Statistical Discrimination at the Extensive & Intensive Margin: Explaining Racial Employment & Wage Gap

Md Moshi Ul Alam $^*\!\mathrm{and}$ Irina Shaorshadze †

September 15, 2018

Abstract

This paper provides evidence, the first of its kind on the existence of statistical discrimination in employment decisions (the extensive margin) in US labor market. The literature has usually used AFQT as the measure of productivity observable to the resarcher but unobservable to the employer. We extend this set by incorporating data on non-cognitive skills from the NLSY 79 and 97. Unlike existing literature which finds no evidence of statistical discrimination in wages,- after incorporating non-cognitive skills, our model provides evidence on statistical discrimination in wages (intensive margin). Our results also reverses those of learning models of productivity which finds no evidence of statistical discrimination in the long run.

1 Introduction

Most studies of racial differences in labor market outcomes focus on differences in wages. However differences in employment are even more dramatic than differences in wages. Black workers have higher unemployment duration and job turnover than whites. This is

^{*}Department of Economics, University of Wisconsin-Madison

[†]The Bank of America

true even if we do not include the incarcerated, and is true for both men and women. It has been suggested that racial discrimination is important determinant of labor market outcomes, however most theoretical and empirical work on discrimination studies the effect of discrimination on wages, and not on unemployment. This paper analyses the effect of screening statistical discrimination on unemployment duration and job turnover gap between black and white workers in labor markets characterized search frictions.

Since unemployment differences between blacks and whites is even more dramatic than that of wages, a story of the inter-group labor market outcomes is not complete without grappling with the puzzle of differential unemployment outcomes. While frictionless world could explain unemployment through particularly high taste for leisure or high outside option, the magnitude of differences in unemployment between black and white men suggest that leisure-based explanations is at best a part of the story. The alternative way to explain the differential employment outcome of black and white workers is through search frictions combined with discrimination. In the current version of the paper we provide reduced for evidence for the existence of the unemployment gap between black and white workers. This evidence will serve as a motivation for the structural model that will use combination of search and employer learning to explain the unemployment gap between black and white workers.

1.1 Stylized Facts

The following stylized facts are from the survey by Lang and Lehmann (2012). We focus here on black and white men.

- There is a raw wage gap between blacks and whites. This gap almost disappears for the high skilled workers. The gap disappears if one controls for AFQT, but it re-emerges if one controls both for AFQT and education. This is because for a given AFQT blacks tend to get more years of schooling.
- From 2003 to 2008, the ratio of mean incomplete unemployment duration of black

men 16 and older relative to white men was between 1.28 and 1.33.

- Overall unemployment duration is about 30% longer for black men.
- There are clear differences in rates of entry into unemployment from employment.
 Mean employment duration of black men is 70% of that of white men.
- Unemployment and skills are negatively correlated. However while during last 30% years there was convergence in earnings, there was no convergence in unemployment.

2 Literature

There are two strands of models of discrimination. Taste-based discrimination explains inter-group labor-market differences through taste (prejudice) of employers or their customers. Canonical model of staistical discrimination is by Becker (2010). Meanwhile statistical discrimination does not rely on prejudice to explain the labor-market outcomes. Instead this literature focuses on imperfect information on productivity of the workers Phelps (1972), Arrow (1998). There are two types of models of statistical discrimination. Models of rational stereotyping assume that groups are ex-ante equally productive, however negative beliefs of employers about the group productivity are ex-post self-fulfilling . Screening discrimination models admit that there may be underlying differences productivity between the groups, but any given worker's actual productivity is imperfectly observed. Most of the studies of discrimination focus on wage differences.

To date there is a small literature that combines taste-based discrimination models with models of search. Black (1995) and Bowlus and Eckstein (2002) develop the model that combines random search and taste discrimination. Lang et al. (2005) present the model that combines taste based discrimination with directed search. Random search models can reproduce racial gaps in both unemployment and wages. The unsatisfactory feature of such models is that they rely on presence of a lot of employers that are strongly prejudiced. It is also unsatisfactory that black workers are not able to avoid the employers that are discriminatory. Meanwhile, models of directed search (posted wages) produce counterfactual results that unemployment duration for blacks is shorter than that of whites.

Rosen (1997) develops search and matching model that does not rely on prejudice, but rather on private information about match quality. She shows that if whites have higher reservation productivity, the firms will know that the white applicants are more productive, and will only choose white applicants. While this model has elements reminiscent of statistical discrimination, the underlying factor in having blacks have lower reservation productivity and not the precision of signals. Therefore the underlying mechanism is in essence similar to low outside option-type story. Meanwhile the model is able to produce both plausible wage differentials and unemployment differentials.

Several papers have sought investigate the presence of statistical discrimination by postulating that there is a proxi for unobservable worker ability (such as AFQT test scores), and by looking if AFQT is revealed to employers overtime and thus reflected in wages. In a seminal paper that employes this approach to investigate presence of statistical discrimination against blacks and employer learning, Altonji and Pierret (2001) do not find evidence of statistical discrimination on wages. Meanwhile, using similar methods, Arcidiacono et al. (2010) find that the AFQT is revealed upon hiring for college graduate workers, but black high school graduates do experience statistical discrimination, as their AFQT scores are not reflected in wages.

Recent survey by Lang and Lehmann (2012) suggest that search models should be important part of explaining racial unemployment puzzle, and they suggest that combining them with statistical discrimination model would be an interesting avenue to pursue. That statistical discrimination can fruitfully applied to investigating racial differencials in hiring and unemployment were suggested by Altonji (2005) and Arrow (1998). However very few papers incorporate statistical discrimination in explaining unemployment.

Altonji (2005) sketches out a model that includes statistical discrimination, hiring and

minimum wages, but does not take the model to the data. Ahn et al. (2010) present the two-sided one-shot matching model that includes black and white workers, different ability levels, and noisy signals of the abilities. The authors find that the employers find it easier to screen based on race rather than age. As a result 18 year old blacks earn as much as 16 year old whites, but have much higher unemployment rates. Cavounidis and Lang (2015) develop the model that includes search and screening. In this model, black workers are monitored more once hired, and are fired more often than whites as a result. Because of differencial churning of white and black workers, a typical unemployed black worker is of lower quality than a typical unemployed white worker. Therefore in equilibrium unemploymet duration is longer for the black workers, while turnover is higher. However the model is stylized and the authors do not take it to the data. Fryer Jr et al. (2013) look at the evidence consistent with statistical discrimination in the data from Unemployment Insurance recipients in New Jersey. They find the supporting evidence for statistical discrimination in hiring.

The issue with learning models is that uncertainty often resolves quickly Lange (2007). However this does not have to be the case if search frictions are present. It is reasonable to start from model such as that of Rosen (1997), flash out stronger foundations based on statistical screening discrimination, incorporate learning about the productivity over time and take it to the data.

3 Data

This study uses NLSY79 data for the years 1979-2012. The selection criteria for the sample follows that of Arcidiacono et al. (2010). We keep the data for the white and black men. Actual experience is the weeks worked divided by 50, while the actual experience is the years since the respondent left school. Wages and earnings are not deflated.

Consistent with the literature (Altonji and Pierret, 2001), we take AFQT scores to

measure worker's ability.

4 Descriptive Evidence

In this section we review the descriptive evidence from NLSY79 on the differences in labor market experience between black and white workers.

Figure 1a shows the unconditional density of AFQT score for black and white workers. It is evidence that distribution of AFQT scores between black and white workers is very different. While AFQT density appears uniform across most of the AFQT range, AFQT scores for the black workers are concentrated at the lower tail. figures 1b, 1c, 1d show the distribution of AFQT scores by education levels. These figures confirm that for a given AFQT score, black workers get more education (Lang and Lehmann, 2012). It appears that out of there four education groups, high school dropout blacks and whites are relatively more similar in terms of AFQT scores. This suggests that in order to have the black and white workers similar in terms of AFQT, analysis might best consider low skill labor market of high school dropouts.

Figure 3 shows the transition rates between employment and unemployment. In order to generate these graphs we follow Arcidiacono et al. (2010) and consider that individuals enter the labor market in week 27th of the year of their highest completed education level. Quarterly transition rate out of unemployment (conditional on remaining in the labor force) is almost 0.15 higher for whites than for blacks at the time of labor market entry. The gap narrows only after a decade from the entry into the labor market. However even 20 years after the entry into the labor market, the 0.1 gap still remains.

Figure 3b shows that at the time of labor market entry there is close to 0.1 gap in the transition from employment into unemployment. However this gap narrows after around 5 years. This is consistent with the statistical discrimination story, where information about worker's skills is noisy when workers enter the labor market, and it is relatively more

noisy for black workers. After the workers get hired, their productivity is revealed. Those whose productivity is lower than some threshold, get terminated. If the productivity signal of black workers is more noisy, such termination will be more frequent for black workers in the beginning of their careers, before their true productivity is revealed. It would be interesting to look at the transition rates by the level of education.

Figure 4 shows the age profile of wages by education levels. In order to generate the weekly wages, we divide the yearly labor earnings by weeks employed in a year. These graphs indicate that at the age 21, there is no gap in wages between the two races. However the gap emerges and widens through next 2 decades of life. It is important to note that actual labor market experience acquired is lower for blacks than whites for all education levels. This accumulated gap in experience likely plays role in explaining the diverging wage gap.

