

# Debt Maturity and Asymmetric Information: Evidence from Default Risk Changes\*

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## ABSTRACT

Asymmetric information models suggest that borrowers' choices of debt maturity depend on their private information about the probability of default. Borrowers with favorable information prefer short-term debt, and those with unfavorable information prefer long-term debt. We test this implication by tracing the evolution of debt issuers' default risk following issues of corporate debt classified by maturity. We find that short-term debt issuance leads to a decline in borrowers' asset volatility and an increase in their distance-to-default. The opposite is true for long-term debt issues. The results strongly suggest that borrowers' private information about their default risk determine the maturity of their debt issues.

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*Keywords:* Debt Maturity, Debt Issuance, Default Risk, Signaling, Timing, Distance-to-Default, Asset Volatility

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# I. Introduction

In this paper, we examine the predictions of asymmetric information models by tracing the evolution of default risk of firms conditional on debt maturity choices. Asymmetric information models, such as those by Flannery (1986) show that information asymmetries result in debt securities being mispriced in a way that varies with debt maturity.<sup>1</sup> Information theories model a borrower's willingness to subject its financing costs to new information as a tradeoff between the information effect of expecting that future news will be favorable, and the refinancing risk. For borrowers with favorable private information about future default risk changes, the market imputes a higher likelihood of credit quality deterioration than does the borrower. Consequently, borrowers that expect an improvement in the market perception of their credit quality will raise short-term debt to benefit from refinancing on favorable terms when their true credit quality is revealed to the market at a later date. Conversely, when borrowers have unfavorable private information about future default risk or when their true default risk is higher than that imputed by the market, they prefer long-term debt and thereby eliminate uncertainty about the future refunding costs or exposure to liquidity risks.

This raises the question whether rational market participants can reliably infer a firm's true quality from its maturity decisions and adjust the rates charged to compensate for differences in default risk. Flannery shows that with costless financial transactions, the debt market will have a pooling equilibrium and no borrower will choose long-term debt contract as these will be significantly more expensive. However, with positive transactions

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<sup>1</sup>Diamond (1991) also presents a model of debt maturity choice. In his model, some very low-rated borrowers who have no choice but to settle for short-term debt. These supply-side factors complicate inferences from the information models. However, as Diamond suggests, very low-rated borrowers with restricted access to public debt market most likely use short-term bank debt.

costs there would be a separating equilibrium in which good quality firms issue short-term debt and incur transactions costs of refinancing it.

If private information about default risk is guiding a firm's choice of its debt maturity, then we expect borrowers issuing short-term debt to exhibit a decline in perceived default risk, and those issuing long-term debt to exhibit an increase in perceived default risk. We test this key prediction of information models by examining how default risk measures evolve in the period following issues of corporate debt classified by maturity. Our tests focus on two market-based measures of a firm's default risk: asset volatility, which is directly related to a firm's default risk; and distance-to-default, which is inversely related to default risk.

Based on a sample of 4,089 debt issues for the period 1983-2003, we find that short-term debt issuers experience a significant decline in market perceptions of their default risk, i.e., their distance-to-default increases and asset volatility declines in the period following the short-term debt issue.<sup>2</sup> Firms that complete long-term debt issues exhibit the opposite pattern. These patterns of default risk changes around debt issues match the predictions of the information asymmetry theories. These changes in default risks are also economically large. Our tests control for initial debt ratings and key firm characteristics so our evidence of changes in default risk following debt issue cannot be attributed to changes in time-varying firm characteristics or ex ante risk information about the borrower as reflected in its risk ratings.

We provide further evidence on information theories by examining the effect of deviations from predicted maturity on future default risk changes. These tests are based

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<sup>2</sup>The call provision provides a firm with an early opportunity to refinance its debt. Robbins and Schatzberg (1986) argue that callable bonds can also signal a firm's better prospects in the presence of asymmetric information. Thus, in defining short- and long-term debt, we employ an adjusted maturity measure that replaces stated maturity of the bond with the first call date for callable bonds.

on the argument that debt maturity choices of issuers that are contrary to their type are more likely to be guided by private information. To do these tests, we first predict debt maturities of firms based on observable firm characteristics at the time of issue. We then examine future default risk changes of issuers classified by the deviation between predicted maturity and actual maturity. Firms that were predicted to issue short-term debt but actually issued long-term debt have a much larger deterioration in default risk measures relative to a broader population of long-term debt issuers. Conversely, firms that were predicted to issue long-term debt but actually issued short-term debt have a much larger improvement in default risk measures following debt issues.

While our evidence suggests that information considerations play an important role in a firm's choice of its debt maturity, there are concerns that alternative explanations may be consistent with our results. First, the agency theory suggests that firms issuing long-term debt will have incentives to engage in riskier investments and consequently increase the risk of their assets. Another possibility is that all information is observable and creditors know the expected default risk changes, and offer optimal contracts. Thus, creditors will offer long-term debt contracts to borrowers that are investing in longer-term fixed assets as they do not want to expose the borrowers to refinancing risk.

We test the agency view directly by examining the evolution of firm characteristics conditional on their debt maturity decision. We find no difference in investment policies of firms issuing debt of different maturities. The results show that both short- and long-term debt issuers exhibit similar evolution in tangibility, profitability, variability of operating cash, leverage, market/book assets, and investments in tangible and intangible assets in the two years following the debt issue. Thus, we do not find strong support for the agency view or the full-information view that long-term debt invest differently in the period following the issue. Importantly, we control for several key firm characteristics

that the agency theories indicate are important determinants of a firm's maturity choice. These tests show that changes in market perceptions of default risk of issuers are related to the deviation between actual and predicted maturity, even after controlling for time-varying firm characteristics and initial ratings of issuers.

The paper is organized as follows. In Section II we present a brief review of previous studies. We define our key variables in Section III. Data and univariate analysis are presented in Section IV. We show our key multivariate results in Section V. In Section VI we examine default risk changes for issuers whose debt maturity choices differ from those that are based on standard maturity models. Section VII concludes.

## II. Background

The literature on debt maturity typically tests information models of debt maturity choice in a cross-sectional setting. These tests relate debt maturity to risk ratings, and to variables reflecting the degree to which a firm's ex ante private information is favorable or unfavorable. The existing studies that examine the relation between risk ratings and debt maturity assume that firms' maturity choices allow creditors to infer some of what was previously firm-specific private information, and that creditors use this information in assigning risk ratings. Since the models predict that firms with high ratings (those with favorable information) will prefer to issue short-term debt and those with low ratings will prefer to issue long-term debt, the prediction in these papers is that debt maturity will be positively associated with debt ratings.

Barclay and Smith (1995) relate the maturity structure of existing debt to bond ratings. They find that lower-rated firms use more short-term debt than do higher-rated firms. For firms with rated debt, there is a strict monotonic relation between the ma-

turity of existing debt and bond rating. They also find that nonrated firms have more short-term debt. Since nonrated debt is mostly private debt, which usually has shorter maturity, it is unclear whether the nonmonotonicity is driven by factors other than a firm's credit risk. A study by Stohs and Mauer (1996), which uses bond ratings for publicly traded industrial firms, and a study by Scherr and Hulbert (2001), which uses an accounting measure of risk (Altman Z-score), both find similar evidence of nonmonotonic relation between risk ratings and debt maturity structure.<sup>3</sup> Because these studies focus on debt maturity structure, which reflects the stock of debt that has been built up over time, the researchers cannot distinguish the maturities of new debt from the remaining time on the stock of existing debt contracts. The maturity structure of the existing stock of debt reflects decisions made at different historical points and may not correspond with asymmetric information during the sample period.

Recognizing that asymmetric information models are more about the maturity of new debt at the time of origination and less about the maturity structure of existing debt, many studies relate the maturity of new debt issues to risk measures. These studies present conflicting results. Mitchell (1993) finds that issuers with higher bond ratings issue longer-maturity debt. But Guedes and Opler (1996) find that firms that are rated investment grade issue both shorter- and longer-term debt. Noninvestment grade firms issue intermediate-maturity debt. Berger, Espinosa-Vega, Frame, and Miller (2005) test information asymmetry models on a large sample of bank loans and find that the maturity of new loans to small businesses is positively related to risk ratings. Ortiz-Molina and Penas (2008) use an accounting measure of risk and find that firms rated as low risk issue longer-maturity debt.

