Moumanti Podder Joint work with Prof. Krishanu Maulik

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## Weighted Sums of Regularly Varying Random Variables With Dependent Weights

### Moumanti Podder Joint work with Prof. Krishanu Maulik

Indian Statistical Institute Kolkata

January 17, 2013

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Regularly Varying Distributions

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# $\{X_t\}$ independent and identically distributed, with regularly varying tails.

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# $\{X_t\}$ independent and identically distributed, with regularly varying tails.

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 $\{\Theta_t\}$  nonnegative random variables.

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 $\{X_t\}$  independent and identically distributed, with regularly varying tails.

 $\{\Theta_t\}$  nonnegative random variables.

 $\sum_{t} \Theta_t X_t$ 

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 $\{\Theta_t\}$  nonnegative random variables.

 $\sum \Theta_t X_t$ 

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Arises in

• Risk model proposed by Nyrhinen.

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 $\sum \Theta_t X_t$ 

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### Arises in

- Risk model proposed by Nyrhinen.
- As stationary solutions to stochastic recurrence equations.

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 $\{X_t\}$  independent and identically distributed, with regularly varying tails.

 $\{\Theta_t\}$  nonnegative random variables.

 $\sum \Theta_t X_t$ 

### Arises in

- Risk model proposed by Nyrhinen.
- As stationary solutions to stochastic recurrence equations.
- In random coefficient linear processes.

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### **Fundamental question:**

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### Fundamental question:

If  $X_t \in RV_{-\alpha}$ ,



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### Fundamental question:

If  $X_t \in RV_{-\alpha}$ , when is  $\sum_t \Theta_t X_t$  regularly varying?

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### Fundamental question:

If  $X_t \in RV_{-\alpha}$ , when is  $\sum_t \Theta_t X_t$  regularly varying? In particular, when  $\{\Theta_t\}$  and  $\{X_t\}$  are NOT independent?

## Deterministic weights

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•  $\{c_t\}$  positive real numbers.

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•  $\{c_t\}$  positive real numbers.

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•  $\{X_t\}$  i.i.d.  $RV_{-\alpha}$ .

## Deterministic weights

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{c<sub>t</sub>} positive real numbers.
{X<sub>t</sub>} i.i.d. RV<sub>-α</sub>.

•

 $\sum_{t=1}^{\infty} c_t^{\delta} < \infty \quad ext{for some } 0 < \delta < lpha \wedge 1.$ 

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• Then  $\sum_{t=1}^{\infty} c_t X_t$  regularly varying with index  $-\alpha$ .

## Random weights

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### • What if weights are random?

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## Random weights

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- What if weights are random?
- Have to consider

$$\sum_t \Theta_t X_t$$

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where  $\Theta_t$  random.

## Random weights

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- What if weights are random?
- Have to consider

$$\sum_t \Theta_t X_t$$

where  $\Theta_t$  random.

• First must study behaviour of product

 $\Theta_t X_t$ .

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### Theorem (Breiman; 1965)

### Nonnegative X with tail distribution $\overline{F} \in RV_{-\alpha}$ with $\alpha > 0$ .

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### Theorem (Breiman; 1965)

Nonnegative X with tail distribution  $\overline{F} \in RV_{-\alpha}$  with  $\alpha > 0$ . Nonnegative  $\Theta$  independent of X.

 $E(\Theta^{lpha+arepsilon})<\infty$ 

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for some  $\varepsilon > 0$ .

### Regularly Varying Distributions

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 $E(\Theta^{lpha+arepsilon})<\infty$ 

for some  $\varepsilon > 0$ . Then

 $P(\Theta X > x) \sim E(\Theta^{\alpha})P(X > x) \text{ as } x \to \infty,$ 

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 $E(\Theta^{lpha+arepsilon})<\infty$ 

for some  $\varepsilon > 0$ . Then

 $P(\Theta X > x) \sim E(\Theta^{lpha})P(X > x) \text{ as } x \to \infty,$ 

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hence  $P(\Theta X > x) \in RV_{-\alpha}$ .