Figure 5 shows the weeks employed in a given year. the gap in weeks employed between the two races appears remarkably consistent during till the age of 40.

Next we investigate the reasons the employment spell ended. The motivation for this is to find evidence for for asymmetric information on worker productivity between workers and employers. For instance, if workers knew their own productivity when starting employment better than employers, then if the spell ended, this maybe more likely to be because due to being fired. If there is difference in the information about worker productivity that employer have between black and white workers, this difference may manifest in different patterns of employment termination. However Table 23 does not reveal any systematic differences in reasons the employment spell ended.

Table 24 shows the duration of the unemployment spells by the level of education. Blacks have lower duration of unemployment for low levels of education. On the other hand, for individuals with doctoral or professional degrees the duration unemployment for blacks is lower. However table 26, shows that there are only 4 black individuals in this category. Table 25 shows the mean duration of employment spells by education. Mean employment spell duration is lower for black workers for all education levels other than professional and doctoral.

5 Reduced Form Evidence

5.1 Black and White Differences in Wages

I start by looking at the evidence of black and white wage gap. Following Altonji and Pierret (2001) and Arcidiacono et al. (2010), we formulate the simple econometric model of employer learning we estimate the log wage equation

$$w_{i} = \beta_{0} + \beta_{1}r_{i} + \beta_{AFQT}AFQT_{i} + \beta_{r,x}(r_{i} \times x_{i}) + \beta_{AFQT,x}(AFQT_{i} \times x_{i}) + \beta_{AFQT,r}(AFQT_{i} \times r_{i}) + \beta_{AFQT,r,x}(AFQT_{i} \times r \times x_{i}) + f(x_{i}) + \beta_{\Phi}'\Phi_{i} + \epsilon_{i}$$

$$(1)$$

In equation 1, r stands for race, x stands for experience, Phi stands for controls. If employers do not observe AFQT at the time of hire, but learn it over time, then coefficient on AFQT should be small initially, but should grow over time. If employers use race to statistically discriminate upon hiring, we should expect the coefficient on *Black* to be negative at the time of hiring.

Dependent variable in Table 1 is the wage at the main job. ¹ we do not see evidence of wage discrimination. After controlling for AFQT, coefficient on *black* is close to zero and not statistically significant. Furthermore, coefficient on AFQT is positive and significant, and there is evidence that there returns to AFQT increase in experience. When potential experience is used rather than actual experience, coefficient on *Black* appears positive, but interaction of black and potential experience is negative. This is due to the fact that black workers spend longer in unemployed and non-participation states.

¹Individual may hold more than one job at the time.

Lack of statistical discrimination in wages is corroborated by the evidence in tables 2 and 3. Dependent variable in Table 2 is the starting wage in the main job, while dependent variable in Table 3 is the first job after graduation.

5.2 Black and White Differences in Unemployment

This section extends the results by Altonji and Pierret (2001) and Arcidiacono et al. (2010) to analyse unemployment. In 2, U stands for either indicator of having unemployment spell or for number of weeks unemployed.

$$U_{i} = \beta_{0} + \beta_{1}r_{i} + \beta_{AFQT}AFQT_{i} + \beta_{r,x}(r_{i} \times x_{i}) + \beta_{AFQT,x}(AFQT_{i} \times x_{i}) + \beta_{AFQT,r}(AFQT_{i} \times r_{i}) + \beta_{AFQT,r,x}(AFQT_{i} \times r \times x_{i}) + f(x_{i}) + \beta_{\Phi}'\Phi_{i} + \epsilon_{i}$$

$$(2)$$

Table 13 estimates equation similar tot that in 1, except the dependent variable is the indicator for existence of unemployment spell is the previous year. Column 1 shows that black workers are 30% more likely to be unemployed in a given year, even if we control of AFQT. Further controlling for experience and interactions of experience and race or AFQT do not get rid of the significant black coefficient.

Table 14 estimates equation where the dependent variable is the indicator if the individual was unemployed at the time of interview. Evidence here is consistent with that in table 13 - coefficient on Black is positive and significant.

Table 19 estimates equation where the dependent variable is the number of weeks unemployed in the year. There is evidence that black workers spend longer time being unemployed, but the gap decreases over time. Table 20 presents the same evidence from a negative binomial regression where the dependent variable is the weeks unemployed in the calendar year.

Wage and unemployment regressions show that there is no gap in wages between black

and white workers. However there is gap in unemployment even when controlling for AFQT.

6 Non-Cognitive Skills

The results discussed above are all in line with AFQT being used as the variable which is a correlate of productivity, being unobservable to the employer but observed to the researcher. AFQT has widely been used in the literature, with its advent from Neal and Johnson (1996) In this section, we discuss results by incorporating non-cognitive skills as the correlates of productivity which affects wages and is unobservable to the employer. Figure 2 shows the distribution of coding speed (one of the measures of non-cognitive skills) by different levels of education across races. The figures reveal a difference in the distribution of this non-cognitive skill by race.

A growing branch of literature document significant returns to both cognitive and noncognitive skills (Heckman (2000), Heckman and Kautz (2012), Kautz et al. (2014), Heckman and Rubinstein (2001), Almlund et al. (2011), Borghans et al. (2008)). Heckman and Rubinstein (2001) document that although in the raw data GED earners earn more than high school dropouts,- hoowever once measured cognitive ability is controlled for, GED earners earn as much or less than high school dropouts. They attribute the difference in non-cognitive skill towards this gap,- GED earners although have better cognitive skills than high school dropouts, but have similar or worse non-cognitive skills than high school dropouts. Heckman et al. (2006), ² document the that higher levels of cognitive and non-cognitive skills are associated with higher probabilities of graduating from high school and higher returns to schooling. The authors go on to show that increasing non-cognitive ability to the highest level reduces the probability of being a high school dropout to virtually zero for females with average cognitive ability and evidence for similar

 $^{^{2}}$ also see Urzua (2006)

patterns are documented in case of college attendance as well. ³ Non-cognitive skills have also been shown to affect measured ability on achievement tests, because non-cognitive skills affect schooling and schooling in turn affects success in achievement tests (Hansen et al. (2004), Heckman et al. (2004)). Cunha et al. (2006) presents a substantial overview of the literature on the importance of non-cognitive skills and its contribution towards economic success (graduation probabilities, returns to education, incarceration, health among others). The literature on early childhood development shows early childhood interventions having most of its effects through bolstering non-cognitive skills (Karoly and Levaux (1998), Blau and Currie (2006), Heckman (2000)).

Tables 9-11, 12-15 and 19-21 show evidence on existence of statistical discrimination at the extensive margin (dummy of being unemployed during the interview, dummy of being unemployed in the last year and number of weeks unemployed respectively) by incorporating different measures of non-cognitive skills like the locus of control measure and the self esteem measure.

In wages, tables 4-6 provide evidence of statistical discrimination in wages, overturning results in the literature which talked about non-existence of statistical discrimination in wages conditional on AFQT or models of learning too. The wage gap due to discrimination is as high as 12% in beginning wages when the employee enters the labor market, it however drops to 5.5% for wages in the main job at the time of survey

7 Description of Proposed Model

The model will explain differential employment outcomes, not just wage dispersion. There are two issues to explain: high unemployment duration and high turnover. Job destruction rate could not just be different between workers because they are blacks and whites

 $^{^{3}\}mathrm{Heckman}$ et al. (2006) also document importance of both cognitive and non-cognitive skills on non-marital pregnancy for females.

- that would be illegal. If anything, firing black workers should be harder because of anti-discrimination laws. So either job destruction should be exogenously heterogeneous among firms and black workers disproportionately sort to such firms, or this destruction should be different in equilibrium for blacks and whites. Intuitively, because black workers have higher uncertainty on their productivity (either by themselves or by the employer or both), that makes them more risk loving. That could make such workers sort to more insecure jobs.

Belowwelist some of the model features

- The model will allow for ex-ante heterogeneity. In other words, it will allow for the possibility that ex-ante the distribution of abilities is different, and that the distribution of whites first order stochastically dominates the distribution of blacks, as in the data. It will take as distribution of human capital at the time of labor market entry as exogenous. Therefore it will explain only the part of discrimination that emerges after human capital investment have been made, but not the lack of human capital investment.
- The model will include only the low wage workers. Low and high skill jobs constitute separate labor markets. In the data the gap in employment and wages for college educated black and white workers is small.
- The model will be base on posted wages and directed search. Model with posted wages is preferable, as it will capture the legal constraint facing the firms that they can not legally condition the wages on the race of the applicant.
- the model will avoid the issues of asymmetric information by assuming that the ability is similarly noisy as viewed by the employer and the worker themselves.
- The model will includes three states (employment, unemployment and nonparticipation), since differences in nonparticipation are very large, and likely reflect the expectations on the condition of the labor market. Nonparticipation may be included as non-market activity, or can be modeled as crime

8 Further Work

In work in progress we build the structural model that will explain black and white unemployment gap and wage gap in terms of statistical discrimination. We plan to use the reduced form estimates as auxiliary parameters, to estimate model parameter using indirect inference. Target moments will be key moments of unemployment rates and wages, and their respective racial differentials. The counterfactual I am interested in is to how much does the wage gap and unemployment gap reduce, once employers have perfect information about the correlates of productivity, i.e. AFQT and non-cognitive skills.

