

High-frequency sampled stable CARMA processes

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joint work with Florian Fuchs

- Continuous-time ARMA (CARMA) processes

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- Spectral density estimation

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- Parameter estimation

CARMA Processes

An **ARMA**(p, q), $p, q \in \mathbb{N}_0$ process $(Y_k)_{k \in \mathbb{Z}}$ is the solution of the difference equation:

$$a(B) Y_k = b(B) \xi_k \quad \text{for } k \in \mathbb{N},$$

where

- B is the backward shift operator,
- $(\xi_k)_{k \in \mathbb{Z}}$ is an iid sequence of random variables,
- $a(z) := z^p + a_1 z^{p-1} + \dots + a_p$ with $a_1, \dots, a_p \in \mathbb{R}$,
- $b(z) := b_0 z^q + b_1 z^{q-1} + \dots + b_q$ with $b_0, \dots, b_q \in \mathbb{R}$.

Motivation:

A **continuous-time ARMA(p, q)** (CARMA(p, q)) process $(Y(t))_{t \in \mathbb{R}}$, $p, q \in \mathbb{N}_0$, $p > q$, is the solution of the stochastic differential equation:

$$a(D)Y(t) = b(D)DL(t) \quad \text{for } t \in \mathbb{R},$$

where

- D is the differential-operator (after t),
- $(L(t))_{t \in \mathbb{R}}$ is a Lévy process,
- $a(z) := z^p + a_1 z^{p-1} + \dots + a_p$ with $a_1, \dots, a_p \in \mathbb{R}$,
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- $\int_{\|x\| > 1} \log \|x\| v_L(dx) < \infty.$

Let

$$\mathbf{X}(t) = \int_{-\infty}^t e^{A(t-s)} \mathbf{e}_p dL(s) \quad \text{for } t \in \mathbb{R}$$

in \mathbb{R}^p with $\mathbf{e}_p = (0, \dots, 0, 1)' \in \mathbb{R}^p$ be the solution of the stochastic differential equation

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Example: Ornstein-Uhlenbeck process