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### Note:

 Extra moment condition beyond α-th moment on Θ, but only α-th moment of Θ appears in final result.

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 Extra moment condition beyond α-th moment on Θ, but only α-th moment of Θ appears in final result.

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• Independence between X and  $\Theta$  assumed.



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### Theorem (Resnick and Willekens; 1991)

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### Theorem (Resnick and Willekens; 1991)

Nonnegative  $\{X_t, t \ge 1\}, \{\Theta_t, t \ge 1\}$  independent of each other.

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### Theorem (Resnick and Willekens; 1991)

Nonnegative  $\{X_t, t \ge 1\}, \{\Theta_t, t \ge 1\}$  independent of each other.  $\{X_t\}$  i.i.d with  $\overline{F} \in RV_{-\alpha}, \alpha > 0$ .

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Nonnegative  $\{X_t, t \ge 1\}$ ,  $\{\Theta_t, t \ge 1\}$  independent of each other.  $\{X_t\}$  i.i.d with  $\overline{F} \in RV_{-\alpha}$ ,  $\alpha > 0$ . When  $0 < \alpha < 1$ , there exists  $\varepsilon > 0$  such that  $\alpha + \varepsilon < 1$  and

$$\sum_{t} E(\Theta_t^{\alpha \pm \varepsilon}) < \infty.$$

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$$\sum_t \mathsf{E}(\Theta_t^{\alpha\pm\varepsilon}) < \infty.$$

When  $\alpha \geq 1$ , there exists  $\varepsilon > 0$  such that

$$\sum_{t} [E(\Theta_t^{\alpha+\varepsilon})]^{\frac{1}{\alpha+\varepsilon}} < \infty \quad \text{and} \quad \sum_{t} [E(\Theta_t^{\alpha-\varepsilon})]^{\frac{1}{\alpha+\varepsilon}} < \infty.$$

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$$\sum_t \mathsf{E}(\Theta_t^{\alpha\pm\varepsilon}) < \infty.$$

When  $\alpha \geq 1$ , there exists  $\varepsilon > 0$  such that

$$\sum_{t} [E(\Theta_t^{\alpha+\varepsilon})]^{\frac{1}{\alpha+\varepsilon}} < \infty \quad \text{and} \quad \sum_{t} [E(\Theta_t^{\alpha-\varepsilon})]^{\frac{1}{\alpha+\varepsilon}} < \infty.$$

Then 
$$\lim_{x\to\infty} \frac{P[\sum_{t=1}^{\infty} \Theta_t X_t > x]}{P[X_1 > x]} = \sum_{t=1}^{\infty} E(\Theta_t^{\alpha}).$$

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### Note:

 Extra moment condition beyond α-th moment on Θ<sub>t</sub>, but only α-th moment of Θ<sub>t</sub> appears in final result.

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- Note:
  - Extra moment condition beyond α-th moment on Θ<sub>t</sub>, but only α-th moment of Θ<sub>t</sub> appears in final result..
  - Independence between  $\{X_t\}$  and  $\{\Theta_t\}$  still retained.

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### Theorem (Zhang, Shen and Weng; 2008)

{X<sub>t</sub>} i.i.d regularly varying with index -α. [No longer nonnegative].

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### Theorem (Zhang, Shen and Weng; 2008)

- {X<sub>t</sub>} i.i.d regularly varying with index -α. [No longer nonnegative].
- Negligible left tail of X<sub>t</sub>:

$$\frac{P[X_t < -x]}{P[X_t > x]} \to 0.$$

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•  $\{\Theta_t\}$  nonnegative, independent of  $\{X_t\}$ .

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- Negligible left tail of X<sub>t</sub>:

$$\frac{P[X_t < -x]}{P[X_t > x]} \to 0.$$

- $\{\Theta_t\}$  nonnegative, independent of  $\{X_t\}$ .
- $\{\Theta_t\}$  satisfies Resnick-Willekens conditions.