	Linear Re	egression C	pefficients S	Snown		
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.007	0.006	0.004	0.003	0.078^{***}	-0.009
	(0.006)	(0.006)	(0.006)	(0.010)	(0.011)	(0.010)
AFQT	0.205^{***}	0.203^{***}	0.160^{***}	0.160^{***}	0.170^{***}	0.157^{***}
	(0.003)	(0.003)	(0.004)	(0.005)	(0.005)	(0.005)
Ac.Expr		0.045^{***}	0.050^{***}	0.050^{***}	0.050^{***}	0.050^{***}
		(0.003)	(0.003)	(0.003)	(0.003)	(0.003)
AFQT x Ac.Expr.			0.005^{***}	0.005^{***}	0.004^{***}	0.005^{***}
			(0.000)	(0.000)	(0.000)	(0.000)
Black x Ac.Expr				0.000		0.001
				(0.001)		(0.001)
Black x Pot. Experience					-0.007***	
					(0.001)	
Highest Grade Attd.						0.003^{***}
						(0.000)
Constant	6.133^{***}	6.055^{***}	6.058^{***}	6.059^{***}	6.039^{***}	6.046^{***}
	(0.010)	(0.012)	(0.012)	(0.012)	(0.012)	(0.012)
Observations	33,746	33,746	33,746	33,746	33,746	33,732
R-squared	0.458	0.461	0.464	0.464	0.465	0.465

Table 1:	Wages in the	e Main	(CPS) Job
Linea	ar Regression	Coeffic	ients Shown

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Log wage in the CPI (main) job

b. Observations before completing finmal education are dropped

c. AFQT score is standadized within each birth year cohort

	Linear Re	egression C	oefficients S	Shown		
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.001	-0.000	-0.002	0.018	0.071^{***}	0.010
	(0.010)	(0.010)	(0.010)	(0.016)	(0.017)	(0.016)
AFQT	0.199^{***}	0.197^{***}	0.138^{***}	0.143^{***}	0.148^{***}	0.141^{***}
	(0.005)	(0.005)	(0.007)	(0.007)	(0.007)	(0.007)
Ac.Expr		0.038^{***}	0.048^{***}	0.049^{***}	0.048^{***}	0.049^{***}
		(0.005)	(0.005)	(0.006)	(0.005)	(0.006)
AFQT x Ac.Expr.			0.008^{***}	0.007^{***}	0.007^{***}	0.007^{***}
			(0.001)	(0.001)	(0.001)	(0.001)
Black x Ac.Expr				-0.003*		-0.003
				(0.002)		(0.002)
Black x Pot. Experience					-0.008***	
					(0.001)	
Highest Grade Attd.						0.002^{***}
						(0.000)
Constant	6.080***	6.031^{***}	6.032^{***}	6.025^{***}	6.011^{***}	6.016^{***}
	(0.016)	(0.017)	(0.017)	(0.018)	(0.018)	(0.018)
Observations	$12,\!995$	$12,\!995$	$12,\!995$	$12,\!995$	$12,\!995$	$12,\!992$
R-squared	0.482	0.484	0.490	0.490	0.491	0.490

 Table 2: Beginning Wages in the Main (CPS) Job

 Lincor Begression Coefficients Shown

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Log wage in the CPI (main) job

b. Observations before completing finmal education are dropped

c. AFQT score is standadized within each birth year cohort

	Linear Re	gression Co	pefficients S	shown		
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.035	0.043	0.043	0.113^{***}	0.036	0.111^{**}
	(0.030)	(0.030)	(0.030)	(0.043)	(0.042)	(0.043)
AFQT	0.112^{***}	0.107^{***}	0.107^{***}	0.125^{***}	0.106^{***}	0.124^{***}
	(0.015)	(0.015)	(0.019)	(0.020)	(0.019)	(0.020)
Ac.Expr		0.093^{***}	0.093^{***}	0.108^{***}	0.094^{***}	0.108^{***}
		(0.029)	(0.029)	(0.030)	(0.029)	(0.030)
AFQT x Ac.Expr.			0.000	-0.019	0.001	-0.019
			(0.011)	(0.013)	(0.011)	(0.013)
Black x $Ac.Expr$				-0.076**		-0.076**
				(0.033)		(0.033)
Black x Pot. Experience					0.003	
					(0.013)	
Highest Grade Attd.						0.001
						(0.001)
Constant	5.972***	5.909^{***}	5.909^{***}	5.893***	5.911***	5.890^{***}
	(0.041)	(0.045)	(0.045)	(0.046)	(0.046)	(0.046)
Observations	1,042	$1,\!042$	1,042	1,042	$1,\!042$	1,042
R-squared	0.177	0.185	0.185	0.189	0.185	0.189

 Table 3: First wage after education in the Main (CPS) Job

 Linear Regression Coefficients Shown

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Log first wage in CPS(main) job after completing education

b. Observations before completing finmal education are dropped

c. AFQT score is standadized within each birth year cohort

	Linear R	egression C	Coefficients S	Shown		
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.013**	0.013**	-0.050***	-0.003	0.067***	0.056**
	(0.006)	(0.006)	(0.014)	(0.023)	(0.024)	(0.024)
AFQT	0.169***	0.164***	0.166***	0.132***	0.132***	0.130***
	(0.004)	(0.004)	(0.004)	(0.006)	(0.006)	(0.006)
Coding Speed	0.004^{***}	0.004^{***}	0.001^{**}	0.003***	0.003^{***}	0.003^{***}
	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)
Ac.Expr		0.048^{***}	0.038^{***}	0.050^{***}	0.047^{***}	0.048^{***}
		(0.003)	(0.003)	(0.004)	(0.004)	(0.004)
Coding x Black			0.002^{***}	0.000	-0.000	-0.000
			(0.000)	(0.001)	(0.001)	(0.001)
Coding x Ac.Expr			0.000***	0.000**	0.000**	0.000**
			(0.000)	(0.000)	(0.000)	(0.000)
AFQT x Ac.Expr.			· · · ·	0.004***	0.004***	0.004***
v 1				(0.001)	(0.001)	(0.001)
Coding x Ac.Expr x Bl			-0.000	0.000***	0.000***	0.000***
eea			(0.000)	(0.000)	(0.000)	(0.000)
Black x Ac.Expr			(0.000)	-0.006**	0.002	0.003
Diden x He.Lxpi				(0.002)	(0.002)	(0.003)
AFQT x Ac.Expr. $= 0$,				(0.002)	(0.003)	(0.000)
$AFQT \times AC.Expt. = 0,$				-	-	-
Black x Pot. Experience					-0.012***	-0.012***
Diack x I of. Experience					(0.0012)	
High ant Care de Attal					(0.001)	(0.001) 0.002^{***}
Highest Grade Attd.						
						(0.000)
Constant	5.989***	5.897***	6.017***	5.948***	5.933***	5.920***
	(0.013)	(0.015)	(0.018)	(0.021)	(0.021)	(0.021)
Observations	33,746	33,746	33,746	33,746	33,746	33,732
R-squared	0.463	0.466	0.469	0.470	0.471	0.472
it squarou			in parenthe		0.111	0.112
	Juluk		in parentine			

Table 4: Wages in the Main (CPS) Job

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Log wage in the CPI (main) job

b. Observations before completing finmal education are dropped

c. AFQT score is standadized within each birth year cohort

	Linear R	egression C	Coefficients	Shown		
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
		0.004	0.000		0 100***	0 100***
Black	0.005	0.004	0.003	0.079**	0.133***	0.126***
	(0.010)	(0.010)	(0.022)	(0.034)	(0.035)	(0.035)
AFQT	0.171***	0.167***	0.166***	0.126***	0.127***	0.125***
	(0.006)	(0.006)	(0.006)	(0.009)	(0.009)	(0.009)
Coding Speed	0.003***	0.003^{***}	-0.000	0.002^{***}	0.002^{***}	0.002^{***}
	(0.000)	(0.000)	(0.001)	(0.001)	(0.001)	(0.001)
Ac.Expr	× ,	0.042***	0.029***	0.046***	0.043***	0.043***
1		(0.005)	(0.006)	(0.006)	(0.006)	(0.006)
Coding x Black		(01000)	0.001	-0.002**	-0.002**	-0.002**
			(0.001)	(0.002)	(0.002)	(0.002)
Coding x Ac.Expr			0.000***	0.000**	0.000**	0.000**
Counig x Ac.Expi			(0.000)			
			(0.000)	(0.000)	(0.000)	(0.000)
$AFQT \ge Ac.Expr.$				0.005***	0.005***	0.005***
				(0.001)	(0.001)	(0.001)
Coding x Ac.Expr x Bl			-0.000**	0.000**	0.000**	0.000**
			(0.000)	(0.000)	(0.000)	(0.000)
Black x Ac.Expr				-0.011***	-0.003	-0.003
				(0.004)	(0.004)	(0.004)
AFQT x Ac.Expr. $= 0$,				-	-	-
Black x Pot. Experience					-0.011***	-0.010***
Diack x 1 of. Experience						
					(0.002)	(0.002)
Highest Grade Attd.						0.002***
						(0.000)
Constant	5.969***	5.910^{***}	6.034^{***}	5.941***	5.932***	5.923***
	(0.021)	(0.023)	(0.027)	(0.031)	(0.031)	(0.031)
Observations	12,995	$12,\!995$	12,995	12,995	12,995	12,992
R-squared	0.485	0.487	0.492	0.494	0.495	0.495
			in parenthe			