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<sup>3</sup>A limitation of models that assess risk based on accounting models is that the information contained in these measures may be backward-looking so their desirability as measures of default risk is uncertain. In addition, accounting models do not consider volatility of a firm's assets in estimating its risk of default.

Many papers test the signaling models by including variables that reflect the degree to which a firm's ex ante information is favorable or unfavorable. The papers by Barclay and Smith (1995), Stohs and Mauer (1996), and Johnson (2003) include future abnormal earnings in debt maturity regressions. While the evidence reported in these papers is consistent with the predictions of information models, the economic magnitudes of these effects remain small. One explanation for these weak results is that the ex post variables, such as future abnormal earnings or stock returns (as in Guedes and Opler (1996)), are noisy measures of ex ante private information. In addition, there is a severe identification problem in these tests: firms with significant growth opportunities are likely to experience high earnings growth, and a random walk model of normal earnings will identify growth firms as experiencing positive future abnormal earnings.

As this brief review suggests, most of the previous studies have examined implications of asymmetric information models in a cross-sectional framework – relating debt maturity choices to risk ratings and its interaction with ex-ante information asymmetry measures. There is very little work done on testing the time-series predictions of the models. The exceptions are studies that test the models using post-issuance stock returns and ratings. Spiess and Affleck-Graves (1999) document long-run underperformance of firms following long-term debt issuance. Covitz and Harrison (1999) study the changes of firm default risk post debt issuance by examining expected rating migrations. They find that firms that expect to receive rating downgrades are more likely to issue debt with longer maturity.

However, several commentators claim that bond ratings are a noisy estimate of a firm's likelihood of default (see, for example, Vassalou and Xing (2004)). Ratings adjust slowly because rating agencies generally adopt a through-the-cycle approach, a policy that is aimed at avoiding excessive rating reversals. With this philosophy, rating agencies disregard short-term fluctuations in default risk. Ratings only partially adjust to the

actual level of the permanent component of default risk (see discussion in Cantor and Mann (2003), and Altman and Rijken (2004)). Furthermore, small changes in borrower’s financial default risk are unlikely to affect ratings because ratings follow a grid.

Instead, we track changes in perceived default risk following corporate issuances of debt of different maturities by using distance-to-default and asset volatility as measures of default risk. These risk measures improve upon the accounting measures of default risk such as such as the Z-score in Altman (1968), or the conditional logit model in Ohlson (1980), and debt ratings as has been done in previous studies. In the next section, we describe the construction of these measures.

### III. Key Variables

#### *A. Distance-to-default*

The distance-to-default, an inverse measure of a firm’s likelihood of default, has been used as a measure of default risk in several recent studies.<sup>4</sup> Our method closely follows that of Vassalou and Xing (2004) and is based on the contingent claims method of Black and Scholes (1973) and Merton (1974).

The procedure starts with an estimate of the volatility of equity ( $\sigma_E$ ) using daily data from the past 12 months. We use this as an initial value for the estimation of the volatility of assets ( $\sigma_A$ ). We then use the Black and Scholes formula to compute the value of assets ( $V_A$ ) for each trading day, using the market value of equity ( $V_E$ ) of that day. We use these daily values to compute the standard deviation of  $V_A$ , which we then use as an estimate for

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<sup>4</sup>See, for example, Vassalou and Xing (2004), Guner (2006) and Acharya, Bharath, and Srinivasan (2007).



the next iteration. We repeat this procedure until the  $\sigma_A$  from two consecutive iterations converge. With the converged value of  $\sigma_A$ , we can back out  $V_A$  through the Black-Scholes formula. Using a 12 month rolling estimation window, we can estimate the value of  $\sigma_A$  at the end of every month. In the Black-Scholes estimation, the risk-free rate is the one-year T-bill.

We then estimate the distance-to-default (DTD) as:

$$DTD_{i,t} = \frac{\ln(V_{A_{i,t}}/K_{i,t}) + \left(\mu_{i,t} - \frac{1}{2}\sigma_{A_{i,t}}^2\right)T}{\sigma_{A_{i,t}}\sqrt{T}} \quad (1)$$

where  $K_{i,t}$  is the book value of debt at time t, estimated as the short-term debt plus one-half of the long-term debt.<sup>5</sup> We estimate the drift,  $\mu_{i,t}$ , as the mean of the change in the natural logarithm of the value of assets.<sup>6</sup>

## B. Asset volatility

One concern with the use of distance-to-default is that short- and long-term debt issues might affect the future evolution of leverage differently for the two sets of issuers. If debt issues affect the evolution of leverage differently for short- and long-term debt issuers, then this may bias our inferences. Although our regression control for leverage changes, we

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<sup>5</sup>Crosbie and Bohn (2002) argue that it is important to include long-term debt for two reasons. First, firms need to service long-term debt and therefore interest payments are part of their short-term obligations. Second, a firm's ability to roll over its short-term debt depends on the size of its long-term debt.

<sup>6</sup>Based on the average ratio of half long-term plus short-term debt to book assets of 0.18, the average drift  $\mu_{i,t}$  of 0.10, the average asset volatility  $\sigma_A$  of 0.27, the average distance-to-default calculated based on book assets is  $DTD = \frac{\ln(1/0.18) + (0.10 - \frac{1}{2}0.27^2)1}{0.27\sqrt{1}} = 6.59$ . Several other studies use a simple approximation (e.g., Sundaram and Yermack (2007)) that estimates distance-to-default as  $\frac{V-K}{\sigma V}$ . This produces a slightly lower estimate of distance-to-default. Based on our assumed parameters, the simple approximation yields a distance-to-default of 3.04.

focus on a second measure of default risk that is unaffected by variation in leverage. This measure of default risk is the asset volatility,  $\sigma_A$ , which we estimate above in Equation 1. Debt issuers in our sample exhibit somewhat lower asset volatility than that reported for a broader population of Compustat firms by Vassalou and Xing (2004). In our view, this difference reflects the relatively better credit quality of public debt issuers.

### *C. Control Variables*

In default risk regressions, we control for leverage, firm size, profitability, asset tangibility, coefficient of variation of operating income ( $CV(OI)$ ), and growth opportunities. The variables are defined in Table I. We include leverage, since leverage has substantial effects on default probabilities. We control for firm size because larger firms are more diversified and are likely to have greater financial flexibility. Consequently, they face lower default risk. We control for profitability because profitable firms are considered less risky. Their higher margins contribute to internal equity, thus reducing their default risk. We include coefficient of variation of operating income because firms with greater income volatility may not be able to meet their fixed obligations and are generally considered more risky.

We also include the tangibility of assets, because tangible assets are easier for outsiders to value. Therefore, the asymmetric information problems are less severe when a firm's assets are mostly tangible. It is also difficult for managers to increase the risk of the firm when a firm has more tangible assets.

We also control for growth opportunities by including the ratio of market to book value of assets.<sup>7</sup> Higher growth firms have higher default risk. In addition, managers of high growth firms have a greater ability to increase the risk of assets.

In addition, macroeconomic variables that proxy for the variation in aggregate default risk over time are included in the tests. These include short-term interest rates and default spread. The level of short-term interest rates affects the aggregate level of default risk – credit risk is low when debt is issued in an environment of low interest rates. Similarly, default spread is a proxy for aggregated default risk. Debt issued during an environment when default spreads are generally high will have higher default risk, on average.

Recent studies show that the maturity of aggregate debt issues also varies with macroeconomic conditions (Kaplin and Levy (2001), Baker, Greenwood, and Wurgler (2002)). Kaplin and Levy (2001) show that the ratio of aggregate short-term debt to long-term debt varies countercyclically. There is pronounced increase in short-term debt during recessions. It could be explained by the transitory nature of the shock. Second, there is need to finance inventory buildup, in part due to reduction in sales. Thus, in our tests, we also examine the robustness of results after including indicator variables to identify recessionary periods and its interaction with debt maturity variables.

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<sup>7</sup>Adam and Goyal (2008) show that the market-to-book assets ratio has the highest information content with respect to investment opportunities.