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### Theorem (Zhang, Shen and Weng; 2008)

- {X<sub>t</sub>} i.i.d regularly varying with index -α. [No longer nonnegative].
- Negligible left tail of X<sub>t</sub>:

$$\frac{P[X_t < -x]}{P[X_t > x]} \to 0$$

- $\{\Theta_t\}$  nonnegative, independent of  $\{X_t\}$ .
- {Θ<sub>t</sub>} satisfies Resnick-Willekens conditions.

Then  $\sum_{t=1}^{\infty} \Theta_t X_t^+$  and  $\sup_{1 \le n < \infty} \sum_{t=1}^n \Theta_t X_t$  are regularly varying with index  $-\alpha$ .

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### Theorem (Denisov and Zwart; 2007)

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## Theorem (Denisov and Zwart; 2007)

• Nonnegative X, with  $P[X > x] = \overline{F}(x) = x^{-\alpha}L(x), 0 < \alpha < \infty.$ 

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### Theorem (Denisov and Zwart; 2007)

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(DZ)  $\lim_{x\to\infty} \sup_{y\in[1,x]} L(y)/L(x) < \infty.$ 

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# Theorem (Denisov and Zwart; 2007)

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(DZ) 
$$\lim_{x\to\infty}\sup_{y\in[1,x]}L(y)/L(x)<\infty.$$

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• Nonnegative  $\Theta$  independent of X.

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(DZ) 
$$\lim_{x\to\infty} \sup_{y\in[1,x]} L(y)/L(x) < \infty.$$

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Nonnegative Θ independent of X.
E(Θ<sup>α</sup>) < ∞ and P{Θ > x} = o(P{X > x}).

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Nonnegative Θ independent of X.
E(Θ<sup>α</sup>) < ∞ and P{Θ > x} = o(P{X > x}). Then P(ΘX > x) ~ E(Θ<sup>α</sup>)P(X > x).

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### Theorem (Hazra and Maulik; 2012)

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### Theorem (Hazra and Maulik; 2012)

•  $\bar{F} \in RV_{-\alpha}$  with  $\lim_{x\to\infty} \sup_{y\in[1,x]} L(y)/L(x) < \infty$ .

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#### Introduction

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### Theorem (Hazra and Maulik; 2012)

- $\overline{F} \in RV_{-\alpha}$  with  $\lim_{x\to\infty} \sup_{y\in[1,x]} L(y)/L(x) < \infty$ .
- $\{\Theta_t, t \ge 1\}$  independent of  $\{X_t, t \ge 1\}$ ,

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- $\{\Theta_t, t \ge 1\}$  independent of  $\{X_t, t \ge 1\}$ ,

• 
$$P[\Theta_t > x] = o(P[X_t > x])$$

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### Theorem (Hazra and Maulik; 2012)

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- $\{\Theta_t, t \ge 1\}$  independent of  $\{X_t, t \ge 1\}$ ,
- $P[\Theta_t > x] = o(P[X_t > x])$
- For  $0 < \alpha < 1$ ,  $\sum_{t=1}^{\infty} E[\Theta_t^{\alpha}] < \infty$ ;

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- $\{\Theta_t, t \ge 1\}$  independent of  $\{X_t, t \ge 1\}$ ,
- $P[\Theta_t > x] = o(P[X_t > x])$
- For  $0 < \alpha < 1$ ,  $\sum_{t=1}^{\infty} E[\Theta_t^{\alpha}] < \infty$ ;
- For  $\alpha \geq 1$ ,  $\sum_{t=1}^{\infty} (E[\Theta_t^{\alpha}])^{\frac{1}{\alpha+\varepsilon}} < \infty$  for some  $\varepsilon > 0$ .

