100}$ , where  $X_1, X_2, \dots, X_{100}$  are **i.i.d.** such that  $\mathbf{P}(X_1 = 2) = \mathbf{P}(X_1 = -1) = 3/10$ ,  $\mathbf{P}(X_1 = 0) = 4/10$ . Use **Normal approximation** method to conclude that  $\mathbf{P}(Z > 45) \approx 0.1038$ .
6. (a) Note that  $X$  takes values in  $(0, 1)$ . Fix  $0 < x < 1$ , then

$$f_X(x) = \int_0^1 f(x, y) dy = \frac{1}{5} (3x^2 + 4x + 2).$$

(b)

$$\begin{aligned} \mathbf{P}(X \leq Y) &= \int_{x \leq y} \int f(x, y) dx dy \\ &= \frac{1}{5} \int_0^1 \left( \int_0^y (3x^2 + 4xy + 6y^2 + 2x) dx \right) dy \\ &= \frac{1}{5} \int_0^1 (9y^3 + y^2) dy \\ &= \frac{31}{60} \end{aligned}$$

7. (a) The values of  $T$  are  $\{1, 2, 3, \dots\}$ . Fix  $k \geq 1$ , then

$$\mathbf{P}(T = k) = \frac{2}{3} \left(\frac{1}{2}\right)^k + \frac{1}{3} \left(\frac{1}{2} + \theta\right) \left(\frac{1}{2} - \theta\right)^{k-1}.$$

(b)

$$\mathbf{P}(\text{I got the biased coin} | T = 10) = \frac{\frac{1}{3} \left(\frac{1}{2} + \theta\right) \left(\frac{1}{2} - \theta\right)^9}{\frac{2}{3} \left(\frac{1}{2}\right)^{10} + \frac{1}{3} \left(\frac{1}{2} + \theta\right) \left(\frac{1}{2} - \theta\right)^9}$$

8. (a) The values of  $Y$  are  $\{0, 1, 2, \dots\}$ .  
 (b) Fix  $k \geq 0$ , then

$$\begin{aligned} \mathbf{P}(Y = k) &= \mathbf{P}(k \leq X < k + 1) \\ &= e^{-k} - e^{-(k+1)} \\ &= (1 - e^{-1}) e^{-k}. \end{aligned}$$

(c) So  $Y$  follows Geometric distribution supported on  $\{0, 1, 2, \dots\}$ , and with success probability  $(1 - e^{-1})$ .

9. Let  $X$  be the time when Hermione arrives and  $Y$  be the time when Harry arrives. From the given information we get that  $X$  and  $Y$  are independent normal random variables with  $X - 12 : 00 \text{ noon} \sim \text{Normal}(-5, 9)$  and  $Y - 12 : 00 \text{ noon} \sim \text{Normal}(5, 9)$ .

(a)  $\mathbf{P}(Y < X) = \mathbf{P}(X - Y > 0) = 1 - \Phi\left(\frac{10}{\sqrt{18}}\right) \approx 0.0091$ .

(b)  $\mathbf{P}(Y - X > 10) = 0.5$ .

10. Notice that  $X \sim \text{Exponential}(\lambda)$ .

(a)  $\mathbf{E}[X^n]$  can now be computed by looking at the form of Gamma  $(n + 1, \lambda)$  density, and it is  $n!/\lambda^n$ .

(b) By geometric series formula the required sum is  $\lambda/(\lambda - z)$ .