# IV. Data, Descriptive Statistics and Univariate Results

## *A. Data and descriptive statistics*

Our sample, which we obtain from the Securities Data Company (SDC) New Issues database, comprises public straight debt issues by U.S. firms from 1983 to 2003.<sup>8</sup> We exclude debt issues with missing maturities and issue amounts, and debt issues by financial firms (6000-6999), financial leasing firms (7359), and utilities (4910-4940). We obtain financial statement data from Compustat. The daily market values and daily stock returns are from the CRSP daily files. We require issuers to be listed on both CRSP and Compustat in the year prior to the issue.

Table II shows the time-series and cross-sectional distribution of the sample debt issues. The sample consists of 4,089 debt offerings by 647 firms over the period from 1983 to 2003. Panel A shows that compared to the 1980s, debt issues are significantly more numerous in the 1990s and the early 2000s. Column (2) of Panel A reports the average issue amount (in constant dollars as at the year 2000) by year. Over the entire period, the issue size averages to about \$179 million. Column (3), which reports the stated debt maturity shows that the average stated debt maturity is about 12 years. As the table shows, stated maturities have declined in the more recent period.

Almost one third of our debt issues are callable, with call dates concentrating around five, seven, and ten years from the date of issuance. The call provisions provide firms with an opportunity to redeem their bonds at the first call date and effectively determine the earliest opportunity for a firm to refinance its existing debt (King and Mauer, 2000).

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<sup>8</sup>The sample ends in 2003 as we require three years of data after issuance.

Therefore, we use adjusted maturities that replace the maturity of callable bonds with time to first call.<sup>9</sup> The adjusted maturity averages to about eight years and shows a pattern similar to that of stated maturities.

Panel B shows debt characteristics classified by adjusted maturity. We classify debt as short-term if the adjusted maturity is less than or equal to three years, as medium-term if between three and seven years, and as long-term if it exceeds seven years. The studies by Barclay and Smith (1995) and Guedes and Opler (1996) use similar cutoffs. Short-term debt issues are larger in amount than are debt issues of longer maturity. Both short- and long-term debt issues have higher debt ratings and lower yield spreads compared to medium-term debt.

Having presented the descriptive statistics for the debt issue sample, we now construct a weighted average term-to-maturity for issuers that offer multiple debt securities in any given month where the weights reflect the amount issued. Our subsequent tests use this collapsed sample of 2,829 issues (more precisely “issuer-months”), where multiple issues by a firm in a given month are replaced by a single observation that aggregates these multiple issues.

Table III reports average borrower characteristics for different maturity classes in the period before the issue. Short-term debt issuers are larger compared to medium- and long-term debt issuers. In addition, short-term issuers have lower leverage, higher market-to-book ratios, lower tangibility of assets, higher profits, and lower debt ratings than do long-term debt issuers.

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<sup>9</sup>To check robustness, we replicate our results by using stated maturities, and find similar results.

### *B. Changes in default risk: plots*

To examine how default risk changes for short- and long-term debt issuers, we start by plotting the two default risk measures over the six-year period surrounding the debt issue. Figure 1(a) plots the average distance-to-default before and after debt issues by maturity classes. The plot shows that for the short-term issuers, the distance-to-default declines before the issue and increases substantially after the issue. By contrast, for long-term debt issuers, the distance-to-default shows no change before the issue and then a marked decline in the period following the issue.

Figure 1(b) plots the evolution of asset volatility. Consistent with the evidence from the earlier figure, the plot of asset volatility shows that for short-term debt issuers, asset volatility increases in the period before the issue and then declines significantly in the post-issue period. As expected, the pattern of asset volatility for long-term debt issuers is the reverse.

### *C. Changes in default risk: univariate results*

We present formal tests of the default risk differences before and after issuances for short- and long-term debt issuers in Table IV. The table presents average default risk measures in the period around debt issues. The column titled “-1” presents the average distance-to-default and asset volatility for the 12 month period before the issue month. The column titled “0” shows the average risk measures in the month of issuance. Columns titled “1” and “2” present averages for one and two years after issue, respectively. For short-term issuers, the distance-to-default increases from 6.69 in the month of the issuance to 7.24 two years after (the difference is significant with a p-value of zero). Consistent

with a decline in default risk for short-term debt issuers, we find that asset volatility declines during the two years following the issue (from 0.31 at the time of issuance to 0.29 in year 2). This difference is also statistically significant at the 1% level.

By contrast, for the long-term issuers, both of the default risk measures increase following issuance. The distance-to-default drops from 7.79 in the month of the issue to 7.28 two years later, and asset volatility increases from 0.26 to 0.27. Both of these changes are significant at the 1% level. Firms that issue medium-term debt also show a decline in asset volatilities but no change in distance-to-default following debt issue.

Overall, the time-series changes in default risk for short- and long-term debt issuers are consistent with the main predictions of the information models.

#### *D. Changes in financial condition of issuers surrounding debt issuances*

This section examines if ex post changes in firm characteristics differ for short- and long-term debt issuers. If debt maturity choices are motivated by agency considerations, we would expect to see long-term debt issuers investing relatively more in fixed assets, show declining profitability, greater variability of operating income, increasing leverage, and declining cash balances. According to the asymmetric information models, the underlying business risk should not change, only the market's perception of a firm's default risk changes. In other words, we would expect no relative differences in how firm characteristics evolve for short- and long-term debt issuers.

Table V presents the average values of profitability,  $CV(OI)$ , cash holdings, leverage, and market-to-book assets ratio, tangibility of assets, capital expenditures, total investments, and net investments. As the table shows, the changes in most firm characteristics

do not differ significantly across short- and long-term debt issuers. For example, profitability, market-to-book ratio, investments, and tangibility decline for both short- and long-term debt issuers following issuances. Similarly,  $CV(OI)$ , cash and leverage increase for both sets of issuers following issuances. Overall, the changes in default risk measures of short- and long-term debt issuers do not reflect the evolution of firm characteristics following debt issuances. Thus, we do not find strong support for the agency view in the data. The default risk changes that we documented earlier are more consistent with changing market perceptions of default risk of issuers, rather than changing underlying firm fundamentals.

## V. Regression Results

### A. Main results

Table VI examines the changes in distance-to-default and asset volatility around debt offerings. These tests are conducted on a panel where for each issuer-month in the sample we examine the risk measures in the year before the month of issuance, the month of the issue itself, and one year after the issue, and two-years after issue. Thus, we have four observations per issue. As indicated earlier, we use the collapsed issuance data where we replace multiple issues by a firm in a month with a single aggregated.<sup>10</sup>

The key variables of interest in Table VI are the time-period indicator variables that trace out changes in default risk measures from year -1 to year 2 relative to the offering month. For example,  $I_{-1}$  takes a value of one if the observation pertains to one year prior to debt issuance, and picks up the difference in default risk in the preceding 12 months,

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<sup>10</sup>There are no apparent concerns that survivorship bias affects our findings as we have about 2,696 debt issues on average over the four year period.



relative to its value in the offer month;  $I_{+1}$  takes a value of one if the observation pertains to the 12 months following the offer month, and picks up incremental default risk increase in the first year relative to the offer month;  $I_{+2}$  takes a value of one if the observation is for months 13 to 24 relative to the offer month, and picks up incremental default risk in the second year. The missing indicator is the offer month indicator ( $I_0$ ). In these tests, we control for firm size, market-to-book assets, leverage, profitability, asset tangibility, operating-income variability, and term structure variables, such as the Treasury rate and the spread of Baa bond yield over the one-year Treasury yield.

Our tests also control for debt ratings. The ratings reflect risk characteristics that rating agencies can observe, and control for credit risk of issuers at the time of issue. If ratings reflect some of the private information that issuers have about their future default risk changes, then our tests are decidedly conservative. Therefore, the three included time-period indicator variables ( $I_{-1}$ ,  $I_{+1}$ , and  $I_{+2}$ ) pick up changes in default risk that are not reflected in time-varying firm characteristics or in the debt ratings at the time of origination. We also include the industry indicator variables (based on Fama-French 38 industry classifications) to control for industry fixed effects. To account for multiple debt issues in a given year, we cluster standard errors at the issuer-year level and apply corrections for heteroscedasticity.

If default risk declines for short-term debt issuers and rises for long-term debt issuers, then the coefficient on  $I_{+1}$  and  $I_{+2}$  in the distance-to-default regressions should be positive when firms issue short-term debt and negative when they issue long-term debt. We also expect the coefficient on  $I_{+1}$  and  $I_{+2}$  in the asset volatility regressions to be negative when firms issue short-term debt and positive when they issue long-term debt.