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- $\overline{F} \in RV_{-\alpha}$  with  $\lim_{x\to\infty} \sup_{y\in[1,x]} L(y)/L(x) < \infty$ .
- $\{\Theta_t, t \ge 1\}$  independent of  $\{X_t, t \ge 1\}$ ,
- $P[\Theta_t > x] = o(P[X_t > x])$
- For  $0 < \alpha < 1$ ,  $\sum_{t=1}^{\infty} E[\Theta_t^{\alpha}] < \infty$ ;
- For  $\alpha \geq 1$ ,  $\sum_{t=1}^{\infty} (E[\Theta_t^{\alpha}])^{\frac{1}{\alpha+\varepsilon}} < \infty$  for some  $\varepsilon > 0$ .

$$\mathsf{P}[\sup_{1 \le n < \infty} \sum_{t=1}^{n} \Theta_t X_t > x] \sim \mathsf{P}[X_{\infty} > x] \sim \mathsf{P}[X_1 > x] \sum_{t=1}^{\infty} \mathsf{E}[\Theta_t^{\alpha}]$$

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### Theorem (Hazra and Maulik; 2012)

- $\overline{F} \in RV_{-\alpha}$  with  $\lim_{x\to\infty} \sup_{y\in[1,x]} L(y)/L(x) < \infty$ .
- $\{\Theta_t, t \ge 1\}$  independent of  $\{X_t, t \ge 1\}$ ,
- $P[\Theta_t > x] = o(P[X_t > x])$
- For  $0 < \alpha < 1$ ,  $\sum_{t=1}^{\infty} E[\Theta_t^{\alpha}] < \infty$ ;
- For  $\alpha \geq 1$ ,  $\sum_{t=1}^{\infty} (E[\Theta_t^{\alpha}])^{\frac{1}{\alpha+\varepsilon}} < \infty$  for some  $\varepsilon > 0$ .

$$\mathsf{P}[\sup_{1 \le n < \infty} \sum_{t=1}^n \Theta_t X_t > x] \sim \mathsf{P}[X_\infty > x] \sim \mathsf{P}[X_1 > x] \sum_{t=1}^\infty \mathsf{E}[\Theta_t^\alpha]$$

and  $X_{\infty} = \sum_{t=1}^{\infty} \Theta_t X_t^+$  is almost surely finite.

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### Note

 Extra moment condition on Θ<sub>t</sub> no longer assumed beyond α-th moment.

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### Note

 Extra moment condition on Θ<sub>t</sub> no longer assumed beyond α-th moment.

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• But  $\{X_t\}$  and  $\{\Theta_t\}$  still assumed independent.

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## Note

 Extra moment condition on Θ<sub>t</sub> no longer assumed beyond α-th moment.

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• But  $\{X_t\}$  and  $\{\Theta_t\}$  still assumed independent.

## Next obvious question

• What if  $\{X_t\}$  and  $\{\Theta_t\}$  are not independent?

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# Definition (Sarmanov bivariate distributions)

(X, Y) follows bivariate Sarmanov distribution if:

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# Definition (Sarmanov bivariate distributions)

(X, Y) follows bivariate Sarmanov distribution if:

$$P(X \in dx, Y \in dy) = (1 + \theta \phi_1(x)\phi_2(y))F(dx)G(dy),$$

•  $x \in \mathbb{R}$  and  $y \ge 0$ ,

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$$P(X \in dx, Y \in dy) = (1 + \theta \phi_1(x)\phi_2(y))F(dx)G(dy),$$

• 
$$x \in \mathbb{R}$$
 and  $y \ge 0$ ,  $X \sim F$ ,  $Y \sim G$ ,

•  $\phi_1$  and  $\phi_2$  bounded, real-valued,

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 and  $y \geq$  0,  $X \sim F, Y \sim G$ ,

•  $\phi_1$  and  $\phi_2$  bounded, real-valued,

$$E\{\phi_1(X)\} = E\{\phi_2(Y)\} = 0,$$

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θ real constant,

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 and  $y \geq$  0,  $X \sim F, Y \sim G$ ,