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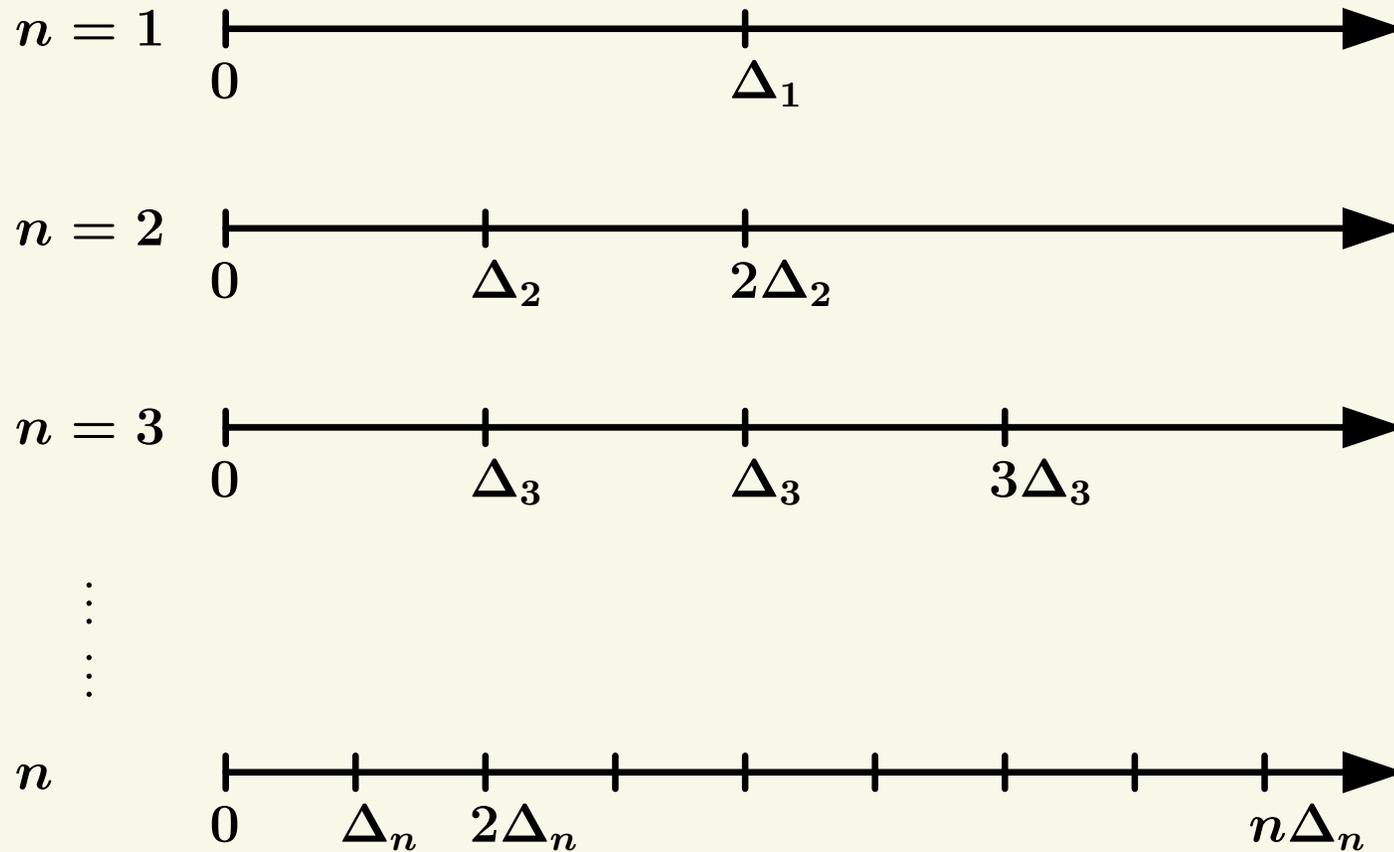
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Spectral Density Estimation

High-frequency observations



$$\Delta_n \downarrow 0 \text{ as } n \rightarrow \infty \quad \text{and} \quad \lim_{n \rightarrow \infty} n\Delta_n = \infty.$$

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Then

$$\lim_{n \rightarrow \infty} \Delta_n g_{\Delta_n}(\omega\Delta_n) \mathbf{1}_{[-\frac{\pi}{\Delta_n}, \frac{\pi}{\Delta_n}]}(\omega) = g_Y(\omega), \quad \omega \in \mathbb{R}.$$

Theorem (Fasen and Fuchs (2013))

The (normalized) periodogram of the sampled sequence

$Y^{\Delta_n} := (Y(k\Delta_n))_{k \in \mathbb{N}}$ is

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Then as $n \rightarrow \infty$,

$$\Delta_n^{2-\frac{2}{\alpha}} I_{n, Y^{\Delta_n}}(\omega\Delta_n) \xrightarrow{\mathcal{D}} \frac{|b(i\omega)|^2}{|a(i\omega)|^2} \cdot \left| \int_{[0,1)} e^{2\pi i s} dL_s \right|^2, \quad \omega \in \mathbb{R} \setminus \{0\}.$$

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\Rightarrow If $\alpha = 2$ this means that as $n \rightarrow \infty$,

$$\Delta_n I_{n, Y^{\Delta_n}}(\omega\Delta_n) \xrightarrow{\mathcal{D}} \frac{\pi}{\sigma_L^2} g_Y(\omega) \cdot \left| \int_{[0,1)} e^{2\pi i s} dL_s \right|^2.$$

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where $(S_j^{\Re}(\tilde{\omega}), S_j^{\Im}(\tilde{\omega}))_{j \in \{1, \dots, m\}}$ is a $(2m)$ -dimensional **stable random vector** with characteristic function

$$\mathbb{E} \left[\exp \left\{ i \left(\sum_{j=1}^m \theta_j S_j^{\Re}(\tilde{\omega}) + \nu_j S_j^{\Im}(\tilde{\omega}) \right) \right\} \right] = \exp \{ -\sigma_L^\alpha \cdot K_{\tilde{\omega}}(\tilde{\theta}, \tilde{\nu}) \},$$