Table 5: Beginning Wages in the Main (CPS) Job

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Log wage in the CPI (main) job

b. Observations before completing finmal education are dropped

c. AFQT score is standadized within each birth year cohort

	Linear R	egression C	Coefficients	Shown		
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.037	0.045	0.132**	0.329***	0.304***	0.301***
	(0.031)	(0.030)	(0.066)	(0.093)	(0.097)	(0.097)
AFQT	0.095^{***}	0.090***	0.085^{***}	0.109^{***}	0.108^{***}	0.108^{***}
	(0.019)	(0.019)	(0.019)	(0.026)	(0.026)	(0.026)
Coding Speed	0.002	0.002	0.002	0.003	0.003	0.003
	(0.001)	(0.001)	(0.001)	(0.002)	(0.002)	(0.002)
Ac.Expr		0.093***	0.093^{**}	0.108^{**}	0.112^{**}	0.113^{**}
		(0.029)	(0.037)	(0.044)	(0.044)	(0.044)
Coding x Black			-0.002	-0.007***	-0.007**	-0.007**
			(0.002)	(0.003)	(0.003)	(0.003)
Coding x Ac.Expr			0.000	-0.000	-0.000	-0.000
			(0.001)	(0.001)	(0.001)	(0.001)
AFQT x Ac.Expr.			. ,	-0.027	-0.027	-0.027
				(0.016)	(0.016)	(0.016)
Coding x Ac.Expr x Bl			-0.001	0.005**	0.005**	0.005**
			(0.001)	(0.002)	(0.002)	(0.002)
Black x Ac.Expr			× ,	-0.222***	-0.228***	-0.229***
Ŧ				(0.073)	(0.073)	(0.073)
AFQT x Ac.Expr. $= 0$,				-	-	-
Black x Pot. Experience					0.012	0.013
1					(0.014)	(0.014)
Highest Grade Attd.						0.001
0						(0.001)
Constant	5.908***	5.846***	5.809***	5.791***	5.797***	5.794***
	(0.060)	(0.063)	(0.073)	(0.079)	(0.079)	(0.080)
Observations	1,042	1,042	1,042	1,042	1,042	1,042
R-squared	0.178	0.187	0.189	0.197	0.198	0.198

Table 6: First wage after education in the Main (CPS) Job

Standard errors in parentheses *** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Log first wage in CPS(main) job after completing education

b. Observations before completing finmal education are dropped

c. AFQT score is standadized within each birth year cohort

	Logistic I	Regression (Coefficients S	Shown		
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.129***	0.129***	0.129***	0.152***	0.117**	0.167***
DIACK	(0.033)	(0.034)	(0.034)	(0.132) (0.049)	(0.055)	(0.050)
AFQT	(0.033) - 0.364^{***}	(0.034) - 0.335^{***}	(0.034) - 0.246^{***}	(0.049) - 0.240^{***}	(0.055) -0.247^{***}	-0.237***
•	(0.017)	(0.017)	(0.023)	(0.025)	(0.024)	(0.025)
Ac.Expr	× ,	-0.300***	-0.314***	-0.313***	-0.314***	-0.312***
-		(0.018)	(0.019)	(0.019)	(0.019)	(0.019)
$AFQT \ge Ac.Expr$		· · · ·	-0.017***	-0.018***	-0.017***	-0.018***
			(0.003)	(0.004)	(0.003)	(0.004)
Black x Ac.Expr				-0.004		-0.005
				(0.007)		(0.007)
Black x Pot. Experience					0.002	
					(0.005)	
Highest Grade Attd.						-0.004**
						(0.002)
Constant	-0.178^{***}	0.271^{***}	0.286^{***}	0.278^{***}	0.290^{***}	0.293^{***}
	(0.052)	(0.059)	(0.059)	(0.061)	(0.061)	(0.061)
Observations	32,356	$32,\!356$	32,356	32,356	32,356	32,342

Table 7: Unemployed During Last Year Logistic Regression Coefficients She

Standard errors in parentheses *** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Dummy for if one had unemployment spell during the past year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

	Logistic I	Regression (Coefficients S	Shown		
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.423***	0.389***	0.389***	0.161	0.008	0.236**
AFQT	(0.074) - 0.530^{***}	(0.076) - 0.447^{***}	(0.076) - 0.302^{***}	(0.102) - 0.365^{***}	(0.114) - 0.307^{***}	(0.104) - 0.343^{***}
Ac.Expr	(0.043)	(0.044) -1.384***	(0.056) -1.449***	(0.059) -1.494***	(0.056) -1.445***	(0.059) -1.493***
AFQT x Ac.Expr		(0.087)	(0.088) -0.090*** (0.024)	(0.089) - 0.050^{*}	(0.088) - 0.087^{***}	(0.089) - 0.052^{**} (0.027)
Black x Ac.Expr			(0.024)	(0.026) 0.144^{***} (0.044)	(0.024)	(0.027) 0.143^{***} (0.044)
Black x Pot. Experience				(0.044)	0.085^{***} (0.019)	(0.044)
Highest Grade Attd.					(0.010)	-0.016^{***} (0.004)
Constant	-1.593***	-0.352***	-0.262*	-0.168	-0.134	-0.100
	(0.112)	(0.134)	(0.135)	(0.138)	(0.137)	(0.139)
Observations	31,544	31,544	31,544	31,544	31,544	31,530

Table 8: Unemployed During the InterviewLogistic Regression Coefficients Show

Standard errors in parentheses *** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Dummy for if one was unemployed during the interview

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

	$\frac{\text{Logistic I}}{(1)}$	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.113***	0.110***	-0.066	-0.206**	-0.228**	-0.212*
Ditter	(0.034)	(0.034)	(0.072)	(0.104)	(0.111)	(0.111)
AFQT	-0.259***	-0.221***	-0.210***	-0.058*	-0.059*	-0.056*
Ū	(0.021)	(0.021)	(0.021)	(0.032)	(0.032)	(0.032)
Coding Speed	-0.010***	-0.011***	-0.011***	-0.019***	-0.019***	-0.019***
	(0.001)	(0.001)	(0.002)	(0.002)	(0.002)	(0.002)
Ac.Expr		-0.306***	-0.302***	-0.371***	-0.370***	-0.369***
		(0.018)	(0.019)	(0.023)	(0.023)	(0.023)
Coding x Black			0.005^{**}	0.010***	0.011^{***}	0.011***
			(0.002)	(0.003)	(0.003)	(0.003)
Coding x Ac.Expr			-0.000	0.001^{***}	0.001^{***}	0.001***
			(0.000)	(0.000)	(0.000)	(0.000)
AFQT x Ac. Expr				-0.029***	-0.029***	-0.029***
				(0.005)	(0.005)	(0.005)
$Coding \ge Ac.Expr \ge Bl$			0.000	-0.001**	-0.001**	-0.001**
			(0.000)	(0.000)	(0.000)	(0.000)
Black x Ac.Expr				0.027^{*}	0.025	0.025
				(0.015)	(0.016)	(0.016)
$AFQT \ge Ac.Expr = o,$				-	-	-
Black x Pot. Experience					0.003	0.003
					(0.006)	(0.006)
Highest Grade Attd.						-0.003**
						(0.002)
Constant	0.210***	0.701^{***}	0.750^{***}	1.043***	1.047^{***}	1.061***
	(0.071)	(0.077)	(0.091)	(0.106)	(0.106)	(0.106)
Observations	32,356	32,356	32,356	32,356	32,356	32,342
		dard errors i $< 0.01, ** p <$	in parenthes	es		

a. Dependent Variable: Dummy for if one had unemployment spell during the past year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.393***	0.351***	0.123	-0.581***	-0.838***	-0.761***
	(0.074)	(0.077)	(0.158)	(0.220)	(0.233)	(0.235)
AFQT	-0.302***	-0.194***	-0.170***	0.060	0.058	0.074
	(0.054)	(0.055)	(0.056)	(0.077)	(0.077)	(0.077)
Coding Speed	-0.020***	-0.022***	-0.024^{***}	-0.045***	-0.044***	-0.044***
	(0.003)	(0.003)	(0.004)	(0.005)	(0.005)	(0.005)
Ac.Expr		-1.401***	-1.418***	-1.968***	-1.946^{***}	-1.941***
		(0.087)	(0.095)	(0.131)	(0.132)	(0.132)
Coding x Black			0.001	0.022^{***}	0.022^{***}	0.022***
			(0.005)	(0.006)	(0.006)	(0.006)
Coding x Ac.Expr			-0.001	0.012^{***}	0.012^{***}	0.012***
			(0.001)	(0.002)	(0.003)	(0.003)
AFQT x Ac.Expr				-0.142***	-0.143***	-0.145***
				(0.034)	(0.035)	(0.035)
Coding x Ac.Expr x Bl			0.004^{***}	-0.009***	-0.009***	-0.009***
-			(0.001)	(0.003)	(0.003)	(0.003)
Black x Ac.Expr				0.461***	0.433***	0.427***
1				(0.105)	(0.107)	(0.107)
AFQT x Ac.Expr = o,				-	-	
Black x Pot. Experience					0.066***	0.064***
_					(0.019)	(0.019)
Highest Grade Attd.					· · · ·	-0.014***
0						(0.004)
Constant	-0.824***	0.509***	0.700***	1.522***	1.577***	1.622***
	(0.156)	(0.177)	(0.201)	(0.242)	(0.244)	(0.244)
Observations	31,544	31,544	31,544	31,544	31,544	31,530
	Stand	dard errors i	in parenthes	ses		