•  $\phi_1$  and  $\phi_2$  bounded, real-valued,

 $E\{\phi_1(X)\} = E\{\phi_2(Y)\} = 0,$ 

### θ real constant,

 $1+\theta\phi_1(x)\phi_2(y)\geq 0.$ 

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# Theorem (Yang and Wang; 2012)

• (X, Y) bivariate Sarmanov distribution,

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# Theorem (Yang and Wang; 2012)

- (X, Y) bivariate Sarmanov distribution,
- $P[X > x] = \overline{F}(x) = x^{-\alpha}L(x)$ , with

 $\limsup_{x\to\infty}\sup_{1\leq y\leq x}L(x/y)/L(x)<\infty,$ 

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Theorem (Yang and Wang; 2012)

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•  $E(Y^{\alpha}) < \infty$ ,

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, with

 $\limsup_{x\to\infty}\sup_{1\leq y\leq x}L(x/y)/L(x)<\infty,$ 

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•  $\lim_{x\to\infty} \phi_1(x) = d_1$  exists, finite.

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# Theorem (Yang and Wang; 2012)

• (X, Y) bivariate Sarmanov distribution,

• 
$$P[X > x] = \overline{F}(x) = x^{-\alpha}L(x)$$
, with

 $\limsup_{x\to\infty}\sup_{1\leq y\leq x}L(x/y)/L(x)<\infty,$ 

• 
$$E(Y^{\alpha}) < \infty$$
,  
•  $\overline{G}(x) = o(\overline{F}(x))$ ,

•  $\lim_{x\to\infty} \phi_1(x) = d_1$  exists, finite.

 $P(XY > x) = \overline{H}(x) \sim \{E(Y^{\alpha}) + \theta d_1 E(\phi_2(Y)Y^{\alpha})\}\overline{F}(x).$ 

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### Lemma

(X, Y) bivariate Sarmanov, lim<sub>x→∞</sub> φ<sub>1</sub>(x) = d<sub>1</sub> exists, finite.

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### Lemma

- (X, Y) bivariate Sarmanov,  $\lim_{x\to\infty} \phi_1(x) = d_1$  exists, finite.
- $X^*, Y^*$  independent,  $X^* \sim F, Y^* \sim G$ .

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### Lemma

(X, Y) bivariate Sarmanov, lim<sub>x→∞</sub> φ<sub>1</sub>(x) = d<sub>1</sub> exists, finite.

- $X^*, Y^*$  independent,  $X^* \sim F, Y^* \sim G$ .
- $P[X^*Y^* > x] = \bar{H^*}(x) \in RV_{-\alpha}$  and  $\bar{G}(x) = o(\bar{H^*}(x)),$

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#### Lemma

- (X, Y) bivariate Sarmanov, lim<sub>x→∞</sub> φ<sub>1</sub>(x) = d<sub>1</sub> exists, finite.
- $X^*$ ,  $Y^*$  independent,  $X^* \sim F$ ,  $Y^* \sim G$ .
- $P[X^*Y^* > x] = \bar{H^*}(x) \in RV_{-\alpha}$  and  $\bar{G}(x) = o(\bar{H^*}(x)),$

Then  $P[XY > x] = \overline{H}(x) \sim P[X^*Y^*_{\theta} > x].$ 

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#### Lemma

- (X, Y) bivariate Sarmanov,  $\lim_{x\to\infty} \phi_1(x) = d_1$  exists, finite.
- $X^*$ ,  $Y^*$  independent,  $X^* \sim F$ ,  $Y^* \sim G$ .
- $P[X^*Y^* > x] = \overline{H^*}(x) \in RV_{-\alpha}$  and  $\overline{G}(x) = o(\overline{H^*}(x)),$

Then  $P[XY > x] = \overline{H}(x) \sim P[X^*Y^*_{\theta} > x].$ 

 $Y^*_{\theta}$ , independent of  $X^*$ , with

 $G_{\theta}(dy) = P[Y_{\theta}^* \in dy] = (1 + \theta d_1 \phi_2(y))G(dy).$ 

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### Theorem (Yang and Wang; 2012)

{(X<sub>n</sub>, Y<sub>n</sub>)}<sub>n∈ℕ</sub> *i.i.d* sequence, with generic random vector (X, Y) bivariate Sarmanov.