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$$K_{\tilde{\omega}}(\tilde{\theta}, \tilde{v}) = \frac{1}{\lambda^{(m-s)}(\mathcal{M})} \int_{\mathcal{M}} \left| \sum_{j=1}^m \theta_j \cos(2\pi x_j) + v_j \sin(2\pi x_j) \right|^\alpha d\lambda^{(m-s)}(x_1, \dots, x_m),$$

where $\mathcal{M} = \mathcal{M}(\omega_1, \dots, \omega_m)$ is a $(m-s)$ -dimensional linear manifold in $[0, 1)^m$.

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$\alpha = 2$: Then as $n \rightarrow \infty$,

$$\Delta_n I_{n, Y\Delta_n}(\omega\Delta_n) \xrightarrow{\mathcal{D}} 2\pi g_Y(\omega) \left(\frac{N_1^2}{2} + \frac{N_2^2}{2} \right) \stackrel{d}{=} 2\pi g_Y(\omega) E,$$

where N_1 and N_2 are i.i.d. standard normal random variables and E is a standard exponential random variable.

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where

- m_n is a sequence in \mathbb{N} such that

$$m_n \rightarrow \infty \quad \text{and} \quad \frac{m_n}{n\Delta_n} \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

The sequence of weight functions $W_n : \mathbb{Z} \rightarrow \mathbb{R}$ satisfies the following conditions:

- $W_n(k) = W_n(-k), \quad W_n(k) \geq 0, \quad \forall k \in \mathbb{N},$
- $\sum_{|k| \leq m_n} W_n(k) = 1, \quad \forall n \in \mathbb{N},$
- $\max_{|k| \leq m_n} W_n^2(k) = o\left(\frac{1}{m_n}\right) \quad \text{as } n \rightarrow \infty.$

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Theorem (Fasen and Fuchs (2012))

The smoothed periodogram of Y^{Δ_n} is

$$T_{n, Y^{\Delta_n}}(\omega) = \sum_{|k| \leq m_n} W_n(k) I_{n, Y^{\Delta_n}} \left(\omega + \frac{k}{n} \right), \quad \omega \in [-\pi, \pi].$$

Let $\alpha = 2$. Then as $n \rightarrow \infty$,

$$\Delta_n T_{n, Y^{\Delta_n}}(\omega \Delta_n) \xrightarrow{\mathbb{P}} \frac{\pi}{\sigma_L^2} g_Y(\omega) =: \psi(\omega), \quad \omega \in \mathbb{R} \setminus \{0\}.$$

Parameter estimation

Let

$$a(z) := z^p + a_1 z^{p-1} + \dots + a_p,$$

$$b(z) := z^q + b_1 z^{q-1} + \dots + b_q,$$

and the **normalized power transfer function (spectral density)**

$$\psi(\omega) := \frac{|b(i\omega)|^2}{|a(i\omega)|^2}$$

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$$\psi(\omega) := \frac{|b(i\omega)|^2}{|a(i\omega)|^2} = \frac{\prod_{k=1}^q (\omega + i\mu_k) (\omega - i\bar{\mu}_k)}{\prod_{j=1}^p (\omega + i\lambda_j) (\omega - i\bar{\lambda}_j)}$$

where μ_1, \dots, μ_q are the zeros of $b(\cdot)$ and $\lambda_1, \dots, \lambda_p$ are the zeros of $a(\cdot)$.

Assumption

The zeros μ_1, \dots, μ_q and $\lambda_1, \dots, \lambda_p$ are all distinct and possess strictly negative real parts.