Table 10: Unemployed During the Interview

a. Dependent Variable: Dummy for if one was unemployed during the interview

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

Table		ployed Durin tegression C	•			
	1000000000000000000000000000000000000	(2)	$\frac{\text{beincients S}}{(3)}$	$\frac{10001}{(4)}$	(5)	(6)
VARIABLES	(1)	$\begin{pmatrix} 2 \end{pmatrix}$	(0)	1	(0)	(0)
Black	0.163***	0.163***	-0.106	-0.134	-0.123	-0.124
	(0.034)	(0.035)	(0.165)	(0.240)	(0.175)	(0.240)
AFQT	-0.315***	-0.284***	-0.283***	-0.175***	-0.283***	-0.171***
	(0.018)	(0.018)	(0.018)	(0.027)	(0.018)	(0.027)
Self Esteme	-0.036***	-0.037***	-0.036***	-0.045***	-0.036***	-0.045***
	(0.004)	(0.004)	(0.006)	(0.007)	(0.006)	(0.007)
Ac.Expr		-0.306***	-0.292***	-0.334***	-0.292***	-0.334***
		(0.019)	(0.025)	(0.029)	(0.025)	(0.029)
Selfesteem x Black			0.009	0.015	0.010	0.015
			(0.007)	(0.011)	(0.008)	(0.011)
Selfesteem x Ac.Expr			-0.001	0.001	-0.001	0.001
			(0.001)	(0.001)	(0.001)	(0.001)
Selfesteem x Ac.Expr x Bl			0.001^{**}	-0.000	0.001^{*}	-0.000
			(0.000)	(0.002)	(0.000)	(0.002)
Black x Ac. $Expr$				0.002		0.002
				(0.034)		(0.034)
$AFQT \ge Ac.Expr$				-0.021***		-0.020***
				(0.004)		(0.004)
Black x Pot. Experience					0.002	
					(0.006)	
Highest Grade Attd.					· · · ·	-0.003**
C						(0.002)
Constant	0.606***	1.081***	1.093***	1.270***	1.098***	1.287***
	(0.097)	(0.102)	(0.146)	(0.165)	(0.147)	(0.165)
Observations	31,647	31,647	31,647	31,647	31,647	31,633
		ard errors in	-			
	p<	0.01, ** p<	0.00, ° p<0.	10		

a. Dependent Variable: Dummy for if one had unemployment spell during the past year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

	(1)	(2)	$\frac{\text{oefficients S}}{(3)}$	(4)	(5)	(6)
/ARIABLES	1	1	1	1	1	1
Black	0.462***	0.435***	-0.139	-1.001**	-0.516	-0.923*
JIACK	(0.076)	(0.079)	(0.377)	(0.505)	(0.388)	(0.507)
AFQT	-0.475***	-0.392***	-0.380***	-0.278***	-0.379***	-0.260***
	(0.046)	(0.047)	(0.047)	(0.064)	(0.047)	(0.064)
Self Esteme	-0.038***	-0.037***	-0.046***	-0.077***	-0.047***	-0.076***
	(0.009)	(0.009)	(0.014)	(0.017)	(0.014)	(0.018)
Ac.Expr	(0.000)	-1.342***	-1.351***	-1.834***	-1.344***	-1.832***
1		(0.088)	(0.133)	(0.198)	(0.133)	(0.198)
Selfesteem x Black		()	0.014	0.057**	0.017	0.056**
			(0.017)	(0.023)	(0.017)	(0.023)
Selfesteem x Ac.Expr			-0.004	0.017**	-0.004	0.017**
I			(0.005)	(0.008)	(0.005)	(0.008)
Selfesteem x Ac.Expr x Bl			0.008***	-0.019*	0.007***	-0.019*
-			(0.002)	(0.010)	(0.002)	(0.010)
Black x Ac.Expr			~ /	0.542**	~ /	0.545**
1				(0.216)		(0.216)
AFQT x Ac.Expr				-0.063**		-0.065**
•				(0.029)		(0.029)
Black x Pot. Experience					0.077***	· · · ·
-					(0.020)	
lighest Grade Attd.					× ,	-0.015***
0						(0.004)
Constant	-0.766***	0.427^{*}	0.765^{**}	1.482***	0.889***	1.526***
	(0.221)	(0.239)	(0.336)	(0.401)	(0.337)	(0.402)
Observations	30,850	30,850	30,850	30,850	30,850	30,836

TT 1 10 TT 1 D . . . 1 т .

a. Dependent Variable: Dummy for if one was unemployed during the interview

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

Та	ble 13: Uner	mployed Du Regression (0			
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	(2)	(5)	1	1	(0)
Black	0.135^{***}	0.134^{***}	0.247^{**}	0.350^{**}	0.215^{*}	0.371^{**}
	(0.034)	(0.034)	(0.118)	(0.171)	(0.126)	(0.172)
AFQT	-0.356***	-0.328***	-0.326***	-0.217***	-0.327***	-0.213***
	(0.017)	(0.017)	(0.017)	(0.026)	(0.017)	(0.026)
Locus of Control	0.022^{***}	0.022^{***}	0.035^{***}	0.045^{***}	0.034^{***}	0.046^{***}
	(0.006)	(0.006)	(0.010)	(0.011)	(0.010)	(0.011)
Ac.Expr		-0.304***	-0.297***	-0.288***	-0.298***	-0.288***
		(0.019)	(0.022)	(0.024)	(0.022)	(0.024)
Control x Black			-0.021	-0.022	-0.020	-0.022
			(0.013)	(0.018)	(0.013)	(0.018)
Control x Ac.Expr			-0.002	-0.003**	-0.002	-0.003**
			(0.001)	(0.002)	(0.001)	(0.002)
Control x Ac.Expr x Bl			0.002**	0.002	0.001*	0.002
			(0.001)	(0.003)	(0.001)	(0.003)
Black x Ac.Expr				-0.019		-0.020
				(0.024)		(0.024)
$AFQT \ge Ac.Expr$				-0.021***		-0.021***
				(0.004)		(0.004)
Black x Pot. Experience					0.004	
					(0.006)	
Highest Grade Attd.						-0.004**
						(0.002)
Constant	-0.380***	0.077	-0.009	-0.116	0.002	-0.104
	(0.074)	(0.079)	(0.105)	(0.116)	(0.106)	(0.116)
Observations	32,059	32,059	32,059	32,059	$32,\!059$	32,045
	Stand	dard errors i	in parenthes	ses		

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Dummy for if one had unemployment spell during the past year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

Tabl	e 14: Unem		0			
			Coefficients S			
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	0.415***	0.382***	0.179	-0.368	-0.158	-0.320
	(0.074)	(0.077)	(0.272)	(0.375)	(0.284)	(0.377)
AFQT	-0.542***	-0.467***	-0.460***	-0.356***	-0.460***	-0.334***
	(0.045)	(0.045)	(0.046)	(0.063)	(0.046)	(0.063)
Locus of Control	0.011	0.004	0.047^{*}	0.014	0.045^{*}	0.013
	(0.014)	(0.015)	(0.025)	(0.031)	(0.025)	(0.031)
Ac.Expr		-1.378***	-1.168***	-1.378***	-1.168***	-1.390***
		(0.088)	(0.118)	(0.152)	(0.119)	(0.153)
Control x Black		· · · ·	-0.010	0.057	-0.009	0.060
			(0.029)	(0.040)	(0.029)	(0.040)
Control x Ac.Expr			-0.034***	-0.012	-0.032***	-0.010
			(0.009)	(0.014)	(0.009)	(0.014)
Control x Ac.Expr x Bl			0.021***	-0.020	0.019^{***}	-0.022
			(0.005)	(0.017)	(0.005)	(0.017)
Black x Ac. $Expr$				0.323^{**}		0.340^{**}
				(0.156)		(0.157)
$AFQT \ge Ac.Expr$				-0.067**		-0.069**
				(0.028)		(0.028)
Black x Pot. Experience					0.080^{***}	
					(0.019)	
Highest Grade Attd.						-0.016***
-						(0.004)
Constant	-1.689***	-0.404**	-0.622**	-0.300	-0.493*	-0.217
	(0.164)	(0.185)	(0.257)	(0.298)	(0.258)	(0.300)
Observations	31,266	31,266	31,266	31,266	31,266	31,252
	,	,	in parenthes	,	,	,
			-0.05 * n < 0			

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Dummy for if one was unemployed during the interview

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

	Linear F	Regression C	oefficients S	shown		
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	1.308^{***}	1.325^{***}	1.291^{***}	1.933***	1.627***	1.989***
AFQT	(0.123) -1.045***	(0.121) -0.915***		(0.190) -1.414***	(0.215) -1.521***	(0.192) -1.404***
Ac.Expr	(0.056)	(0.056) -1.367***	(0.085) -1.258***	(0.092) -1.234***	(0.088) -1.256***	(0.092) -1.231***
AFQT x Ac.Expr		(0.060)	(0.061) 0.081^{***}	(0.061) 0.062^{***}	(0.061) 0.076^{***}	(0.061) 0.062^{***}
Black x Ac.Expr			(0.008)	(0.009) - 0.087^{***}	(0.008)	(0.009) - 0.089^{***}
Black x Pot. Experience				(0.020)	-0.033*	(0.020)
Highest Grade Attd.					(0.017)	-0.013^{**}
Constant	6.146^{***} (0.200)	8.403^{***} (0.222)	8.378^{***} (0.221)	8.161^{***} (0.227)	8.275^{***} (0.228)	$(0.006) \\ 8.207^{***} \\ (0.228)$
Observations R-squared	$30,791 \\ 0.073$	$30,791 \\ 0.089$	$30,791 \\ 0.092$	$30,791 \\ 0.092$	$30,791 \\ 0.092$	30,777 0.092