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- {(X<sub>n</sub>, Y<sub>n</sub>)}<sub>n∈ℕ</sub> *i.i.d* sequence, with generic random vector (X, Y) bivariate Sarmanov.
- $\overline{F}(x) = x^{-\alpha}L(x), \alpha > 0$ , and  $\overline{G}(x) = o(\overline{F}(x))$ .

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- {(X<sub>n</sub>, Y<sub>n</sub>)}<sub>n∈ℕ</sub> i.i.d sequence, with generic random vector (X, Y) bivariate Sarmanov.
- $\overline{F}(x) = x^{-\alpha}L(x), \alpha > 0$ , and  $\overline{G}(x) = o(\overline{F}(x))$ .
- $\limsup_{x\to\infty} \sup_{1\le y\le x} L(x/y)/L(x) < \infty$ .

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- {(X<sub>n</sub>, Y<sub>n</sub>)}<sub>n∈ℕ</sub> i.i.d sequence, with generic random vector (X, Y) bivariate Sarmanov.
- $\overline{F}(x) = x^{-\alpha}L(x), \alpha > 0$ , and  $\overline{G}(x) = o(\overline{F}(x))$ .
- $\limsup_{x\to\infty} \sup_{1\le y\le x} L(x/y)/L(x) < \infty.$
- $E(Y^{\alpha}) < \infty$ .

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- $E(Y^{\alpha}) < \infty$ .
- $\lim_{x\to\infty} \phi_1(x) = d_1$  exists, finite.

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- $E(Y^{\alpha}) < \infty$ .
- $\lim_{x\to\infty} \phi_1(x) = d_1$  exists, finite.

$$P[\sup_{1\leq m\leq n}\sum_{i=1}^{m}X_{i}\prod_{j=1}^{i}Y_{j}>x]\sim \frac{1-\{E(Y^{\alpha})\}^{n}}{1-E(Y^{\alpha})}[E(Y^{\alpha})+\theta d_{1}E(\phi_{2}(Y)Y^{\alpha})]\bar{F}(x).$$

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### Note

 No extra moment of Y<sub>t</sub> beyond the α-th moment assumed.

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### Note

 No extra moment of Y<sub>t</sub> beyond the α-th moment assumed.

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•  $\{X_t\}$  and  $\{Y_t\}$  no longer independent.

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### Note

- No extra moment of Y<sub>t</sub> beyond the α-th moment assumed.
- $\{X_t\}$  and  $\{Y_t\}$  no longer independent.
- Bivariate Sarmanov family includes many generalizations of Farlie-Gumbel-Morgenstern (FGM) distributions.

## Infinite sum case

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#### Theorem

### Same set-up as in finite sum case of Yang and Wang, with

 $E(Y_n^\alpha)<1.$ 

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#### Regularly Varying Distributions

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#### Theorem

Same set-up as in finite sum case of Yang and Wang, with

 $E(Y_n^{\alpha}) < 1.$ 

Then

$$P[\sup_{n\geq 1}\sum_{i=1}^{n}X_{i}\prod_{j=1}^{i}Y_{j}>x]\sim \frac{1}{1-E(Y^{\alpha})}[E(Y^{\alpha})+\theta d_{1}E\{\phi_{2}(Y)Y^{\alpha}\}]\bar{F}(x).$$

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### • Lower bound:

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### Lower bound:

Define  $S_n = \sum_{i=1}^n X_i \prod_{j=1}^i Y_j$ ,

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### Lower bound:

Define 
$$S_n = \sum_{i=1}^n X_i \prod_{j=1}^i Y_j$$
,  
 $\Psi(x, m) = P(\sup_{1 \le n \le m} S_n > x)$ , and

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### • Lower bound:

Define  $S_n = \sum_{i=1}^n X_i \prod_{j=1}^i Y_j$ ,  $\Psi(x, m) = P(\sup_{1 \le n \le m} S_n > x)$ , and  $\Psi(x) = P(\sup_{n \ge 1} S_n > x)$ .