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Define

$$\psi(\omega; \theta) := \frac{|(i\omega)^q + b_1(i\omega)^{q-1} + \dots + b_q|^2}{|(i\omega)^p + a_1(i\omega)^{p-1} + \dots + a_p|^2}$$

with $\theta = (a_1, \dots, a_p, b_1, \dots, b_q)$

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Then the parameters of the CARMA process are **identifiable** from $\psi(\cdot; \theta)$.

Suppose we have observed the CARMA(p, q) process on the time grid $\{\Delta_n, \dots, n\Delta_n\}$. Then we choose $m \in \mathbb{N}$ different frequencies $\omega_j \in (0, \pi/\Delta_n)$, $j = 1, \dots, m$,

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$$\hat{\theta} := \operatorname{argmin}_{\theta \in \Theta} \sum_{j=1}^m \left| \log(\mathbf{C}_\theta \cdot \Psi_\theta(\omega_j)) - \log \left(\underbrace{\Delta_n \hat{T}_{n, \gamma \Delta_n}(\omega_j \Delta_n)}_{\substack{\mathbb{P} \\ \rightarrow \mathbf{C}_{\theta^*} \cdot \Psi_{\theta^*}(\omega_j)}} \right) \right|^2.$$

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\rightsquigarrow variance stabilizing technique

Consider a CARMA(2, 1) process which is the strictly stationary solution to

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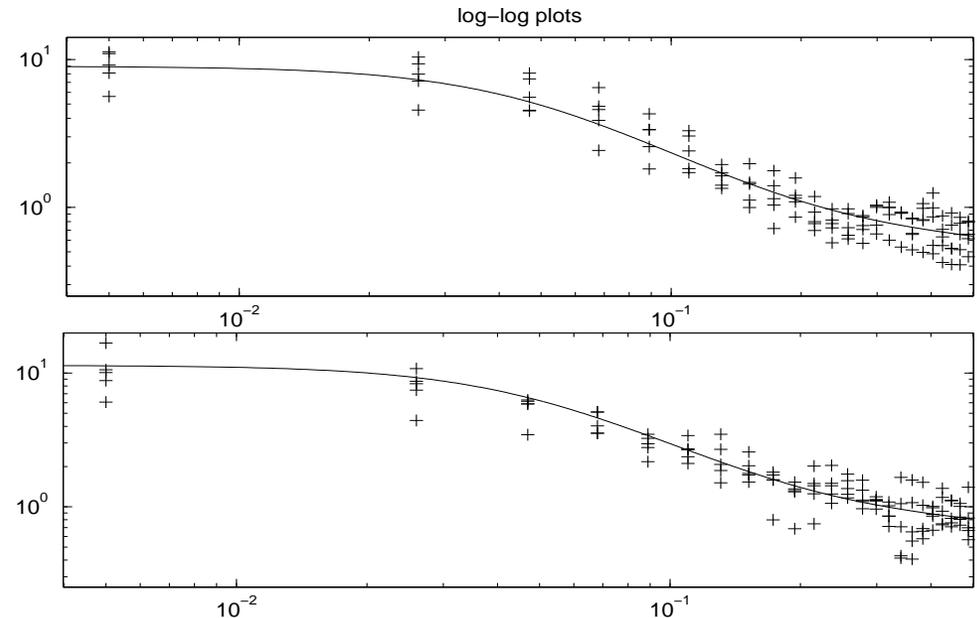
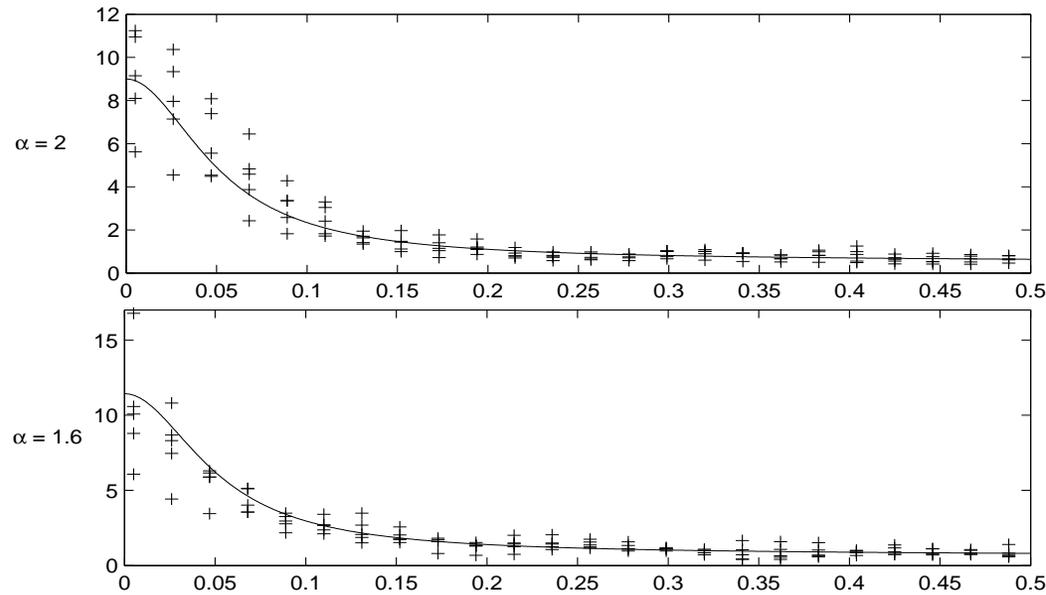
$$(D^2 + a_1 D + a_2) Y_t = (D + \mu) D L_t, \quad t \in \mathbb{R},$$