We report the results for the distance-to-default regressions in Table VI, columns (1) to (3). The coefficient estimates on time-period indicators confirm our predictions. In

the two-year period following issue, both the short- and medium-term debt issuers show a significantly higher distance-to-default relative to the issuance month. By contrast, the long-term debt issuers show a large decline in the distance-to-default in the two-years following the issue.

The results for the asset volatility regressions reported in columns (4) to (6) are consistent with those reported earlier. The asset volatility declines significantly in the two-year period after issuance for the short-term debt issuers and increases significantly for long-term debt issuers.

The results on other control variables mostly confirm our expectations. Firm size negatively affects asset volatility and positively affects distance-to-default. Firms with higher market/book assets ratio have higher asset volatility. Leverage negatively affects both distance-to-default and asset volatility. Profitability is positively related to distance-to-default and negatively related to asset volatility. Tangibility also increases distance-to-default and negatively affects asset volatility. The variability of income has no effect on either measure of default risk. The coefficient estimates on interest rate variables suggest that when Treasury rate and the credit spread are higher, default risk rises. The coefficient estimates on rating indicator variables are not reported separately in the table but the results confirm that as ratings worsen, distance-to-default declines and asset volatility increases.

### *B. Alternative Estimations*

We test the robustness of the results by re-estimating default risk regressions in changes. The advantage of regressions in changes is that they remove the effects of time-invariant omitted firm characteristics. These results are reported in Table VII. The

dependent variable in column (1) is the change in distance-to-default between the issue month and 24 months following the issue. The dependent variable in column (2) is the change in asset volatility over the same period. These regressions also include changes in firm characteristics measured over the same period. The key variables in these regressions are the indicator variables for short- and long-term debt.

In column (1), the coefficient on short-term debt indicator is positive and significant at the 1% level. Relative to medium-term debt, the benchmark category, firms that issue short-term debt have the change of distance-to-default that is higher by 0.90. Since the average distance-to-default is about seven, the increase appears to be economically significant. Also consistent with earlier results, we find that the coefficient on long-term debt is significantly negative. In terms of changes in firm characteristics, only leverage changes appear to be related to changes in distance-to-default. A decline in leverage leads to a significant increase in distance-to-default. Changes in Treasury spreads and Baa spreads also affect changes in the distance-to-default.

In column (2), we report the results from regressions on the changes in asset volatility. The coefficient estimates suggest that the change in asset volatility for short-term issuers is 1.9% lower compared to that for medium-term debt issuers. But for long-term issuers, the change in asset volatility is 2.2% higher compared to that for medium-term issuers. These results are consistent with a decline in default risk for short-term debt issuers and an increase in the risk for long-term debt issuers.

Column (3) reports results from probit estimates in which the dependent variable takes a value of one if the change in distance-to-default is positive over the two periods following the debt issue, and zero otherwise. Consistent with the OLS results reported in column (1), the coefficient on short-term debt is positive and statistically significant at the 1% level, while the coefficient on long-term debt is negative but not significant at

the conventional levels. In column (4), the dependent variable is an indicator variable that takes a value of one if asset volatility increases in the 24-month period following debt issuance. The probit estimates reported in column (4) suggest that asset volatility declines for short-term debt issuers and increases for long-term debt issuers. The F-test statistics on the equality of short- and long-term debt dummies are significant at the 1% level in all four regressions, suggesting that default risk changes following short-term debt issues differ from those around long-term debt issues.

In columns (5) to (8) of Table VII, we control for macroeconomic cycles by including a recession dummy variable that is equal to one if the debt security is issued during NBER recession years. In these tests, we also include interaction variables between various debt maturity indicators and the NBER recession indicator variable. We find that in recession years, default risk declines following both short- and long-term debt issues. During non-recessionary periods, short-term debt issuers experience a reduction while long-term debt issuers experience an increase in default risk after debt issuance.

Our results are robust to how we classify the debt issues. In unreported results, we redefine short-term debt as debt with a maturity less than or equal to five years, medium-term debt as debt with maturity between five and ten years, and long-term debt as debt with a maturity above ten years. The results remain unchanged. We also examine how robust our results are to our definition of modified maturity. We use stated maturity instead of the adjusted maturity and find qualitatively identical results. We also define adjusted maturity as the average between the bond maturity and the number of years of call protection. Again, this change had no material effect on our findings.

Taken together, the results in Tables VI and VII are consistent with the predictions of Flannery (1986). These results show that default risk falls following short-term debt issues and rises after long-term debt issues.

## VI. Predicted Compared to Actual Maturity

The debt maturity literature shows that debt maturity choices are related to observable firm characteristics. If firms have preferences in terms of debt maturity based on their observable firm characteristics, then the signaling implications will be relatively more significant when a firm issues debt of a maturity that is against its type.

Thus, we examine if the changes in default risk subsequent to issuance are greater when a firm issues debt of a maturity that is contrary to the maturity that is expected of it, based on its characteristics. We start by focusing on two firm characteristics that have been shown to affect debt maturity in the previous literature - firm size and growth opportunities. According to the debt maturity literature, small and high-growth firms are more likely to borrow short-term debt while large and low-growth firms are more likely to borrow long-term. Thus, we expect default risk changes to be significantly larger when, for example, large firms and low-growth firms issue short-term debt or when small and high-growth firms issue long-term debt. In these cases, it is likely that the firm's maturity choice is guided by its private information about its default risk. We define large firms as those with assets greater than the sample's median, and small firms as those with assets that are below the sample's median. We define high-growth firms as those with market/book assets ratios greater than the median for the sample. We classify the remaining firms as low-growth firms.

In unreported tables, we find that default risk measures exhibit a relatively larger increase when small firms issue long-term debt compared to an average long-term debt issuer. Conversely, when large firms issue short-term debt, the default risk measures show a larger drop compared to average an average short-term debt issuer. The results are consistent when we classify firms based on growth opportunities.

We extend these tests by including other determinants of debt maturity choice in a multivariate setting. The models of debt maturity such as those by Barclay and Smith (1995), Guedes and Opler (1996), and Stohs and Mauer (1996) suggest that debt maturity is a function of leverage, market-to-book assets, firm size and firm-size squared, asset maturity, abnormal earnings, coefficient of variation of operating income, term spread, an indicator variable for regulatory firms, and rating indicator variables. We therefore estimate debt maturity as a function of these variables and present the results in the appendix. Consistent with findings in other studies, the maturity increases with firm size and asset maturity and decreases with the market-to-book assets ratio, abnormal earnings, and  $CV(OI)$ . Firm size has a nonlinear effect on maturity.

Using these estimates, we predict the maturity choices of issuers in our sample and compare those with actual maturities chosen by the sample firms. Table VIII presents the average distance-to-default and asset volatility for the four groups of issuers, based on predicted and actual maturities. When the data predict that firms will issue short-term debt but instead issue long-term debt, the distance-to-default declines significantly from 9.651 in the year before the issue to 8.102 in the second year after the issue (the p-value for the change is zero). For this group of issuers, the asset volatility increases from 0.283 to 0.310 (the increase has a p-value of 0.03). When we predict that a firm will issue long-term debt and they do so, the default risk also declines but the decline is smaller in magnitude compared to those for cases when the predicted maturity was short.

We also find that when firms are predicted to issue long-term securities but they instead issue short-term debt, the distance-to-default increases significantly from 7.144 in the year before the debt issue to 7.585 in the second year after the offer (p-value equals 0.04). The asset volatility in these cases declines from 0.289 in the year before the offer to 0.275 in the second year after the offer (the p-value for the difference equals 0.01).

Consistent with our conjecture, the differences in distance-to-default is small and non-significant when firms predicted to issue short-term debt do in fact issue short-term debt but the results show a significant reduction in asset volatility.

Overall, when issuers issue debt with actual maturity that is different from the predicted choice, they experience a relatively larger changes in default risk compared to the scenarios when the actual maturity matches the predicted maturity.