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$$S_n = \sum_{i=1}^n X_i \prod_{j=1}^i Y_j$$
,  
 $\Psi(x, m) = P(\sup_{1 \le n \le m} S_n > x)$ , and  
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$$\Psi(x) \geq \Psi(x,m) = P\left[\sup_{1 \leq n \leq m} \sum_{i=1}^n X_i Y_i \prod_{j=1}^{i-1} Y_j > x\right].$$

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$$\Psi(x) \geq \Psi(x,m) = \boldsymbol{P}\left[\sup_{1 \leq n \leq m} \sum_{i=1}^{n} X_i Y_i \prod_{j=1}^{i-1} Y_j > x\right].$$

From Yang and Wang (2012),

 $\Psi(x,m) \sim \frac{1 - \{E(Y^{\alpha})\}^m}{1 - E(Y^{\alpha})} [E(Y^{\alpha}) + \theta d_1 E(\phi_2(Y)Y^{\alpha})]\overline{F}(x).$ 

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Let  $m \to \infty$  to get lower bound.

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### • Upper bound:

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• Upper bound:  
Define 
$$\Theta_i = \prod_{j=1}^{i-1} Y_j$$
 and  $Z_i = X_i Y_i$ .

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## • Upper bound: Define $\Theta_i = \prod_{j=1}^{i-1} Y_j$ and $Z_i = X_i Y_i$ .

For any  $m \in \mathbb{N}, 0 < \delta < 1$  and  $x \ge 0$ , we get

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### • Upper bound: Define $\Theta_i = \prod_{j=1}^{i-1} Y_j$ and $Z_i = X_i Y_i$ .

For any  $m \in \mathbb{N}$ ,  $0 < \delta < 1$  and  $x \ge 0$ , we get

$$P\left[\sup_{1 \le n < \infty} \sum_{t=1}^{n} \Theta_{t} Z_{t} > x\right] \le$$
$$P\left[\max_{1 \le k \le m} \sum_{t=1}^{k} \Theta_{t} Z_{t} > (1-\delta)x\right] + P\left[\sum_{t=m+1}^{\infty} \Theta_{t} Z_{t}^{+} > \delta x\right]$$

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• Upper bound: Define  $\Theta_i = \prod_{j=1}^{i-1} Y_j$  and  $Z_i = X_i Y_i$ . For any  $m \in \mathbb{N}, 0 < \delta < 1$  and  $x \ge 0$ , we get

$$P\left[\sup_{1 \le n < \infty} \sum_{t=1}^{n} \Theta_{t} Z_{t} > x\right] \le$$

$$P\left[\max_{1 \le k \le m} \sum_{t=1}^{k} \Theta_{t} Z_{t} > (1-\delta)x\right] + P\left[\sum_{t=m+1}^{\infty} \Theta_{t} Z_{t}^{+} > \delta x\right].$$

$$P\left[\max_{1 \le k \le m} \sum_{t=1}^{k} \Theta_{t} Z_{t} > (1-\delta)x\right] = \operatorname{Mr}((1-\delta)x,m).$$

• 
$$P\left[\max_{1\leq k\leq m}\sum_{t=1}^{k}\Theta_{t}Z_{t}>(1-\delta)x\right]=\Psi((1-\delta)x,m).$$

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### • Upper bound: Define $\Theta_i = \prod_{i=1}^{i-1} Y_i$ and $Z_i = X_i Y_i$ .