i.e. $a(z) = z^2 + a_1 z + a_2 = (z - \lambda_1)(z - \lambda_2)$ and $b(z) = z + \mu$.

In this case the normalized power transfer function can be written as

$$\left(\int_0^\infty f^2(s) ds \right)^{-1} \psi(\omega) = C(a_1, a_2, \mu) \cdot \frac{\omega^2 + \mu^2}{\omega^4 + (a_1^2 - 2a_2)\omega^2 + a_2^2}$$

$$\text{with } C(a_1, a_2, \mu) = \left(\int_0^\infty f^2(s) ds \right)^{-1} = 2 \frac{a_1 a_2}{\mu^2 + a_2}.$$



Smoothed periodogram values plotted against frequencies for five selected time series (pluses) in the Gaussian case (on top) and the 1.6-stable case (below). The true spectral density and normalized power transfer function is plotted as a solid line, respectively.

$$n = 15000; \Delta_n = 0.1; m = 300; m_n = \lfloor \sqrt{n\Delta_n} \rfloor = 38; W_n(k) = \frac{1}{2m_n+1} \text{ for } |k| \leq m_n.$$

	True	σ_L	a_1	a_2	μ
		1.5	2.0	0.1	0.2
$\alpha = 2$	Mean	1.5127	2.0859	0.1182	0.2159
	Bias	0.0127	0.0859	0.0182	0.0159
	Std. dev.	0.0392	0.1204	0.0358	0.0366
$\alpha = 1.8$	Mean	-	2.0580	0.1108	0.2185
	Bias	-	0.0580	0.0108	0.0185
	Std. dev.	-	0.1240	0.0372	0.0378
$\alpha = 1.6$	Mean	-	2.0626	0.1079	0.2127
	Bias	-	0.0626	0.0079	0.0127
	Std. dev.	-	0.1130	0.0315	0.0361
$\alpha = 1.4$	Mean	-	2.0659	0.1101	0.2129
	Bias	-	0.0659	0.0101	0.0129
	Std. dev.	-	0.1151	0.0311	0.0329

Simulation study for different values of α , based on 250 sample paths each.

$n = 15000$; $\Delta_n = 0.1$; $m = 300$; $m_n = \lfloor \sqrt{n\Delta_n} \rfloor = 38$; $W_n(k) = \frac{1}{2m_n+1}$ for $|k| \leq m_n$.

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- [1] FASEN, V. (2012). Limit theory for high frequency sampled MCARMA models. Submitted for publication.
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End

Thank you!