Table IX confirms these findings in a multivariate setting. We construct a new variable *Actual – Predicted* which is an ordinal variable that takes a value between -2 and +2. This variable is the difference between actual maturity choice (1=short-term, 2=medium-term, and 3=long-term) and predicted maturity choice. The predicted maturity is estimated from a prediction model presented in the Appendix and it also takes three value (1=short-term, 2=medium-term, and 3=long-term). Higher values of *Actual – Predicted* suggest that actual debt was of longer maturity while the firm was predicted to issue shorter-maturity, and vice versa. The table shows that after controlling for changes in firm characteristics and interest rates, the difference between actual and predicted maturity choice is strongly related to issuer default risks in the next two years. When actual debt maturity is longer than that predicted, distance-to-default declines and asset volatility increases suggesting a significant increase in market perceptions of the firm’s default risk in those cases. Conversely, when actual debt maturity is shorter than predicted, there is a substantial decline in the perceived default risk.

## VII. Conclusion

In this paper we test the extent to which information asymmetry plays a role in firms’ debt maturity choices. We examine changes in the market-based default risk characteris-

tics of debt issuers conditional on their maturity choice. With asymmetric information, a firm with favorable private information about its default risk will find that the market's default risk premia are excessive. These distortions are greater for long-term debt. Therefore, because they expect to roll over this debt at a price that reflects what the firm's future condition will be at the time of refinancing. Firms with favorable private information prefer short-term debt. For the opposite reason, firms with unfavorable private information prefer long-term debt. The key testable implication of these models is that firms issuing short-term debt will have favorable private information about their default risk, and firms issuing long-term debt will have unfavorable private information. This implication leads to the prediction that short-term debt issuers will exhibit an improvement in their perceived default risk while long-term debt issuers will show deterioration in their perceived default risk.

Focusing on two market-based default risk measures, asset volatility and distance-to-default, we examine how default risk measures change after debt issues. Our results show that long-term issuers experience a significant increase in perceived default risk, and that short-term debt issuers experience a significant improvement in the period immediately following the debt issue.

We also examine issuers whose debt maturity choices are different from those predicted from standard maturity models. We find significant default risk declines for short-term debt issuers that were predicted to issue long-term debt. This decline is larger for this group of issuers compared to a broader population of short-term debt issuers. We also find that default risk increases for long-term debt issuers that were predicted to issue short-term debt. Again, this increase is larger for this group of issuers compared to that for a broader population of long-term debt issuers. Overall, our evidence strongly supports the predictions of the asymmetric information models of debt maturity choice.

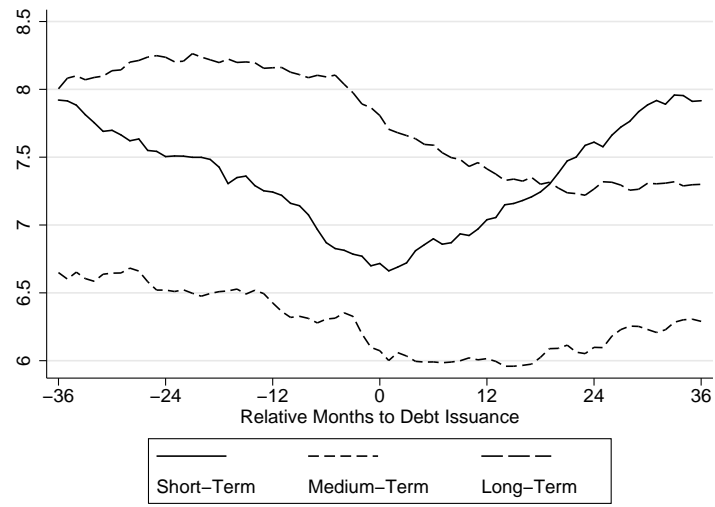


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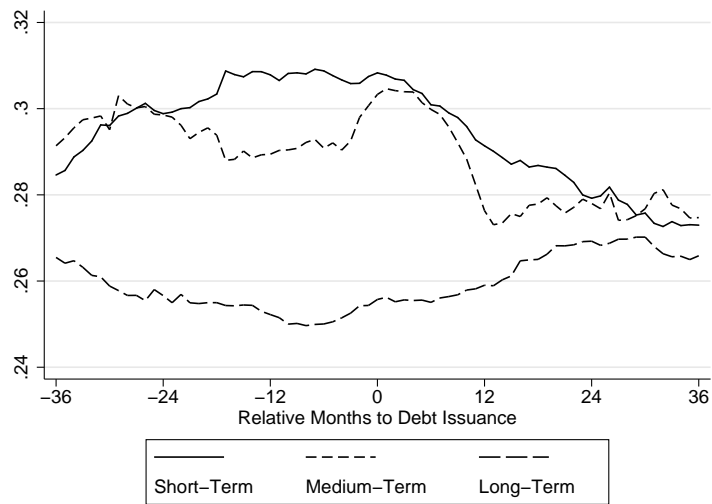
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(a) Distance-to-default by adjusted maturity



(b) Asset volatility by adjusted maturity

**Figure I. Issuer Distance-to-Default and asset volatility by Maturity.**

**Table I**  
**Variable Definitions**

Variable	Definition
Distance-to-Default	Distance-to-default or $DTD_{i,t} = \frac{\ln(V_A/K) + (r - \frac{1}{2}\sigma_A^2)T}{\sigma_A\sqrt{T}}$ , where $V_A$ is the value of assets, $K$ is the book value of debt, $r$ is the one-year T-Bill rate, $\sigma_A$ is the volatility of asset, and $T$ is the maturity of debt.
$\sigma_v$	The volatility of asset values is estimated from the Black-Scholes model through an iterative procedure.
Firm size	The natural logarithm of assets (item 6), where assets are deflated to constant 2000 dollars.
Leverage	Total Debt/Market Value of Assets: the ratio of book value of debt (item 9 + item 34) to MVA, market value of assets. MVA is obtained as the sum of market value of equity (item 199, price-close item 54, shares outstanding) + debt in current liabilities (item 34) + long-term debt (item 9) + preferred- liquidation value (item 10) - deferred taxes and investment tax credit(item 35).
Market/book assets	Estimated as the ratio of market value of assets (MVA) to assets (item 6).
Profit	Profitability is the ratio of operating income before depreciation (item 13) to assets (item 6).
Tangibility	Defined as the ratio of net property, plant and equipment (item 8) to total assets (item 6).
CV(OI)	CVOI is the coefficient of variation of operating income (item 13) measured over a three-year period using annual income statement data.
Rating	Long-term issuer ratings from S&P's (from Compustat). Numerical values are mapped to credit rating in the following way: AAA=20, AA+=19, AA=18, AA-=17, A+=16, A=15, A-=14, BBB+=13, BBB=12, BBB-=11, BB+=10, BB=9, BB-=8, B+=7, B=6, B-=5, CCC+=4, CCC=3, CCC-=2, CC/C=1.
Treasury yield	Yield on a one year Treasury bill (Source: <a href="http://www.federalreserve.gov/release">http://www.federalreserve.gov/release</a> )
Baa spread	Estimated as the difference between yield on Baa-rated bonds and Aaa-rated bonds (Source: <a href="http://www.federalreserve.gov/release">http://www.federalreserve.gov/release</a> )

**Table II**  
**Average Stated and Adjusted Debt Maturities by Year**

Panel A reports the frequency of debt issues during the 1983 to 2003 period. It also reports the average principal amount, stated and adjusted maturity in each of the years during the sample period. Issue amount is expressed as constant 2000 dollars. Stated maturity is the debt maturity stated in the offering prospectus at the time of bond issue. Adjusted maturity adjusts debt maturity to the call start date for bonds that are callable. Panel B presents the debt characteristics grouped by adjusted maturity. Debt issue rating is the S&P bond rating, taken from SDC. Yield spread is the difference between bond yield and comparable Treasury bond yield.