Table 15: Number of Weeks Unemployed in Calendar Year Linear Bogrossion Coefficients Shown

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Number of weeks unemployed in calendar Year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

Negative Binomial Regression Coefficients Shown							
	(1)	(2)	(3)	(4)	(5)	(6)	
VARIABLES	1	1	1	1	1	1	
Black	0.293***	0.294***	0.294***	0.266***	0.326***	0.286***	
AFQT	(0.050) - 0.332^{***}	(0.050) - 0.306^{***}	(0.050) - 0.295^{***}	(0.078) - 0.301^{***}	(0.091) - 0.290^{***}	(0.079) - 0.296^{***}	
ArQI	(0.023)	(0.023)	(0.035)	(0.037)	(0.037)	(0.038)	
Ac.Expr		-0.236^{***}	-0.238^{***}	-0.239^{***}	-0.238^{***}	-0.239***	
AFQT x Ac.Expr		(0.028)	(0.028) -0.002	(0.028) -0.001	(0.028) -0.002	(0.028) -0.001	
			(0.004)	(0.004)	(0.004)	(0.004)	
Black x Ac.Expr				0.004 (0.009)		0.003 (0.009)	
Black x Pot. Experience				()	-0.003	()	
Highest Grade Attd.					(0.008)	-0.004^{*} (0.003)	
Constant	1.893***	2.063***	2.062***	2.071***	2.504***	2.083^{***}	
Constant	(0.084) 2.512^{***}	(0.087) 2.504^{***}	(0.087) 2.504^{***}	(0.089) 2.504^{***}	(0.015) 2.051^{***}	(0.015) 2.504^{***}	
Constant	(0.014)	(0.015)	(0.015)	(0.015)	(0.090)	(0.089)	
Observations	30,791	30,791	30,791	30,791	30,791	30,777	
			in parenthes (0.05×0.05)				

Table 16: Number of Weeks Unemployed in Calendar Year Negative Binomial Regression Coefficients Shown

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Number of weeks unemployed in calendar Year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

	Linear R	Linear Regression Coefficients Shown							
	(1)	(2)	(3)	(4)	(5)	(6)			
VARIABLES	1	1	1	1	1	1			
Black	1.325***	1.337***	0.920**	0.903	0.791*	0.993			
	(0.123)	(0.122)	(0.417)	(0.654)	(0.465)	(0.655)			
AFQT	-1.007***	-0.882***	-0.894***	-1.347***	-0.895***	-1.332***			
~	(0.058)	(0.057)	(0.057)	(0.095)	(0.058)	(0.095)			
Locus of Control	0.090***	0.086***	0.143***	0.116***	0.140***	0.121***			
	(0.022)	(0.021)	(0.037)	(0.041)	(0.038)	(0.041)			
Ac.Expr	()	-1.374***	-1.198***	-1.182***	-1.200***	-1.179***			
-		(0.060)	(0.068)	(0.071)	(0.068)	(0.071)			
Control x Black		· · · ·	0.157***	0.114	0.163***	0.112			
			(0.047)	(0.070)	(0.048)	(0.070)			
Control x Ac.Expr			-0.010***	-0.007*	-0.009***	-0.007*			
			(0.003)	(0.004)	(0.004)	(0.004)			
Control x Ac.Expr x Bl			-0.015***	-0.009	-0.016***	-0.009			
			(0.002)	(0.007)	(0.003)	(0.007)			
Black x Ac.Expr				-0.006		-0.009			
				(0.068)		(0.068)			
$AFQT \ge Ac.Expr$				0.057^{***}		0.057^{***}			
				(0.009)		(0.009)			
Black x Pot. Experience					0.013				
					(0.021)				
Highest Grade Attd.						-0.015**			
					0.0.4.0.4.4.4	(0.006)			
Constant	5.341***	7.636***	6.799***	7.148***	6.842***	7.166***			
	(0.271)	(0.287)	(0.386)	(0.421)	(0.392)	(0.422)			
Observations	30,497	30,497	30,497	$30,\!497$	$30,\!497$	30,483			
R-squared	0.075	0.090	0.093	0.094	0.093	0.094			

Table 17:	Number	of Weeks	Unemployed	in Calendar	Year

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Number of weeks unemployed in calendar Year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

Negative Binomial Regression Coefficients Shown								
	(1)	(2)	(3)	(4)	(5)	(6)		
VARIABLES	1	1	1	1	1	1		
Black	0.300^{***}	0.302^{***}	0.522^{***}	0.519^{*}	0.686^{***}	0.565^{**}		
	(0.050)	(0.050)	(0.183)	(0.283)	(0.219)	(0.284)		
AFQT	-0.332***	-0.306***	-0.305***	-0.271^{***}	-0.302***	-0.266***		
	(0.024)	(0.024)	(0.024)	(0.039)	(0.024)	(0.039)		
Locus of Control	0.018^{*}	0.017^{*}	0.049^{***}	0.051^{***}	0.054^{***}	0.053^{***}		
	(0.009)	(0.009)	(0.016)	(0.018)	(0.017)	(0.018)		
Ac.Expr		-0.238***	-0.213***	-0.213***	-0.210***	-0.212***		
		(0.028)	(0.031)	(0.032)	(0.031)	(0.032)		
Control x Black			-0.032	-0.028	-0.043*	-0.031		
			(0.021)	(0.031)	(0.022)	(0.031)		
Control x Ac. $Expr$			-0.003**	-0.004*	-0.004**	-0.004*		
			(0.002)	(0.002)	(0.002)	(0.002)		
Control x Ac.Expr x Bl			0.001	0.000	0.003^{*}	0.000		
			(0.001)	(0.003)	(0.002)	(0.003)		
Black x Ac.Expr				0.001		-0.001		
				(0.030)		(0.030)		
AFQT x Ac.Expr				-0.005		-0.005		
				(0.004)		(0.004)		
Black x Pot. Experience					-0.018			
					(0.013)			
Highest Grade Attd.						-0.005*		
						(0.003)		
Constant	2.511^{***}	1.888^{***}	1.638^{***}	1.609^{***}	2.502^{***}	1.607^{***}		
	(0.119)	(0.015)	(0.015)	(0.176)	(0.165)	(0.015)		
Constant	1.707^{***}	2.503^{***}	2.502^{***}	2.502^{***}	1.581^{***}	2.501^{***}		
	(0.015)	(0.121)	(0.160)	(0.015)	(0.015)	(0.176)		
Observations	30,497	30,497	30,497	30,497	30,497	30,483		
			in parenthes					
	*** p<	<0.01, ** p<	<0.05, * p<0	0.10				

Table 18: Number of Weeks Unemployed in Calendar Year

a. Dependent Variable: Number of weeks unemployed in calendar Year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

		Regression C	•			
	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	1	1	1	1	1	1
Black	1.405***	1.427***	0.485	-0.472	0.231	-0.430
	(0.125)	(0.124)	(0.593)	(0.921)	(0.646)	(0.921)
AFQT	-0.901***	-0.764***	-0.766***	-1.154***	-0.767***	-1.143***
5	(0.060)	(0.059)	(0.059)	(0.098)	(0.059)	(0.098)
Self Esteme	-0.105***	-0.106***	-0.217***	-0.201***	-0.219***	-0.201***
	(0.013)	(0.013)	(0.022)	(0.025)	(0.022)	(0.025)
Ac.Expr	× ,	-1.374***	-1.554***	-1.484***	-1.556***	-1.482***
1		(0.061)	(0.074)	(0.083)	(0.074)	(0.083)
Selfesteem x Black		()	0.091***	0.116***	0.098***	0.117***
			(0.027)	(0.041)	(0.028)	(0.041)
Selfesteem x Ac.Expr			0.012***	0.011***	0.013***	0.011***
			(0.002)	(0.002)	(0.002)	(0.002)
Selfesteem x Ac.Expr x Bl			-0.007***	-0.011**	-0.007***	-0.011**
			(0.001)	(0.004)	(0.001)	(0.004)
Black x Ac.Expr				0.144		0.146
				(0.097)		(0.097)
AFQT x Ac.Expr				0.049^{***}		0.049^{***}
				(0.010)		(0.010)
Black x Pot. Experience					0.022	
					(0.022)	
Highest Grade Attd.						-0.013**
						(0.006)
Constant	8.420***	10.722^{***}	12.828^{***}	12.586^{***}	12.906^{***}	12.643^{***}
	(0.345)	(0.357)	(0.534)	(0.595)	(0.540)	(0.596)
Observations	30,084	30,084	30,084	30,084	30,084	30,070
R-squared	0.076	0.091	0.094	0.095	0.094	0.095
1		dand onnong i				

Table 19: Number of Weeks Unemployed in Calendar Year

Standard errors in parentheses (** n < 0.01, ** n < 0.05, * n < 0.1)

$$p < 0.01$$
, ** $p < 0.05$, * $p < 0.10$

a. Dependent Variable: Number of weeks unemployed in calendar Year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