F(x)

For any  $m \in \mathbb{N}$ ,  $0 < \delta < 1$  and  $x \ge 0$ , we get

$$P\left[\sup_{1 \le n < \infty} \sum_{t=1}^{n} \Theta_t Z_t > x\right] \le$$

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$$P\left[\max_{1 \le k \le m} \sum_{t=1}^{k} \Theta_t Z_t > (1-\delta)x\right] = \Psi((1-\delta)x, m).$$

$$\lim_{x \to \infty} \frac{\Psi((1-\delta)x, m)}{\overline{F}(x)} =$$

 $(1-\delta)^{-\alpha} \frac{1-\{E(Y^{\alpha})\}^m}{1-E(Y^{\alpha})} E(Y^{\alpha}) + \theta d_1 E(\phi_2(Y)Y^{\alpha}).$ 

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### • Motivated by the principle of one large jump,

$$P\left[\sum_{t=m+1}^{\infty} \Theta_t Z_t^+ > x\right] \le \sum_{t=m+1}^{\infty} P\left[\Theta_t Z_t^+ > x\right] + P\left[\sum_{t=m+1}^{\infty} \Theta_t Z_t^+ \mathbf{1}_{\left[\Theta_t Z_t^+ \le x\right]} > x\right] = A + B.$$

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Motivated by the principle of one large jump,

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B should be small compared to A.

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#### Motivated by the principle of one large jump,

$$P\left[\sum_{t=m+1}^{\infty} \Theta_t Z_t^+ > x\right] \le \sum_{t=m+1}^{\infty} P\left[\Theta_t Z_t^+ > x\right] + P\left[\sum_{t=m+1}^{\infty} \Theta_t Z_t^+ \mathbf{1}_{\left[\Theta_t Z_t^+ \le x\right]} > x\right] = A + B.$$

*B* should be small compared to *A*. Get upper bound for *B* separately for  $\alpha < 1$  and  $\alpha \ge 1$ .

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•  $\alpha < 1$ ,

 $\frac{B}{\overline{F}(x)} \leq C \sum_{t=m+1}^{\infty} \frac{P[\Theta_t Z_t > x]}{\overline{F}(x)}.$ 

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α < 1,</li>

 $\alpha \geq$ 

$$\frac{B}{\overline{F}(x)} \le C \sum_{t=m+1}^{\infty} \frac{P[\Theta_t Z_t > x]}{\overline{F}(x)}.$$
  
$$\ge 1,$$
  
$$\frac{B}{\overline{F}(x)} \le \sum_{t=m+1}^{\infty} \frac{P[\Theta_t Z_t > x]}{\overline{F}(x)}.$$

$$rac{1}{D} \leq \sum_{t=m+1}^{\infty} rac{P[\Theta_t Z_t > x]}{\overline{F}(x)} + C[\sum_{t=m+1}^{\infty} (rac{P[\Theta_t Z_t > x]}{\overline{F}(x)})^{rac{1}{lpha + arepsilon}}]^{lpha + arepsilon}.$$

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$$\frac{P[\Theta_t Z_t > x]}{\overline{F}(x)} \leq B_t$$

for all large values of x, where

$$B_t = BE[\Theta_t^{\alpha}] = B\{E[Y_1^{\alpha}]\}^{t-1}$$

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$$\frac{P[\Theta_t Z_t > x]}{\overline{F}(x)} \leq B_t$$

for all large values of x, where

$$B_t = BE[\Theta_t^{\alpha}] = B\{E[Y_1^{\alpha}]\}^{t-1}$$

with

$$\sum_{t=m+1}^{\infty} B_t < \infty \quad \text{for } \alpha < 1$$

and

$$\sum_{t=m+1}^{\infty} B_t^{\frac{1}{\alpha+\varepsilon}} < \infty \quad \text{for } \alpha \ge 1.$$

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# Thank you $\ddot{-}$

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