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*Panel A: Frequency of Debt Issues*

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Year	N	Principal (\$million)	Stated Maturity	Callable (%)	Adjusted Maturity
1983	59	134.9	15.4	76.3	6.9
1984	50	167.0	13.2	72.0	6.8
1985	87	190.1	14.6	80.5	6.8
1986	147	229.4	16.2	67.4	8.5
1987	99	194.5	14.0	57.6	6.9
1988	67	200.0	12.6	71.6	6.8
1989	95	220.8	12.6	54.7	7.1
1990	97	200.0	9.0	12.4	8.3
1991	236	173.9	12.1	5.1	11.5
1992	222	197.5	12.0	24.3	10.0
1993	241	185.0	15.8	24.9	13.3
1994	163	141.7	9.4	27.6	7.1
1995	225	137.5	12.6	22.2	10.1
1996	289	143.2	13.4	20.1	11.4
1997	389	134.1	15.4	20.6	9.7
1998	505	137.5	12.7	28.1	8.5
1999	216	186.7	10.3	37.5	5.5
2000	164	236.9	7.8	34.8	5.0
2001	257	242.6	9.7	66.2	4.1
2002	294	190.2	8.6	52.0	4.4
2003	187	260.8	9.7	72.2	3.4
Total	4,089	179.1	12.2	37.1	8.1

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**Table II Continued**

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*Panel B: Sorts on firm profitability*

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	Short-term debt	Medium-term debt	Long-term debt
	(N=1,168)	(N=1,094)	(N=1,827)
Principal amount (in \$millions)	214.618	156.843	169.688
Principal amount/book assets	0.046	0.090	0.033
Debt issue rating	14.311	12.717	14.437
Yield spread	1.460	1.855	1.080

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**Table III: Issuer Characteristics by Debt Maturity**

The sample of debt issues is collapsed to construct a weighted average term-to-maturity for issuers that offer multiple debt securities in any given month where the weights reflect the amount issued. This table uses this collapsed sample of 2,829 issues (more precisely “issuer-months”). The table presents average values of lagged characteristics of debt issuers grouped by the adjusted maturity of their debt offering. The variables are defined in the Table I. Columns (4) to (6) report p-values from two-tailed tests of the null hypotheses of no differences in firm characteristics across debt maturities.

	Short-term debt (N=783)	Medium-term debt (N=751)	Long-term debt (N=1295)	Long- short (p-value)	Long- medium (p-value)	Medium- short (p-value)
	(1)	(2)	(3)	(4)	(5)	(6)
Assets (in \$billions)	16.800	12.560	12.610	0.00	0.96	0.00
Market/book assets	1.587	1.291	1.287	0.00	0.92	0.00
Leverage	0.281	0.380	0.301	0.01	0.00	0.00
Profit	0.163	0.152	0.157	0.04	0.09	0.00
Tangibility	0.398	0.424	0.468	0.00	0.00	0.02
CV(OI)	0.944	0.971	0.946	0.85	0.00	0.01
Rating	14.170	13.144	14.488	0.03	0.00	0.00



**Table IV: Default Risk Changes around Debt Issues**

This table reports the average values of distance-to-default and asset volatility ( $\sigma_A$ ) in a four year window surrounding the debt issues. We classify the debt issues as classified as short-term if the adjusted maturity is less than three years. The issues are classified as medium-term if the adjusted maturity is between three and seven years, and as long-term if the adjusted maturity is more than seven years. Column (5) and (6) report p-values from a two-tailed t-test of the null hypothesis that default risk measures between year2 and year0, and between year2 and year-1 are similar.

Debt Maturity	Default risk measure	Year relative to offer					
		-1	0	1	2	Yr.2 - Yr.0 p-value	Yr.2 - Yr.-1 p-value
		(1)	(2)	(3)	(4)	(5)	(6)
Short-term (N=783)	Distance-to-default	6.987	6.685	6.831	7.241	0.00	0.18
	$\sigma_A$	0.308	0.308	0.302	0.288	0.00	0.00
Medium-Term (N=751)	Distance-to-default	6.315	6.056	6.010	5.996	0.76	0.10
	$\sigma_A$	0.293	0.303	0.297	0.280	0.00	0.05
Long-Term (N=1295)	Distance-to-default	8.072	7.792	7.550	7.280	0.00	0.00
	$\sigma_A$	0.252	0.255	0.256	0.267	0.00	0.00

**Table V**  
**Changes in Financial Variables**

This table reports the time-series averages of profitability,  $CV(OI)$ , cash, leverage, the market/book assets ratio, tangibility, capital expenditure, total investment, and net investment around debt issues from annual balance sheet, income statements, and cash flow statements. We define total investment as the sum of capital expenditure (item 128), acquisitions (item 129), and increase in investment (item 113)). We define net investment as total investment plus other use of funds (item 219), minus sale of  $PPE$  (item 107), and minus sale of investment (item 109). Capital expenditure, acquisitions, increase in investment, other use of funds, sale of  $PPE$ , and sale of investment as zero if they were missing. We classify the debt issues as short-term if the adjusted maturity is less than three years, and as long-term if the adjusted maturity is more than seven years.

	Year relative to offer					
	-1	0	1	2	Yr.2- Yr.0 p-value	Yr.2-Yr.- 1 p-value
<i>Panel A: Profitability</i>						
Short-term issuer	0.164	0.158	0.155	0.151	0.00	0.00
Long-term issuer	0.157	0.150	0.148	0.147	0.05	0.00
<i>Panel B: CV (OI)</i>						
Short-term issuer	0.944	0.945	0.954	0.960	0.00	0.00
Long-term issuer	0.946	0.949	0.962	0.970	0.00	0.00
<i>Panel C: Cash</i>						
Short-term issuer	0.044	0.046	0.049	0.053	0.00	0.00
Long-term issuer	0.039	0.041	0.042	0.043	0.19	0.01
<i>Panel D: Leverage</i>						
Short-term issuer	0.281	0.302	0.307	0.305	0.37	0.00
Long-term issuer	0.301	0.321	0.326	0.329	0.03	0.00
<i>Panel E: Market/Book asset</i>						
Short-term issuer	1.587	1.545	1.511	1.509	0.16	0.01
Long-term issuer	1.287	1.294	1.282	1.283	0.56	0.60

**Table V Continued**

	Year relative to offer				Yr.2- Yr.0 p-value	Yr.2-Yr.- 1 p-value
	-1	0	1	2		
<i>Panel F: Tangibility</i>						
Short-term issuer	0.397	0.393	0.389	0.382	0.00	0.00
Long-term issuer	0.468	0.464	0.457	0.452	0.00	0.00
<i>Panel G: Capital expenditure</i>						
Short-term issuer	0.073	0.069	0.064	0.059	0.00	0.00
Long-term issuer	0.088	0.083	0.078	0.074	0.00	0.00
<i>Panel H: Total Investments</i>						
Short-term issuer	0.120	0.114	0.099	0.096	0.00	0.00
Long-term issuer	0.124	0.127	0.115	0.112	0.00	0.00
<i>Panel H: Net Investments</i>						
Short-term issuer	0.098	0.097	0.077	0.073	0.00	0.00
Long-term issuer	0.107	0.111	0.090	0.078	0.00	0.00

**Table VI: Default Risk Changes and Debt Maturity**

The table reports regression results from specifications that regress the two default risk measures on indicator variables for the period before and after debt issue, firm characteristics, and interest rate variables. The accounting variables are from the fiscal year that ends right before the period in which distance-to-default and asset volatility are measured. The regressions include industry fixed effects and rating indicators (the coefficients are suppressed to save space). Numbers in parentheses are standard errors adjusted for heteroscedasticity and they are clustered at the issuer level. <sup>a</sup>, <sup>b</sup> and <sup>c</sup> denote significance at the 1%, 5%, and 10% levels, respectively.