Negative Binomial Regression Coefficients Shown							
	(1)	(2)	(3)	(4)	(5)	(6)	
VARIABLES	1	1	1	1	1	1	
Black	0.346^{***}	0.349^{***}	-0.356	-0.577	-0.243	-0.560	
	(0.051)	(0.051)	(0.251)	(0.396)	(0.298)	(0.396)	
AFQT	-0.290***	-0.263***	-0.266***	-0.239***	-0.265***	-0.235***	
	(0.024)	(0.024)	(0.025)	(0.040)	(0.025)	(0.040)	
Self Esteme	-0.040***	-0.040***	-0.039***	-0.043***	-0.038***	-0.043***	
	(0.005)	(0.005)	(0.009)	(0.010)	(0.009)	(0.010)	
Ac.Expr		-0.241***	-0.212***	-0.226***	-0.211***	-0.224***	
		(0.028)	(0.033)	(0.036)	(0.033)	(0.036)	
Selfesteem x Black			0.030^{***}	0.040^{**}	0.026^{**}	0.040^{**}	
			(0.011)	(0.017)	(0.012)	(0.017)	
Selfesteem x Ac.Expr			-0.002*	-0.001	-0.002*	-0.001	
			(0.001)	(0.001)	(0.001)	(0.001)	
Selfesteem x Ac.Expr x Bl			0.000	-0.001	0.001	-0.001	
			(0.000)	(0.002)	(0.001)	(0.002)	
Black x Ac.Expr				0.030		0.028	
				(0.043)		(0.043)	
AFQT x Ac.Expr				-0.004		-0.003	
				(0.004)		(0.004)	
Black x Pot. Experience				()	-0.010	× ,	
-					(0.014)		
Highest Grade Attd.						-0.004*	
0						(0.003)	
Constant	2.720***	2.500^{***}	2.499^{***}	2.955***	2.820***	2.957***	
	(0.015)	(0.015)	(0.015)	(0.246)	(0.015)	(0.246)	
Constant	2.509***	2.907***	2.854***	2.499***	2.499***	2.499***	
	(0.142)	(0.144)	(0.223)	(0.015)	(0.228)	(0.015)	
	× /	× /	× /	× /	× /	× /	
Observations	30,084	30,084	30,084	30,084	30,084	30,070	
		ard errors ir					
	*** p<	0.01, ** p<	0.05, * p < 0.05	.10			

Table 20: Number of Weeks Unemployed in Calendar Year

a. Dependent Variable: Number of weeks unemployed in calendar Year

b. Observations before completing final education are dropped

c. AFQT score is standadized within each birth year cohort

	(1)	(2)	$\frac{\text{Coefficients}}{(3)}$	(4)	(5)
VARIABLES	1	1	1	1	1
Black	-0.014	-0.028	-0.039	0.081**	0.067*
	(0.029)	(0.029)	(0.029)	(0.034)	(0.035)
AFQT	0.106^{***}	0.089^{***}	0.083^{***}	0.081^{***}	0.075^{***}
	(0.012)	(0.013)	(0.013)	(0.013)	(0.013)
Job Duration		0.003***	0.008***	0.009***	0.009***
		(0.000)	(0.000)	(0.000)	(0.000)
Job Duration Sq.			-0.000***	-0.000***	-0.000***
			(0.000)	(0.000)	(0.000)
Job Duration x Black			(0.000)	-0.002***	-0.002***
505 Duration A Diator				(0.000)	(0.000)
Highest Grade Attd.				(0.000)	0.004^{**}
Ingliest Glade Attu.					
C I I	0 559***	0 700***	0 000***	0.007***	(0.001)
Constant	-0.553***	-0.730***	-0.898***	-0.937***	-0.956***
	(0.015)	(0.017)	(0.019)	(0.020)	(0.021)
Observations	30,560	30,560	30,560	30,560	30,524
	Star	dard errors	in parenthe	ses	
			<0.05 * n <		

Table 21: Quit Job, Given the Spell Ended, Remained in LF Logistic Regression Coefficients Shown

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Dummy if quit job given the job spell ended (but remained in LF)

b. AFQT score is standadized within each birth year cohort

Logistic Regression Coefficients Shown					
	(1)	(2)	(3)	(4)	(5)
VARIABLES	1	1	1	1	1
Black	0.014	0.028	0.039	-0.081**	-0.067*
	(0.029)	(0.029)	(0.029)	(0.034)	(0.035)
AFQT	-0.106***	-0.089***	-0.083***	-0.081***	-0.075***
	(0.012)	(0.013)	(0.013)	(0.013)	(0.013)
Job Duration		-0.003***	-0.008***	-0.009***	-0.009***
		(0.000)	(0.000)	(0.000)	(0.000)
Job Duration Sq.			0.000^{***}	0.000^{***}	0.000^{***}
			(0.000)	(0.000)	(0.000)
Job Duration x Black				0.002^{***}	0.002^{***}
				(0.000)	(0.000)
Highest Grade Attd.					-0.004**
					(0.001)
Constant	0.553^{***}	0.730^{***}	0.898^{***}	0.937^{***}	0.956^{***}
	(0.015)	(0.017)	(0.019)	(0.020)	(0.021)
Observations	$30,\!560$	$30,\!560$	$30,\!560$	$30,\!560$	$30,\!524$
Standard errors in parentheses					

Table 22: Layed-off, Given that the Spell Ended

*** p<0.01, ** p<0.05, * p<0.10

a. Dependent Variable: Dummy if layed off given the job spell endedb. AFQT score is standadized within each birth year cohort

Race
by
Work -
Left
Reasons
23:
Table

		Race	ce		
Reason Left Job - End of Temp Jobs Considered as Layeoffs	Black	NBNH	ΗN	Total	al
	Freq %	Freq %		Freq %	%
Layoff, job eliminated	$5556 \ 48.351191550.281747149.6\overline{5}$	51191	550.28	17471	49.65
Discharged or fired	772 6.72 1184 5.00 1956 5.56	1184	5.00	1956	5.56
Quit for pregnancy, childbirth or adoption of a child	164 1.43	$1.43 \ \ 356 \ \ 1.50 \ \ 520 \ \ 1.48$	1.50	520	1.48
Quit to look for another job	2358 20.525492 23.177850 22.31	25492	23.17	7850	22.31
Quit to take another job	970 8.44	8.44 2313 9.76 3283 9.33	9.76	3283	9.33
Other (SPECIFY)	1248 10.8	$10.861844\ 7.78\ 3092$	7.78	3092	8.79
Quit because Rs ill health, disability, or medical problems	95 0.83	0.83 101	0.43	196	0.56
Moved to another geographic area	53 0.46	75	0.32	128	0.36
Quit to be with children or other family members	17 0.15	31	0.13	48	0.14
Quit because didn't like job, boss, coworkers, pay or benefits	100 0.87	125	0.53	225	0.64
Quit to attend school or traning	59 0.51	170	0.72	229	0.65
Went to jail, prison, had legal problems	38 0.33	19	0.08	57	0.16
Transportation problems	22 0.19	13	0.05	35	0.10
Retired	11 0.10	21	0.09	32	0.09
No desirable assignments available	9 0.08	11	0.05	20	0.06
Temporary job became permanent	3 0.03	9	0.03	6	0.03
Dissatisfied with job matching service	1 0.01	က	0.01	4	0.01
Business failed or bankruptcy	1 0.01	9	0.03	7	0.02
Sold business to another person or firm	5 0.04	1	0.03	12	0.03
Business temporarily inactive	2 0.02	0	0.00	7	0.01
Closed business down or dissolved partnership	7 0.06	7	0.03	14	0.04
Total	11491100.023699100.035190100.00	02369	9100.0	G 5190	100.00

Source: NLSY79

		Race	
Highest Degree Cmptd.	Black	White	Total
	Mean duration	Mean duration	Mean duration
High school diploma (or equivalent)	86	59	71
Associate/Junior College (AA)	69	53	58
Bachelor of Arts Degree (BA)	56	52	53
Bachelor of Science (BS)	54	45	47
Master's Degree (MA, MBA, MS, MSW)	68	44	49
Doctoral Degree (PhD)	40	56	53
Professional Degree (MD, LLD, DDS)	48	61	60
Other (SPECIFY)	118	55	66
Total	80	55	63

Table 24: Unemployment Duration by Highest Degree Attained

Source: NLSY79

Table 25: Employment Duration by	y Highest Degree Attained
----------------------------------	---------------------------

		Race	
Highest Degree Cmptd.	Black Mean duration	White Mean duration	Total Mean duration
High school diploma (or equivalent)	94	123	111
Associate/Junior College (AA)	99	125	117
Bachelor of Arts Degree (BA)	105	123	117
Bachelor of Science (BS)	116	134	130
Master's Degree (MA, MBA, MS, MSW)	117	128	125
Doctoral Degree (PhD)	112	107	108
Professional Degree (MD, LLD, DDS)	119	122	122
Other (SPECIFY)	80	116	111
Total	98	125	115

Source: NLSY79

	Race		
Highest Degree Cmptd.	Black	White	Total
	Count CASE	Count CASE	Count CASE
High school diploma (or equivalent)	480	710	1190
Associate/Junior College (AA)	52	120	172
Bachelor of Arts Degree (BA)	36	89	125
Bachelor of Science (BS)	53	200	253
Master's Degree (MA, MBA, MS, MSW)	31	98	129
Doctoral Degree (PhD)	4	15	19
Professional Degree (MD, LLD, DDS)	1	18	19
Other (SPECIFY)	8	42	50
Total	665	1292	1957