	<i>Distance-to-default</i>						<i>Asset Volatility</i>					
	Short-term debt (1)	Medium-term debt (2)	Long-term debt (3)	Short-term debt (4)	Medium-term debt (5)	Long-term debt (6)	Short-term debt (1)	Medium-term debt (2)	Long-term debt (3)	Short-term debt (4)	Medium-term debt (5)	Long-term debt (6)
$I_{-1}$	0.664 <sup>a</sup> (0.163)	0.471 <sup>a</sup> (0.147)	0.459 <sup>a</sup> (0.118)	-0.009 (0.006)	-0.022 <sup>a</sup> (0.007)	-0.009 <sup>a</sup> (0.003)						
$I_{+1}$	0.448 <sup>a</sup> (0.172)	0.254 <sup>c</sup> (0.151)	-0.138 (0.116)	-0.013 <sup>b</sup> (0.006)	-0.021 <sup>a</sup> (0.007)	0.004 (0.004)						
$I_{+2}$	0.910 <sup>a</sup> (0.238)	0.329 <sup>c</sup> (0.169)	-0.240 <sup>c</sup> (0.134)	-0.019 <sup>a</sup> (0.007)	-0.017 <sup>b</sup> (0.007)	0.016 <sup>a</sup> (0.004)						
Firm size	-0.180 <sup>b</sup> (0.087)	0.329 <sup>a</sup> (0.057)	0.130 <sup>c</sup> (0.068)	-0.010 <sup>b</sup> (0.004)	-0.018 <sup>a</sup> (0.003)	-0.006 <sup>a</sup> (0.002)						
Market/book assets	-0.067 (0.129)	0.160 (0.169)	-0.094 (0.115)	0.027 <sup>a</sup> (0.004)	0.013 <sup>a</sup> (0.005)	0.026 <sup>a</sup> (0.004)						
Leverage	-8.772 <sup>a</sup> (0.648)	-7.846 <sup>a</sup> (0.476)	-8.497 <sup>a</sup> (0.507)	-0.127 <sup>a</sup> (0.027)	-0.173 <sup>a</sup> (0.021)	-0.156 <sup>a</sup> (0.019)						
Profit	5.617 <sup>b</sup> (2.643)	6.020 <sup>a</sup> (1.345)	8.527 <sup>a</sup> (1.813)	-0.384 <sup>a</sup> (0.103)	-0.230 <sup>a</sup> (0.075)	-0.277 <sup>a</sup> (0.069)						

Table VI Continued

	<i>Distance-to-default</i>			<i>Asset Volatility</i>		
	Short-term debt (1)	Medium-term debt (2)	Long-term debt (3)	Short-term debt (4)	Medium-term debt (5)	Long-term debt (6)
Tangibility	1.768 <sup>a</sup> (0.544)	2.080 <sup>a</sup> (0.431)	1.116 <sup>b</sup> (0.452)	-0.045 <sup>b</sup> (0.019)	-0.049 <sup>a</sup> (0.019)	-0.021 (0.013)
CV (OI)	0.629 (0.964)	-0.193 (0.497)	-0.307 (0.682)	-0.045 (0.033)	0.043 (0.039)	-0.007 (0.029)
Treasury	0.775 <sup>a</sup> (0.084)	-0.183 <sup>a</sup> (0.045)	-0.335 <sup>a</sup> (0.048)	-0.017 <sup>a</sup> (0.003)	0.004 <sup>c</sup> (0.002)	0.009 <sup>a</sup> (0.002)
Baa spread	-0.933 <sup>a</sup> (0.106)	0.088 (0.072)	-0.419 <sup>a</sup> (0.084)	0.020 <sup>a</sup> (0.004)	0.003 (0.003)	0.010 <sup>a</sup> (0.002)
F-test Statistics ( $I_-$ )	1.04	0.73	25.24 <sup>a</sup>	1.58	0.66	27.45 <sup>a</sup>
( $_1=I_{+2}$ )	Yes	Yes	Yes	Yes	Yes	Yes
Industry fixed effects	Yes	Yes	Yes	Yes	Yes	Yes
Rating Indicators	0.42	0.53	0.41	0.20	0.24	0.18
N	3020	2822	4941	3020	2822	4941

**Table VII: Change of Distance-to-Default and  $\sigma_A$  and Maturity Choice**

In columns (1) and (5), the dependent variable is the change of distance-to-default. In columns (2) and (6), the dependent variable is the change of asset volatility. The changes are measured from the month of issuance to 24 months after issuance. In columns (3) and (7), the dependent variable equals one if the change of  $DTD$  is positive. Similarly, in columns (4) and (8), the dependent variable equals one if  $\sigma_A$  is positive. The independent variables are changes from the fiscal year in which debt was issued to two years after. NBER recession dummy is equal to one if debt is issued in years 1990,1991,2001 and 2002, and zero otherwise. In parentheses are standard errors adjusted for heteroscedasticity and firm-year clustering. <sup>a</sup>, <sup>b</sup> and <sup>c</sup> denote significance at the 1%, 5%, and 10% levels, respectively. Marginal effects are reported for Probit regressions.

	OLS		Probit		OLS		Probit	
	$\Delta DTD$ (1)	$\Delta \sigma_A$ (2)	$\Delta DTD > 0$ (3)	$\Delta \sigma_A > 0$ (4)	$\Delta DTD$ (5)	$\Delta \sigma_A$ (6)	$\Delta DTD > 0$ (7)	$\Delta \sigma_A > 0$ (8)
Short-Term	0.899 <sup>a</sup> (0.249)	-0.019 <sup>b</sup> (0.009)	0.103 <sup>a</sup> (0.032)	-0.083 <sup>a</sup> (0.032)	0.428 <sup>c</sup> (0.246)	-0.003 (0.011)	0.086 <sup>b</sup> (0.037)	-0.020 (0.038)
Long-Term	-0.466 <sup>b</sup> (0.187)	0.022 <sup>a</sup> (0.007)	-0.026 (0.030)	0.049 <sup>c</sup> (0.028)	-0.585 <sup>a</sup> (0.191)	0.029 <sup>a</sup> (0.009)	-0.028 (0.032)	0.070 <sup>b</sup> (0.030)
NBER recession					1.570 <sup>a</sup> (0.428)	-0.030 <sup>b</sup> (0.013)	0.245 <sup>a</sup> (0.059)	-0.247 <sup>a</sup> (0.054)
NBER recession $\times$ short-term					1.257 <sup>b</sup> (0.546)	-0.049 <sup>a</sup> (0.015)	0.017 (0.074)	-0.230 <sup>a</sup> (0.060)
NBER recession $\times$ long-term					0.900 <sup>b</sup> (0.448)	-0.043 <sup>a</sup> (0.014)	0.029 (0.070)	-0.161 <sup>b</sup> (0.064)
$\Delta$ Size	-0.011 (0.321)	0.058 <sup>a</sup> (0.021)	0.124 <sup>b</sup> (0.059)	0.257 <sup>a</sup> (0.052)	-0.039 (0.300)	0.058 <sup>a</sup> (0.020)	0.118 <sup>b</sup> (0.057)	0.267 <sup>a</sup> (0.050)

Table VII Continued

	OLS		Probit		OLS		Probit	
	$\Delta$ DTD (1)	$\Delta \sigma_A$ (2)	$\Delta$ DTD>0 (3)	$\Delta \sigma_A > 0$ (4)	$\Delta$ DTD (5)	$\Delta \sigma_A$ (6)	$\Delta$ DTD>0 (7)	$\Delta \sigma_A > 0$ (8)
$\Delta$ Market/Book asset	0.456 (0.308)	0.008 (0.007)	0.012 (0.028)	0.077 <sup>a</sup> (0.026)	0.523 <sup>c</sup> (0.301)	0.007 (0.007)	0.022 (0.029)	0.068 <sup>b</sup> (0.027)
$\Delta$ Leverage	-10.953 <sup>a</sup> (0.758)	0.020 (0.047)	-1.954 <sup>a</sup> (0.154)	0.234 <sup>b</sup> (0.111)	-9.917 <sup>a</sup> (0.750)	-0.007 (0.049)	-1.830 <sup>a</sup> (0.152)	0.058 (0.113)
$\Delta$ Profit	2.185 (1.687)	-0.138 <sup>c</sup> (0.074)	-0.005 (0.282)	-0.149 (0.211)	2.335 (1.627)	-0.141 <sup>c</sup> (0.073)	0.042 (0.280)	-0.207 (0.207)
$\Delta$ Tangibility	-1.054 (1.695)	-0.116 <sup>c</sup> (0.066)	0.202 (0.244)	-0.161 (0.226)	-1.268 (1.570)	-0.109 <sup>c</sup> (0.064)	0.154 (0.237)	-0.134 (0.220)
$\Delta$ CV(OI)	-0.679 (0.982)	0.040 (0.054)	0.022 (0.124)	-0.094 (0.096)	-0.943 (1.101)	0.048 (0.057)	-0.012 (0.137)	-0.038 (0.096)
$\Delta$ Treasury	-1.159 <sup>a</sup> (0.098)	0.031 <sup>a</sup> (0.004)	-0.140 <sup>a</sup> (0.014)	0.121 <sup>a</sup> (0.015)	-0.736 <sup>a</sup> (0.088)	0.020 <sup>a</sup> (0.004)	-0.097 <sup>a</sup> (0.015)	0.057 <sup>a</sup> (0.015)
$\Delta$ Baa-spread	-1.724 <sup>a</sup> (0.139)	0.043 <sup>a</sup> (0.005)	-0.194 <sup>a</sup> (0.019)	0.174 <sup>a</sup> (0.019)	-1.485 <sup>a</sup> (0.124)	0.036 <sup>a</sup> (0.005)	-0.174 <sup>a</sup> (0.019)	0.144 <sup>a</sup> (0.019)
F-test/ $\chi^2$ Statistics (Short-Term=Long-Term)	33.86 <sup>a</sup>	33.70 <sup>a</sup>	22.42 <sup>a</sup>	20.35 <sup>a</sup>	19.41 <sup>a</sup>	14.31 <sup>a</sup>	13.11 <sup>a</sup>	7.39 <sup>a</sup>
Industry fixed effects	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Rating Indicators	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
$R^2$ /Pseudo $R^2$	0.35	0.14	0.21	0.08	0.39	0.16	0.23	0.14
N	2602	2602	2602	2602	2602	2602	2602	2602