Table 26: Distribution of Highest Degree Attained - by Race

Source: NLSY79

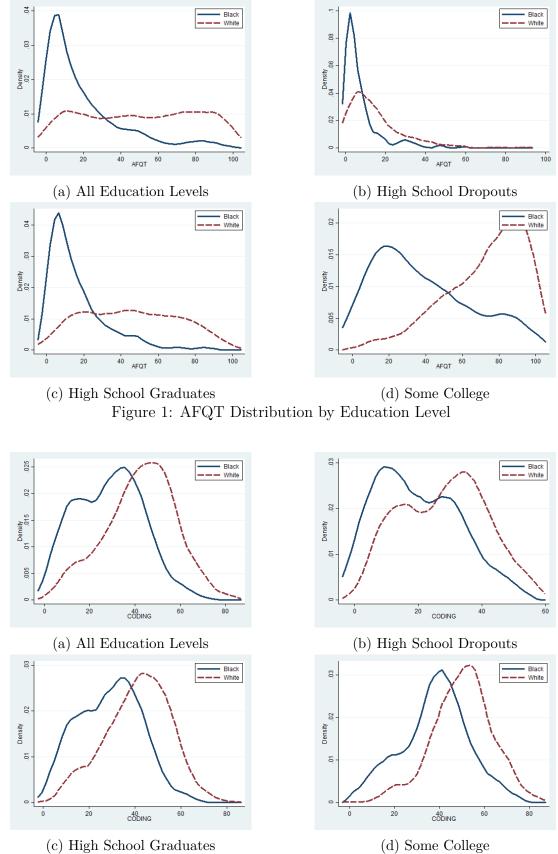
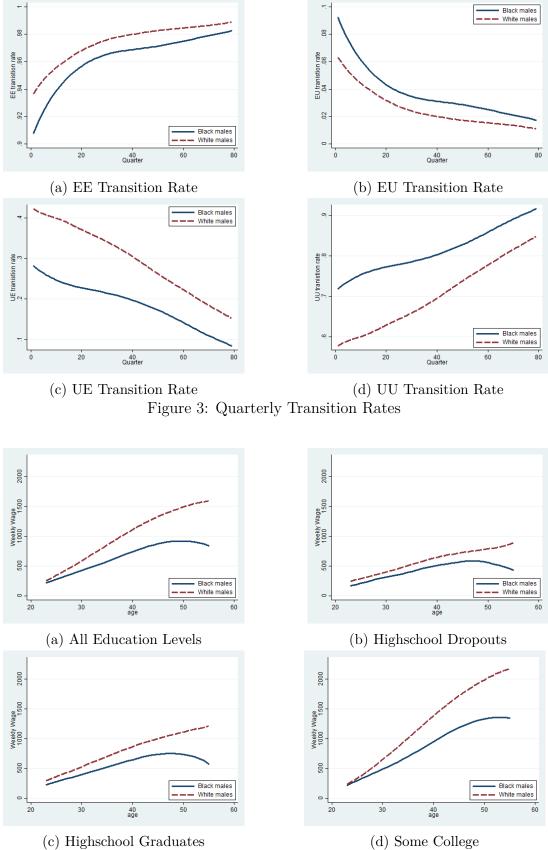
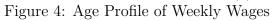
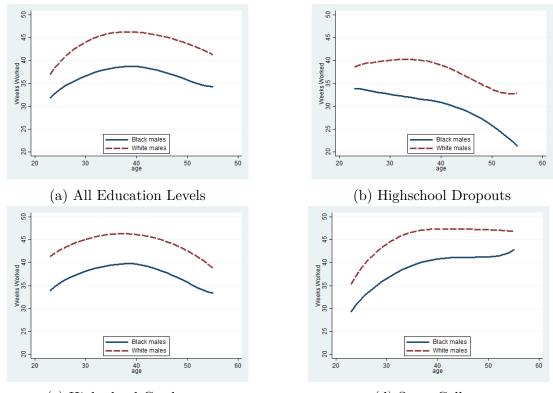


Figure 2: Coding Speed Distribution by Education Level







(c) Highschool Graduates (d) Some College Figure 5: Age Profile of Weeks Worked

References

- Tom Ahn, Peter Arcidiacono, Alvin Murphy, and Omari Swinton. Explaining cross-racial differences in teenage labor force participation: Results from a two-sided matching model. *Journal of Econometrics*, 156(1):201–211, 2010.
- Mathilde Almlund, Angela Lee Duckworth, James Heckman, and Tim Kautz. Personality psychology and economics. In *Handbook of the Economics of Education*, volume 4, pages 1–181. Elsevier, 2011.
- Joseph G. Altonji. Employer Learning, Statistical Discrimination and Occupational Attainment. *American Economic Review*, 95(2):112–117, May 2005.
- Joseph G. Altonji and Charles R. Pierret. Employer Learning and Statistical Discrimination. *The Quarterly Journal of Economics*, 116(1):313–350, 2001.
- Peter Arcidiacono, Patrick Bayer, and Aurel Hizmo. Beyond Signaling and Human Capital: Education and the Revelation of Ability. *American Economic Journal: Applied Economics*, 2(4):76–104, October 2010.
- Kenneth J Arrow. What has economics to say about racial discrimination? *The journal* of economic perspectives, pages 91–100, 1998.
- Gary S Becker. The economics of discrimination. University of Chicago press, 2010.
- Dan A Black. Discrimination in an Equilibrium Search Model. Journal of Labor Economics, 13(2):309–33, April 1995.
- David Blau and Janet Currie. Pre-school, day care, and after-school care: who's minding the kids? *Handbook of the Economics of Education*, 2:1163–1278, 2006.
- Lex Borghans, Angela Lee Duckworth, James J Heckman, and Bas Ter Weel. The economics and psychology of personality traits. *Journal of human Resources*, 43(4):972– 1059, 2008.

- Audra J. Bowlus and Zvi Eckstein. Discrimination and skill differences in an equilibrium search model^{*}. *International Economic Review*, 43(4):1309–1345, 2002.
- Costas Cavounidis and Kevin Lang. Discrimination and Worker Evaluation. NBER Working Papers 21612, National Bureau of Economic Research, October 2015.
- Flavio Cunha, James J Heckman, Lance Lochner, and Dimitriy V Masterov. Interpreting the evidence on life cycle skill formation. *Handbook of the Economics of Education*, 1: 697–812, 2006.
- Roland G Fryer Jr, Devah Pager, and Jörg L Spenkuch. Racial disparities in job finding and offered wages. *Journal of Law and Economics*, 56(3):633–689, 2013.
- Karsten T Hansen, James J Heckman, and Kathleen J Mullen. The effect of schooling and ability on achievement test scores. *Journal of econometrics*, 121(1-2):39–98, 2004.
- James J Heckman. Policies to foster human capital. *Research in economics*, 54(1):3–56, 2000.
- James J Heckman and Tim Kautz. Hard evidence on soft skills. Labour economics, 19 (4):451–464, 2012.
- James J Heckman and Yona Rubinstein. The importance of noncognitive skills: Lessons from the ged testing program. *American Economic Review*, 91(2):145–149, 2001.
- James J Heckman, Maria Isabel Larenas, and Sergio Urzua. Accounting for the effect of schooling and abilities in the analysis of racial and ethnic disparities in achievement test scores. Unpublished manuscript, University of Chicago, Department of Economics, 2004.
- James J Heckman, Jora Stixrud, and Sergio Urzua. The effects of cognitive and noncognitive abilities on labor market outcomes and social behavior. *Journal of Labor economics*, 24(3):411–482, 2006.
- Lynn A Karoly and Hugh P Levaux. Investing in our children: What we know and don't

know about the costs and benefits of early childhood interventions. Rand Corporation, 1998.

- Tim Kautz, James J Heckman, Ron Diris, Bas Ter Weel, and Lex Borghans. Fostering and measuring skills: Improving cognitive and non-cognitive skills to promote lifetime success. Technical report, National Bureau of Economic Research, 2014.
- Kevin Lang and Jee-Yeon K. Lehmann. Racial discrimination in the labor market: Theory and empirics. *Journal of Economic Literature*, 50(4):959–1006, 2012.
- Kevin Lang, Michael Manove, and William T. Dickens. Racial discrimination in labor markets with posted wage offers. *American Economic Review*, 95(4):1327–1340, 2005.
- Fabian Lange. The Speed of Employer Learning. *Journal of Labor Economics*, 25:1–35, 2007.
- Giuseppe Moscarini. Job Matching and the Wage Distribution. *Econometrica*, 73(2): 481–516, 03 2005.
- Derek A Neal and William R Johnson. The role of premarket factors in black-white wage differences. *Journal of political Economy*, 104(5):869–895, 1996.
- Edmund S Phelps. The statistical theory of racism and sexism. *The american economic review*, pages 659–661, 1972.
- Asa Rosen. An equilibrium search-matching model of discrimination. European Economic Review, 41(8):1589–1613, August 1997.
- Sergio Urzua. The effects of cognitive and noncognitive skills on racial and ethnic wage gaps. Unpublished manuscript, Department of Economics, University of Chicago, 2006.