**Table VIII: Default Risk Changes for Issuers with Different Predicted and Actual Maturity**

This table reports the average values of distance-to-default and asset volatility ( $\sigma_A$ ) for subsamples of issuers based on their predicted and actual maturity choices. The predicted maturities are based on results reported in the Appendix. Columns (5) and (6) report p-values from a two-tailed t-test of the null hypothesis that default risk measures between year 2 and year 0, and between year 2 and year -1 are similar.

Predicted and actual Maturity	Default risk measure	Year relative to offer					
		-1	0	1	2	Yr.2 - Yr.0 p-value	Yr.2 - Yr.-1 p-value
		(1)	(2)	(3)	(4)	(5)	(6)
Predicted=Short and Actual=Long (N=127)	Distance-to-default $\sigma_A$	9.651 0.283	9.331 0.288	8.544 0.294	8.102 0.310	0.00 0.07	0.00 0.03
Predicted=Long and Actual = Long (N=1,092)	Distance-to-default $\sigma_A$	8.234 0.249	7.869 0.253	7.683 0.254	7.390 0.264	0.00 0.00	0.00 0.00
Predicted=Long and Actual= Short (N=560)	Distance-to-default $\sigma_A$	7.144 0.289	6.743 0.292	7.026 0.284	7.585 0.275	0.00 0.00	0.04 0.01
Predicted= Short and Actual=Short (N=186)	Distance-to-default $\sigma_A$	7.379 0.342	7.213 0.332	7.150 0.329	7.461 0.316	0.37 0.08	0.78 0.02



**Table IX**  
**Deviations from Predicted Maturity and Default Risk Changes**

Regressions of changes in default risk on actual-predicted maturity choice, changes in firm characteristics, and changes in interest rates. In Columns (1) and (2), the dependent variable is the change in distance-to-default. In Columns (3) and (4), the dependent variable is the change in asset volatility. Both changes are measured from the month of issuance to 24 months following issuance. The changes in firm characteristics similarly reflect changes from the fiscal year in which debt was issued to two years after. (Actual - Predicted) is the difference between actual maturity choice (1=short-term, 2=medium-term, and 3=long-term) and predicted maturity choice. The predicted maturity is estimated from a prediction model presented in the Appendix and it also takes three values (1=short-term, 2=medium-term, and 3=long-term). The standard errors reported in parentheses are corrected for heteroskedasticity and are clustered at the firm-year level. <sup>a</sup>, <sup>b</sup> and <sup>c</sup> denote significance at the 1%, 5%, and 10% level, respectively.

	OLS		Probit	
	$\Delta$ DTD (1)	$\Delta\sigma_A$ (2)	$\Delta$ DTD (3)	$\Delta\sigma_A$ (4)
<b>(Actual - Predicted)</b>	<b>-0.466<sup>a</sup></b> <b>(0.113)</b>	<b>0.011<sup>a</sup></b> <b>(0.003)</b>	<b>-0.049<sup>a</sup></b> <b>(0.012)</b>	<b>0.043<sup>a</sup></b> <b>(0.013)</b>
$\Delta$ Size	-0.193 (0.370)	0.059 <sup>b</sup> (0.024)	0.091 (0.066)	0.285 <sup>a</sup> (0.062)
$\Delta$ Market/book asset	0.356 -0.271	0.012 <sup>c</sup> -0.007	0.009 -0.028	0.091 <sup>a</sup> -0.032
$\Delta$ Leverage	-11.223 <sup>a</sup> (0.762)	0.040 (0.056)	-1.938 <sup>a</sup> (0.162)	0.271 <sup>b</sup> (0.112)
$\Delta$ Profit	2.104 (1.768)	-0.151 <sup>b</sup> (0.075)	0.066 (0.286)	-0.146 (0.228)
$\Delta$ Tangibility	-1.937 (2.712)	-0.096 (0.077)	0.183 (0.259)	-0.13 (0.243)
$\Delta$ CV(OI)	-1.981 <sup>b</sup> (0.965)	0.110 <sup>b</sup> (0.054)	-0.215 (0.151)	-0.021 (0.146)
$\Delta$ Treasury	-1.096 <sup>a</sup> (0.122)	0.032 <sup>a</sup> (0.004)	-0.130 <sup>a</sup> (0.016)	0.122 <sup>a</sup> (0.017)
$\Delta$ Baa spread	-1.578 <sup>a</sup> (0.150)	0.042 <sup>a</sup> (0.005)	-0.175 <sup>a</sup> (0.018)	0.173 <sup>a</sup> (0.021)
Industry fixed effects	Yes	Yes	Yes	Yes
Rating indicators	Yes	Yes	Yes	Yes
$R^2$ /Pseudo- $R^2$	0.33	0.12	0.20	0.08
N	2468	2468	2468	2468

## Appendix I

### Multinomial Logit Regressions Predicting Debt Maturity Choice

The table reports estimates from a multinomial logit regression of debt maturity on leverage, the market/book assets ratio, firm size, firm-size squared, asset maturity, abnormal earnings, income volatility, term spread and regulatory industry indicator. The dependent variable is the debt maturity choice with long-term debt as the omitted category. Asset maturity is defined as the ratio of gross PP&E (item 7) to depreciation (item 125). Abnormal earnings is estimated as the difference between this years earnings per share (item 57) and last years earnings per share divided by last years share price (item 199). The regulatory industry indicator variable takes a value of one for firms in railroads (SIC code 4011) through trucking (4210 and 4213) through 1980, airlines (4512) through 1978, telecommunication (4812 and 4813) through 1982, and gas and electric utilities (4900 and 4939), and zero otherwise. We define term spread as the difference between one-year interest series and ten-year interest series (Source:<http://www.federalreserve.gov/release>). Other variables are defined in Table 1. The regression also includes rating indicators, the coefficient estimates of which are suppressed. Numbers in parentheses are standard errors adjusted for heteroscedasticity and firm-year clustering. <sup>a</sup>, <sup>b</sup> and <sup>c</sup> denote significance at the 1%, 5% and 10% levels, respectively. Marginal effects are reported.

Coefficient Estimates		
	Short-term	Medium-term
Leverage	0.210 (0.620)	2.092 <sup>a</sup> (0.527)
Market/book assets	0.441 <sup>a</sup> (0.115)	0.320 <sup>a</sup> (0.104)
Firm-size	-2.150 <sup>a</sup> (0.575)	-2.662 <sup>a</sup> (0.428)
<i>Firm - size</i> <sup>2</sup>	0.138 <sup>a</sup> (0.033)	0.149 <sup>a</sup> (0.026)
Asset maturity	-0.025 <sup>c</sup> (0.013)	-0.030 <sup>b</sup> (0.012)
Abnormal earnings	0.550 (0.359)	0.536 (0.390)
<i>CV(OI)</i>	0.437 (0.542)	0.602 (0.495)
Term spread	0.057 (0.056)	0.092 (0.057)
Regulatory industry	0.195 (0.326)	0.317 (0.257)
Rating indicator variables	Yes	
Pesudo- <i>R</i> <sup>2</sup>	0.10	
N	2,627	